

Data Rich Hydrology, Perugia – 28 January, 1 February 2019

Beyond traditional extreme value theory: lessons learned from rainfall and hurricane intensity

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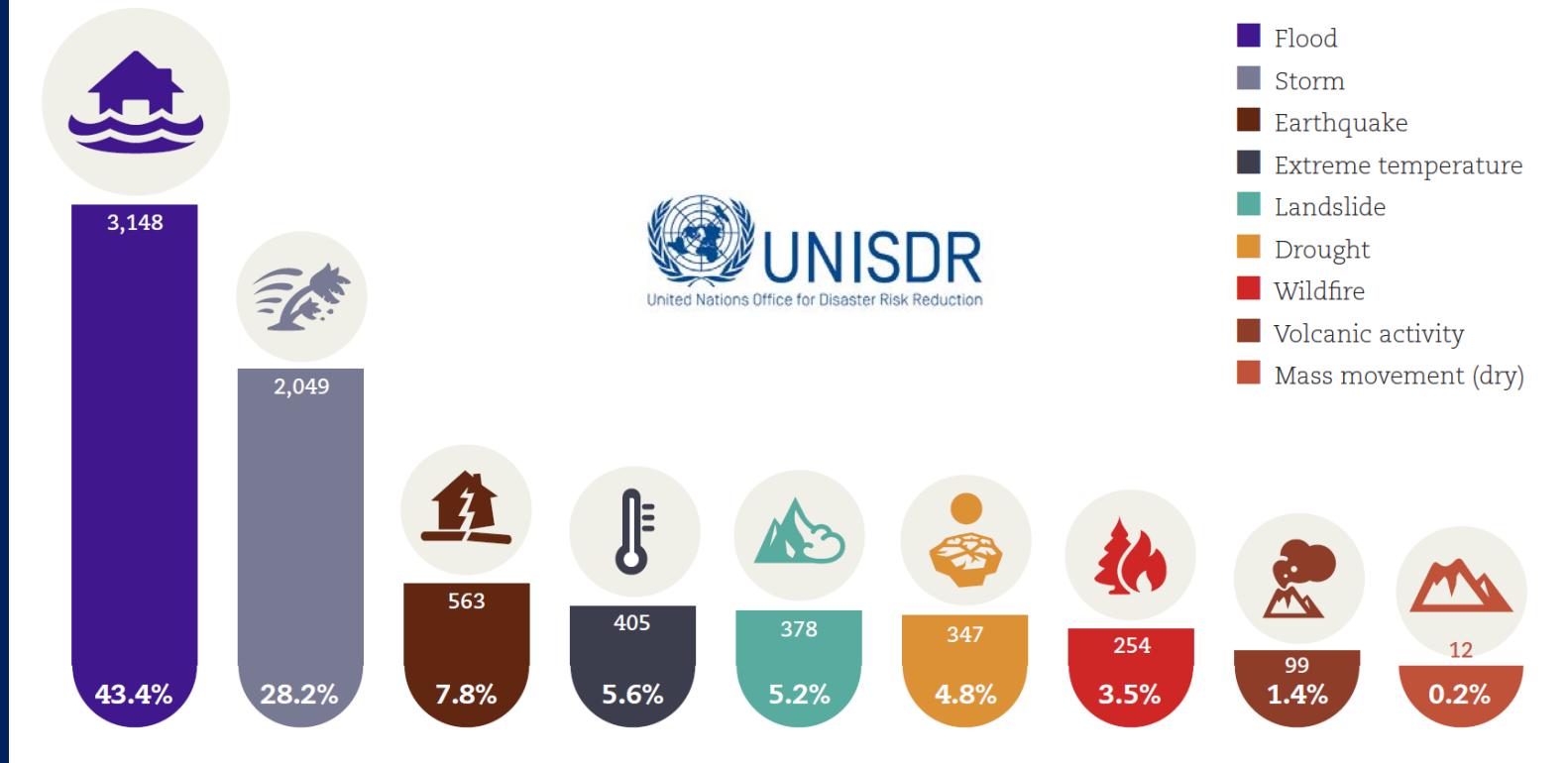


A little motivation. Extremes of all types...



Extreme events globally

Numbers of disasters per type 1998-2017



- 91% of disasters worldwide caused by floods, storms, droughts, heatwaves, and other extreme weather events;
- Extreme storms/floods affected ~3 billion people in 1998-2017, causing 69% of disaster-related economic losses ($\sim 2 \cdot 10^{12}$ US\$).

Some motivation

Given a stochastic quantity x ,

Expected lifetime (V_N)
Level of Risk accepted (R)



Return Time (T_r)



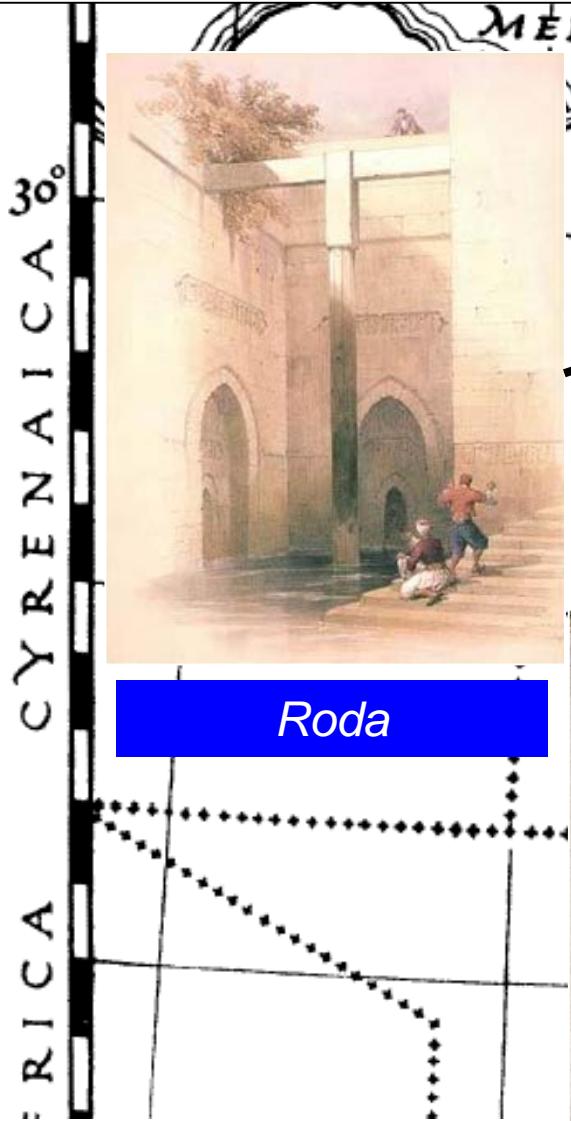
Estimated quantile $\hat{x}(T_r)$



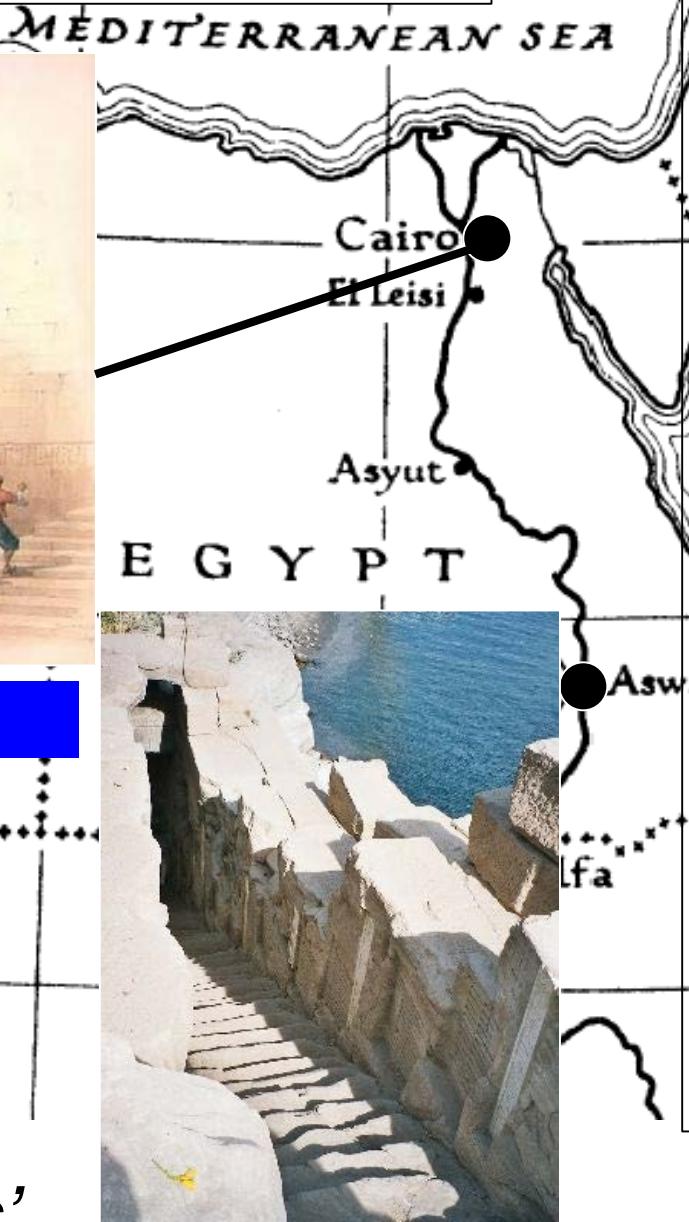
-Design of related structures
-Optimal water planning



Data Rich Hydrology

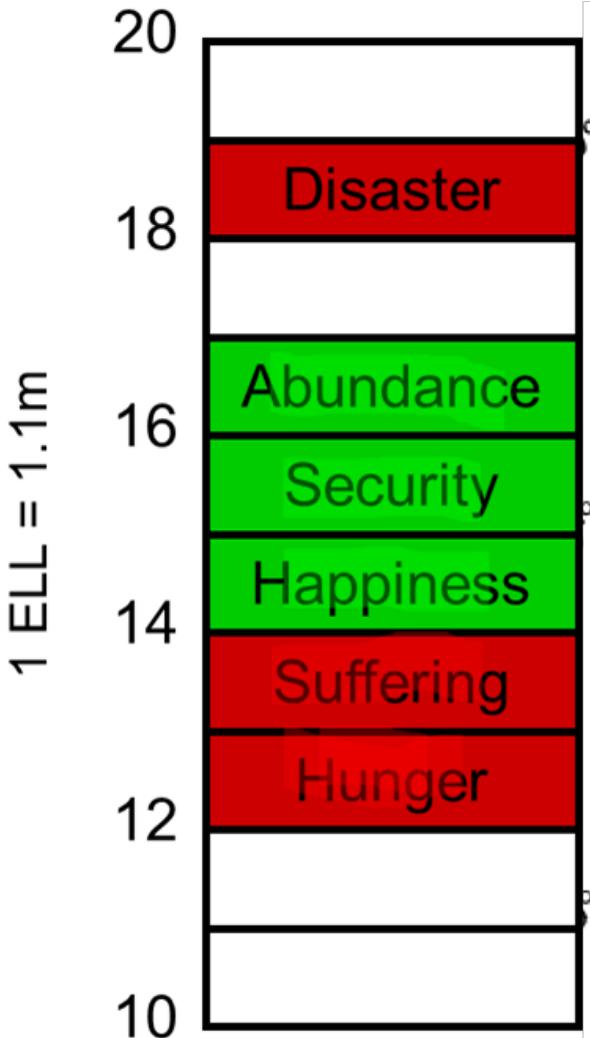


'Nilometers'



Elephantine island

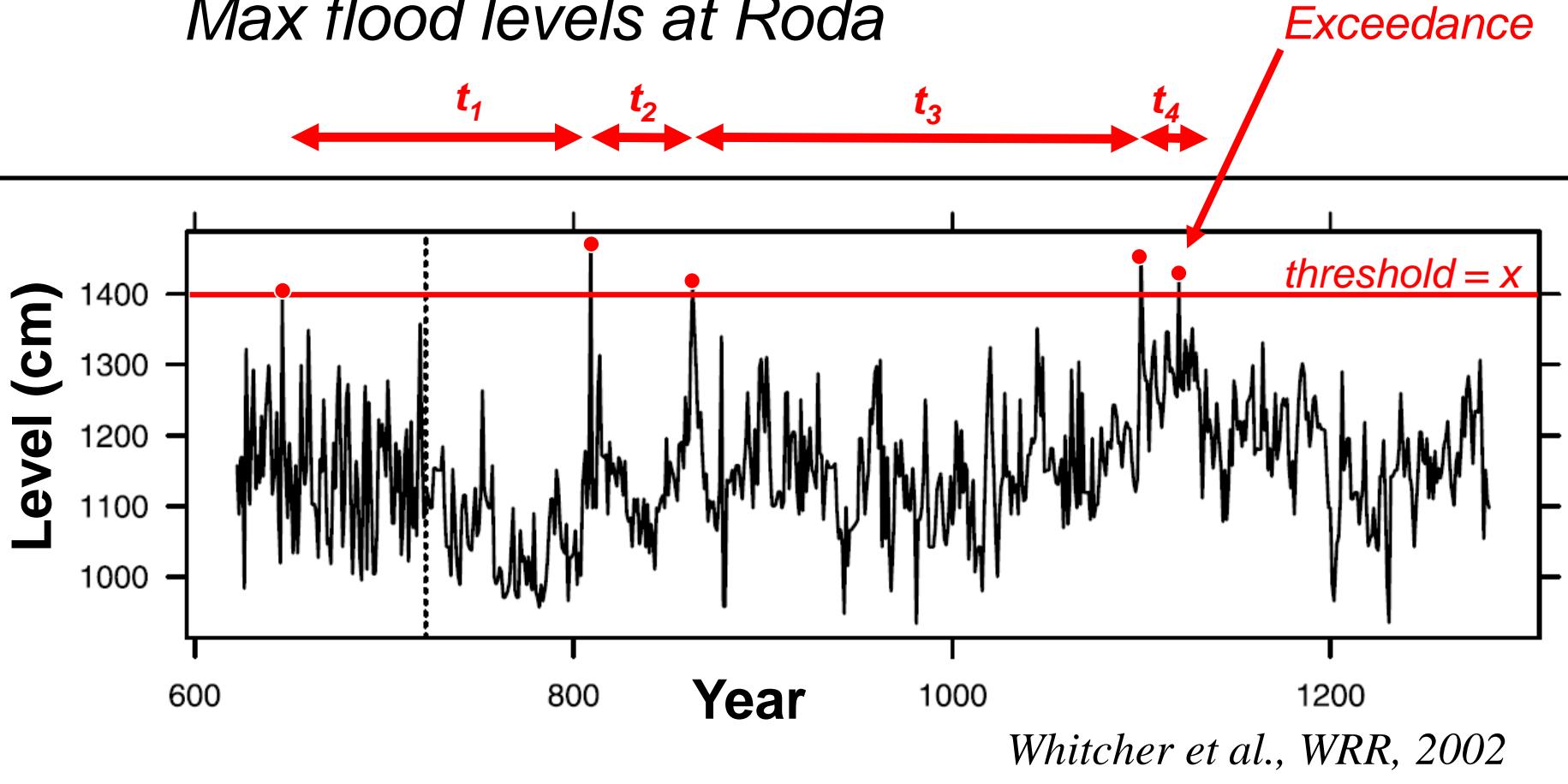
Pliny the Elder's calibration
of the River Nile's stage



Toussoun, 1922

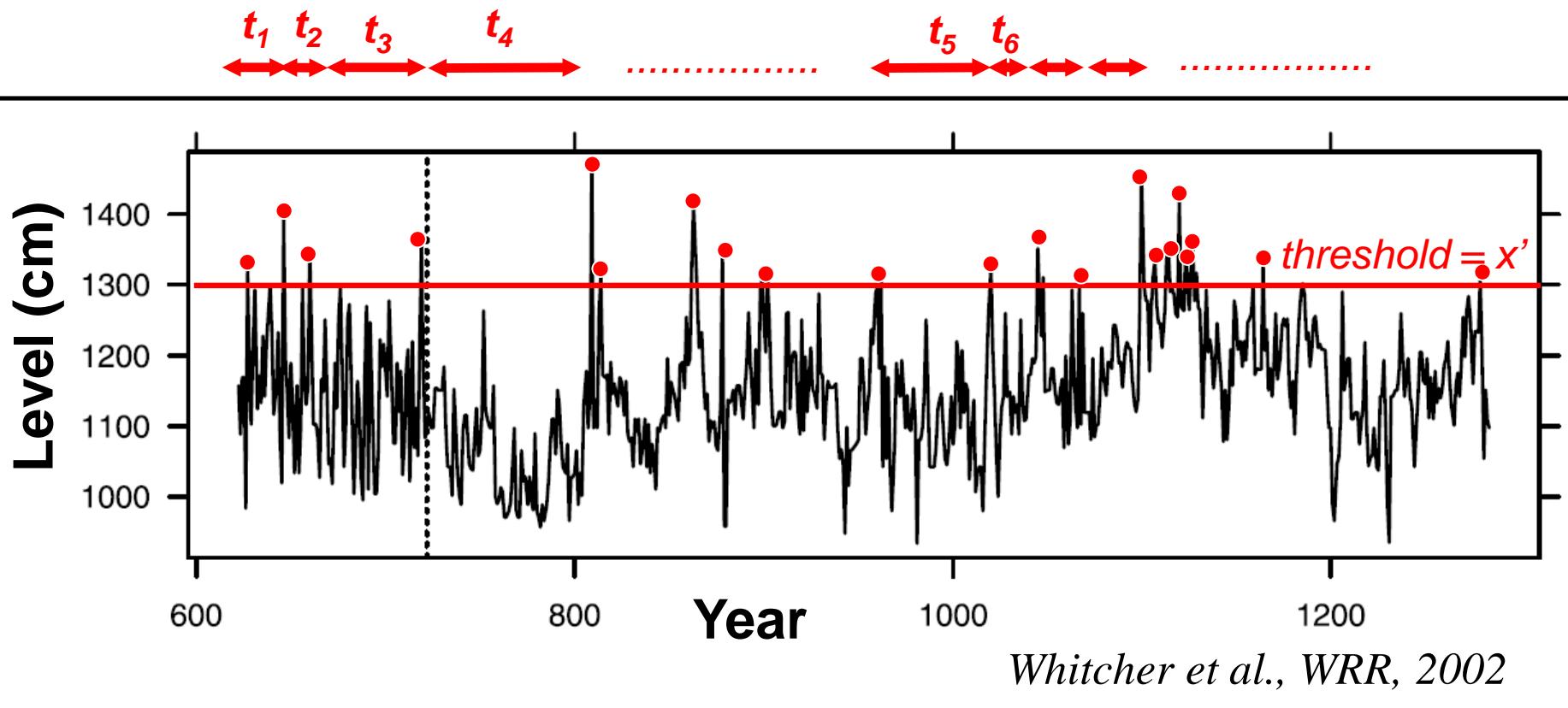
Hurst, 1927

Max flood levels at Roda



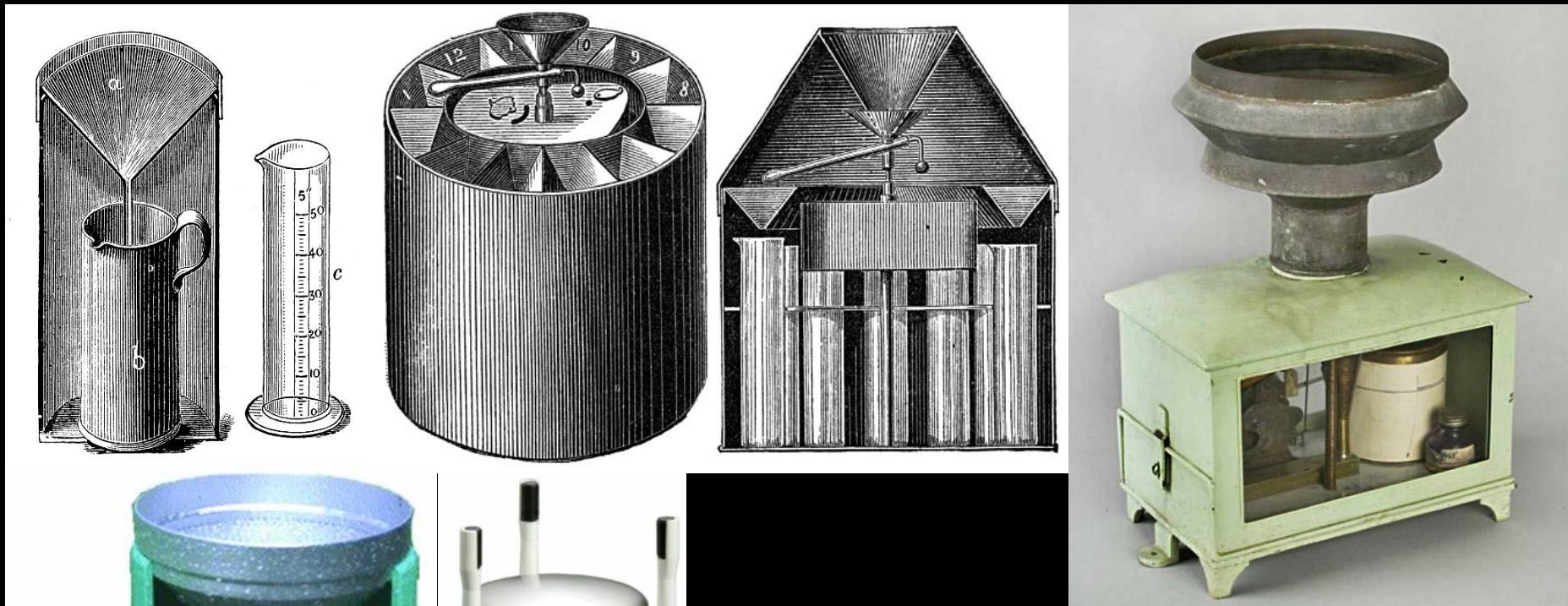
*Average recurrence interval or return time: $Tr(x) =$
average time intervals between two successive
exceedances of x*

Max floods levels at Roda



Because $x' < x$, then $Tr(x') < Tr(x)$

An important application: precipitation.



Data Rich Hydrology:

First obs.: India IV century BC

In Europe: Benedetto Castelli (1578-1643)

Sir Christopher Wren (1632-1723)

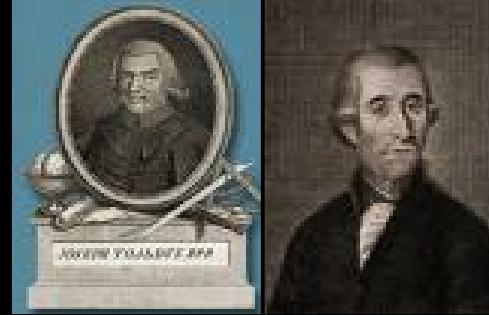
Giovanni Poleni (Venice 1683- Padova 1761)



Data Rich Hydrology: precipitation in Padova (1725- today)



Giovanni e Francesco Poleni
via Beato Pellegrino
1725-1764



Giuseppe Toaldo
e
Vincenzo Chiminello:
Specola 1768-1813



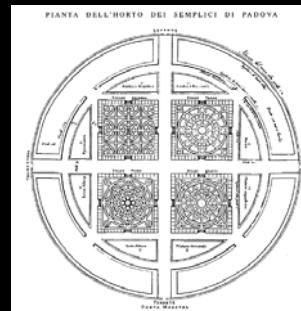
Giovanni Santini:
Specola 1823-1877



Giuseppe Lorenzoni:
Specola 1878-1934



Magistrato alle Acque di
Venezia: Osservatorio
G. Magrini
1936-1996



ARPAV
Orto Botanico:
1997-presente

Tabelle Brunnwasser, am Tage Oktober 1900 Blatt 100-101									
Zug der Gewässer	Name	Menge	Wert	Vorläufige Brunnwasserbeschaffung nach Paul Körner.					
				1	2	3	4	5	6
1	1	4.7	4.6	1	1	1	1	1	1
2	2	2.2	2.2	2	2	2	2	2	2
3	3	2.2	2.2	2	2	2	2	2	2
4	4	2.2	2.2	2	2	2	2	2	2
5	5	2.2	2.2	2	2	2	2	2	2
6	6	2.2	2.2	2	2	2	2	2	2
7	7	2.2	2.2	2	2	2	2	2	2
8	8	2.2	2.2	2	2	2	2	2	2
9	9	2.2	2.2	2	2	2	2	2	2
10	10	2.2	2.2	2	2	2	2	2	2
11	11	2.2	2.2	2	2	2	2	2	2
12	12	2.2	2.2	2	2	2	2	2	2
13	13	2.2	2.2	2	2	2	2	2	2
14	14	2.2	2.2	2	2	2	2	2	2
15	15	2.2	2.2	2	2	2	2	2	2
16	16	2.2	2.2	2	2	2	2	2	2
17	17	2.2	2.2	2	2	2	2	2	2
18	18	2.2	2.2	2	2	2	2	2	2
19	19	2.2	2.2	2	2	2	2	2	2
20	20	2.2	2.2	2	2	2	2	2	2
21	21	2.2	2.2	2	2	2	2	2	2
22	22	2.2	2.2	2	2	2	2	2	2
23	23	2.2	2.2	2	2	2	2	2	2
24	24	2.2	2.2	2	2	2	2	2	2
25	25	2.2	2.2	2	2	2	2	2	2
26	26	2.2	2.2	2	2	2	2	2	2
27	27	2.2	2.2	2	2	2	2	2	2
28	28	2.2	2.2	2	2	2	2	2	2
29	29	2.2	2.2	2	2	2	2	2	2
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31	31	2.2	2.2	2	2	2	2	2	2
32	32	2.2	2.2	2	2	2	2	2	2
33	33	2.2	2.2	2	2	2	2	2	2
34	34	2.2	2.2	2	2	2	2	2	2
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36	36	2.2	2.2	2	2	2	2	2	2
37	37	2.2	2.2	2	2	2	2	2	2
38	38	2.2	2.2	2	2	2	2	2	2
39	39	2.2	2.2	2	2	2	2	2	2
40	40	2.2	2.2	2	2	2	2	2	2
41	41	2.2	2.2	2	2	2	2	2	2
42	42	2.2	2.2	2	2	2	2	2	2
43	43	2.2	2.2	2	2	2	2	2	2
44	44	2.2	2.2	2	2	2	2	2	2
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100	100	2.2	2.2	2	2	2	2	2	2
101	101	2.2	2.2	2	2	2	2	2	2
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144	144	2.2	2.2	2	2	2	2	2	2
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147	147	2.2	2.2	2	2	2	2	2	2
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154	154	2.2	2.2	2	2	2	2	2	2
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161	161	2.2	2.2	2	2	2	2	2	2
162	162	2.2	2.2	2	2	2	2	2	2
163	163	2.2	2.2	2	2	2	2	2	2
164	164	2.2	2.2	2	2	2	2	2	2
165	165	2.2	2.2	2	2	2	2	2	2

Chimi-

1 punto=0,188 mm

Anno	Gennaio	punti	mm
1833	1	0	0
	2	0	0
	3	0	0
	4	0	0
	5	0	0
	6	0	0
	7	0	0
	8	0	0
	9	0	0
	10	0	0
	11	0	0
	12	0	0
	13	0	0
	14	0	0

SE 4	Nuvolone	1 W 0 2 W 0 3 W 0 4 W 0 5 E 0 6 S 1	Nuvolo Nuvola Nuvola sereno sereno
NE 0	Vario		
NO 0	Vario		Sera secca
NW 0	severo		
WS 0	severo		Mattina nebbia giorno; temperato
W 0	severo		
NW 0	Nuvolo		Mattina nebbia giorno vario; so-
W 0	vario		

serno ; giorno varie
sempre sollecito

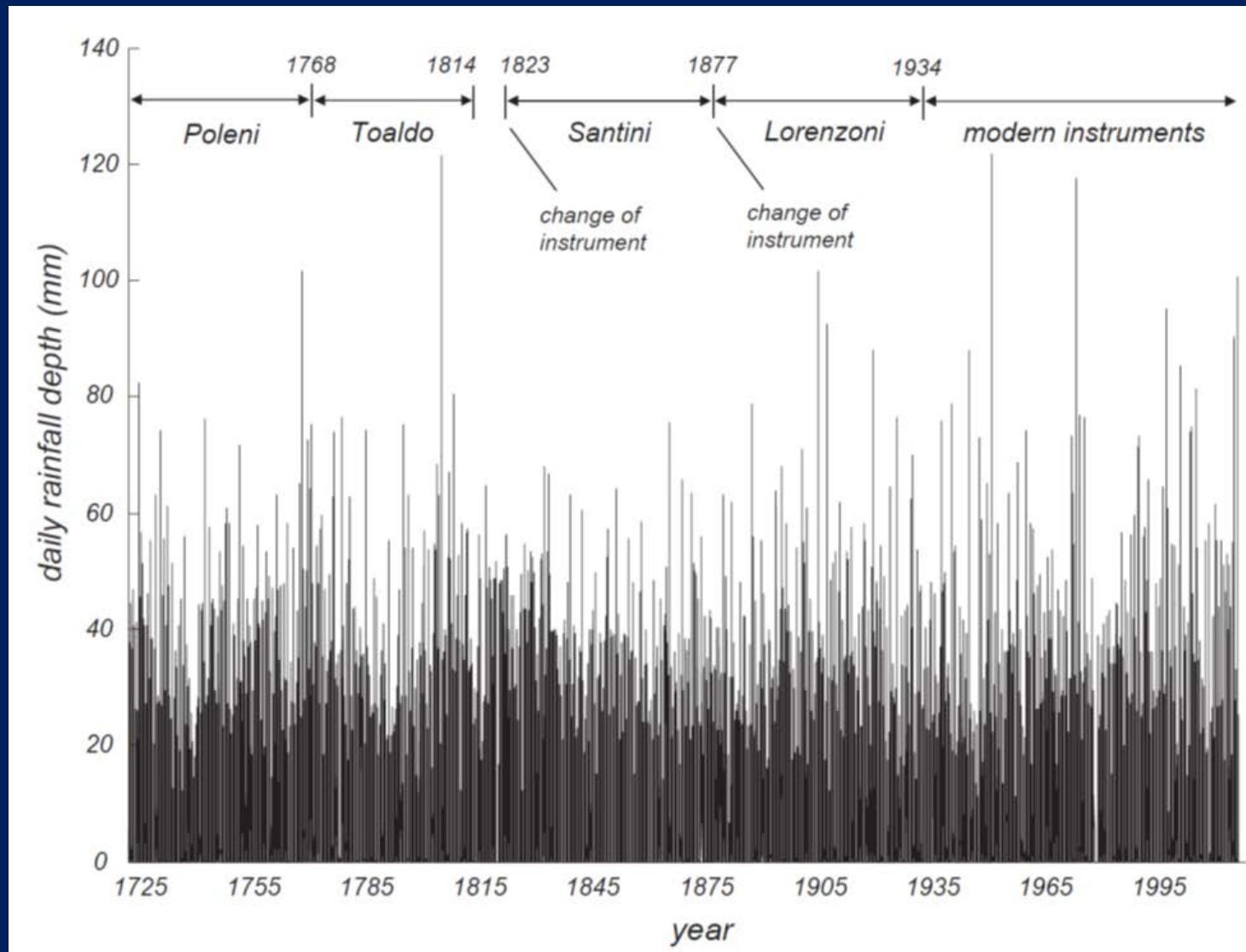
Mattina nebbia. Solazzo tutto il giorno; temperatura dolce.

Brattina nebbia, il resto del giorno vario; temperatura dolce

19	28.2,9A		4,5	2,8	22	W	0	Nuvolo
	28.3,25	28.3,22	4,9	3,2	36	W	0	Nuvolo
	28.5,6A		4,7	3,9	30	W	0	Nuvolo
	28.5,24		4,4	1,8	26	W	0	Nuvolo
20	28.5,25	28.5,20	5,7	4,7	42	SE	0	Nuvolo
	28.5,25		3,0	3,1	36	SE	0	Nuvolo

Original records at the Padova observatory

The Padova daily precipitation time series 1725-2006

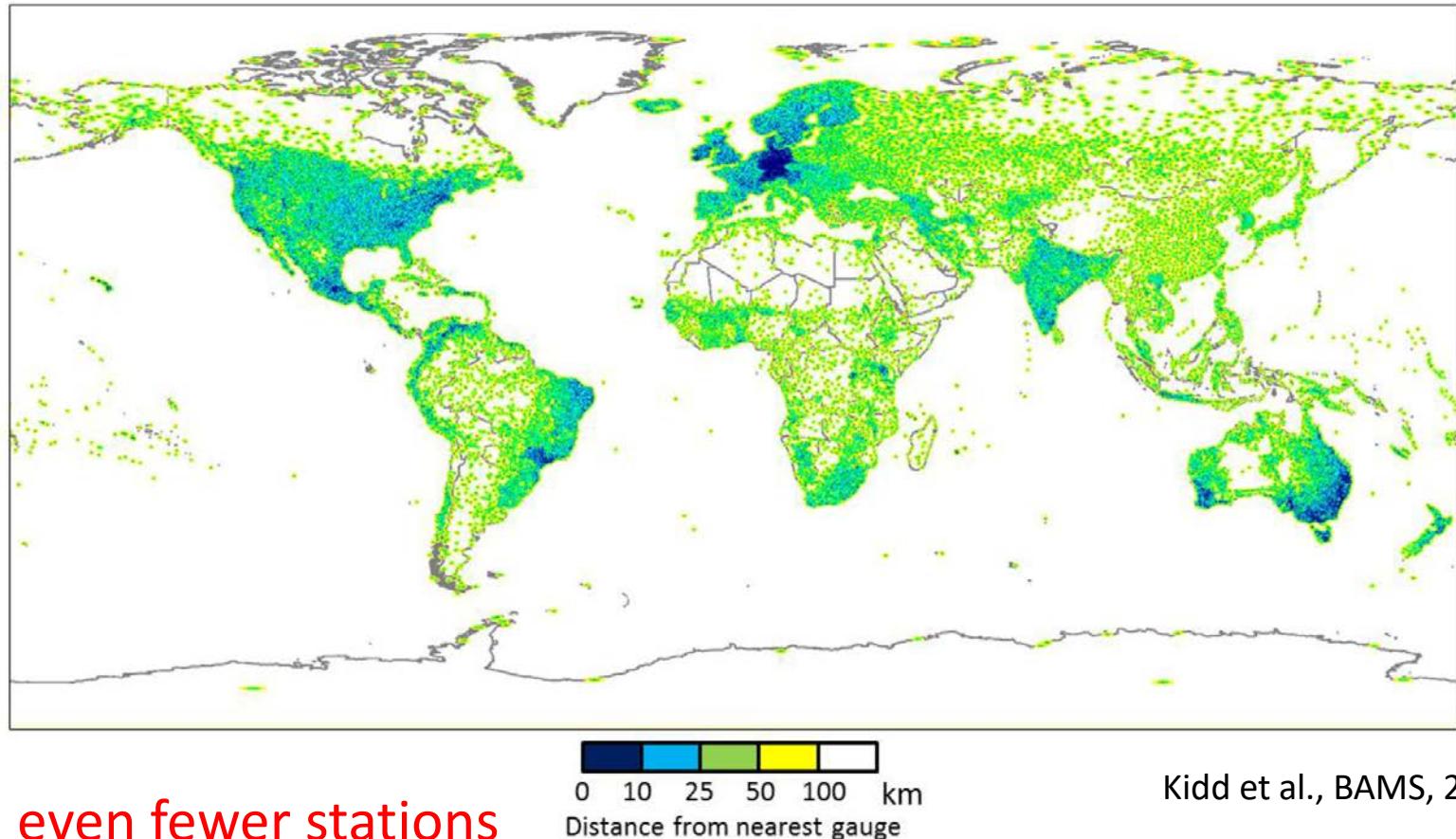


(Marani and Zanetti, 2015)



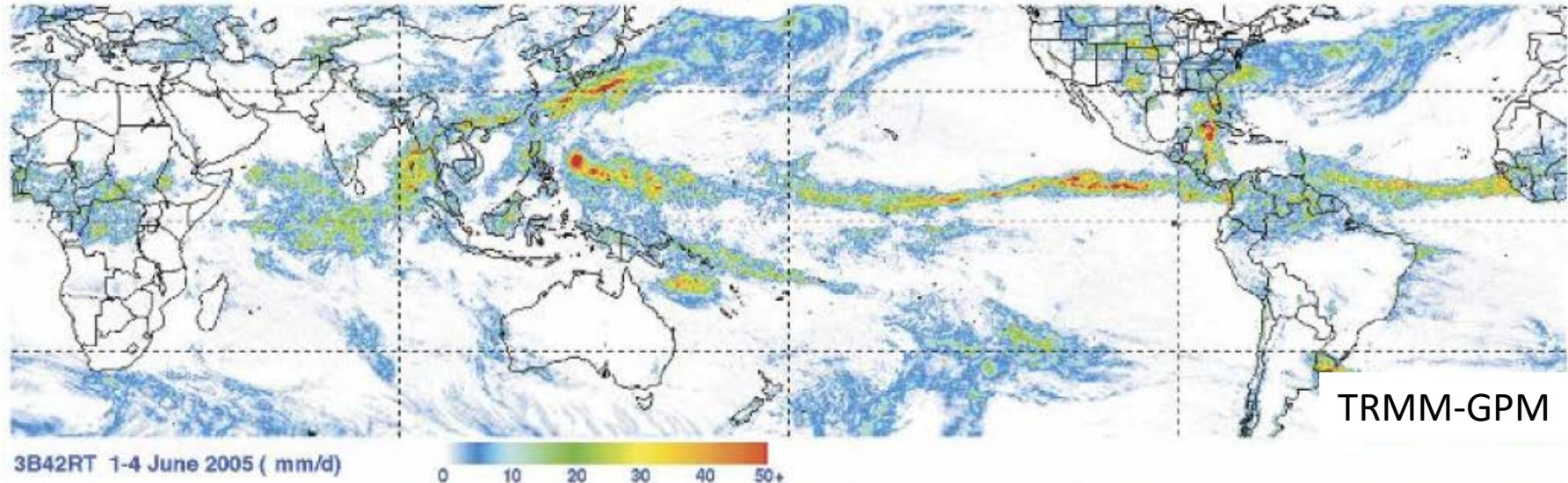
Data Rich Hydrology?

Global Rainfall Extremes ? Large gaps in observational networks



And even fewer stations
are long enough for
extreme rainfall analysis

Kidd et al., BAMS, 2016

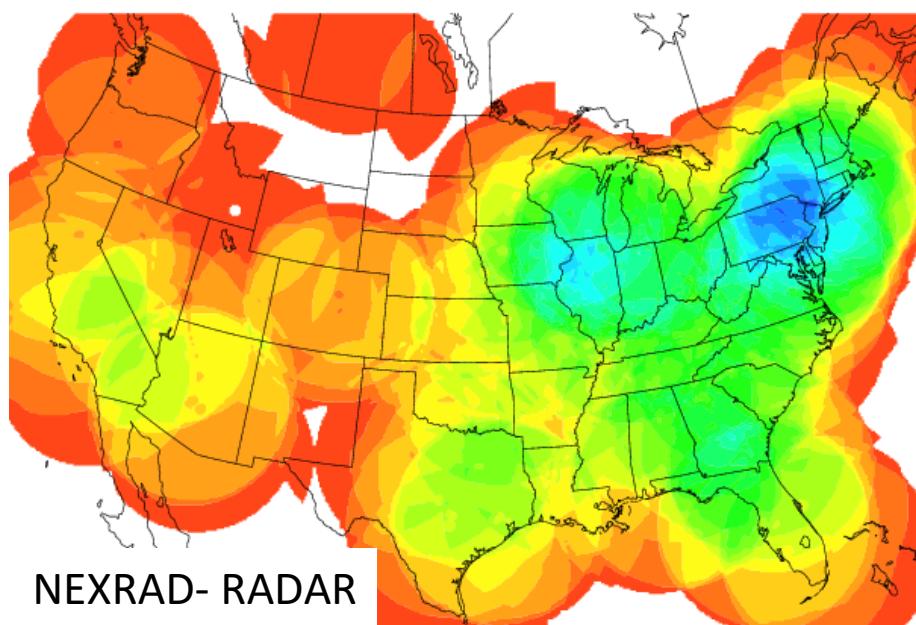


3B42RT 1-4 June 2005 (mm/d)

0 10 20 30 40 50+

TRMM-GPM

Stations in Range



NEXRAD- RADAR

Remote sensing estimates of rainfall

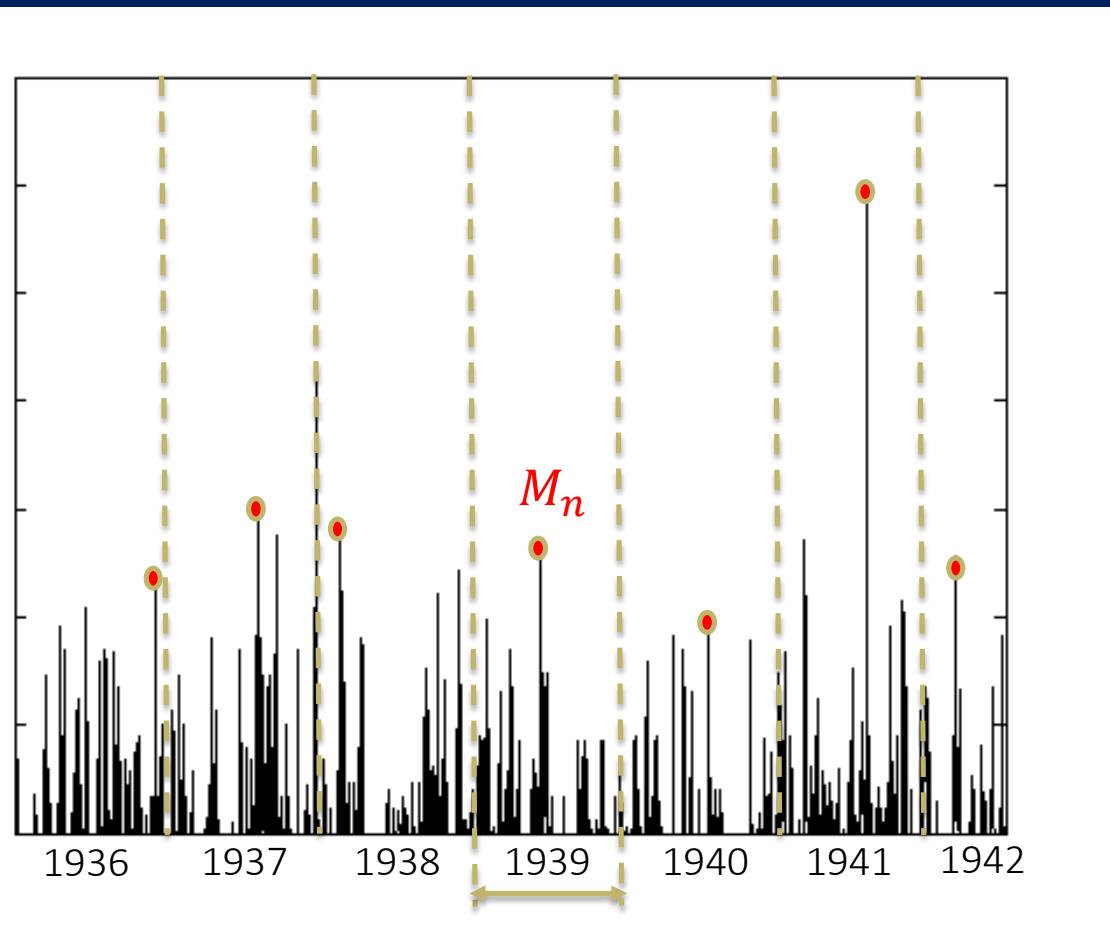
- Continental to quasi-global coverage;
- Short observational period (max < 20 yrs);
- Testing vs point obs ?
- Downscaling ?

Classical Extreme Value Theory (EVT)

[Fischer-Tippett-Gnedenko, 1928-1943]

$$x_n = \max_n(x_i)$$

$$H_n(x) = F(x)^n \quad \text{for independent, identically distributed } x_i$$



Often block size = one year

'Block Maxima':

M_n = Maximum value occurred within a n -event block

What is the distribution of M_n ?

Three-Type Theorem:

-As $n \rightarrow \infty$
-After proper renormalization, are only three 'types' of asymptotic distribution for M_n :

- *Gumbel*
- *Frechet*
- *Weibull*

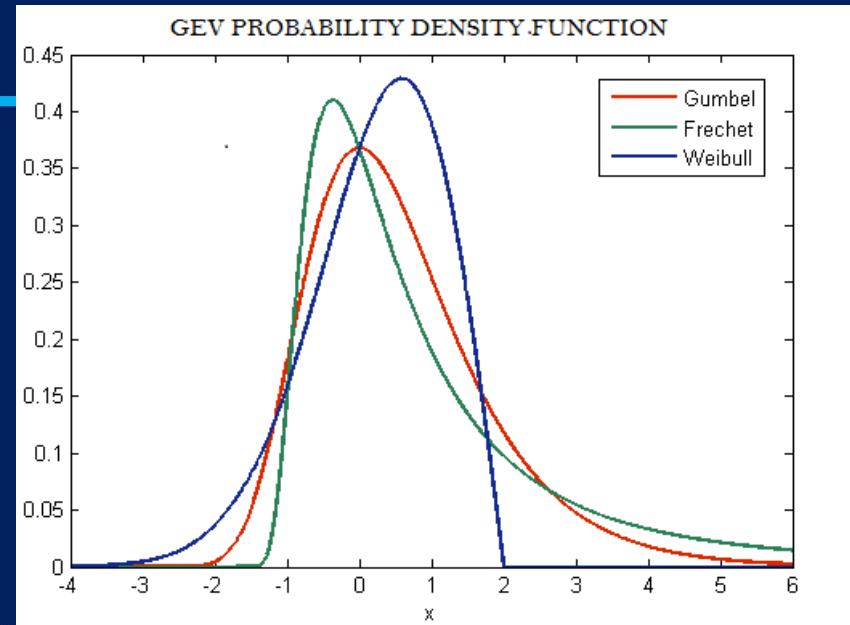
The GEV distribution

[Von Mises, 1936]

$$H(x) = \exp \left\{ - \left[1 + \frac{\xi}{\psi} (x - \mu) \right]_+^{-\frac{1}{\xi}} \right\}$$

Shape Parameter:

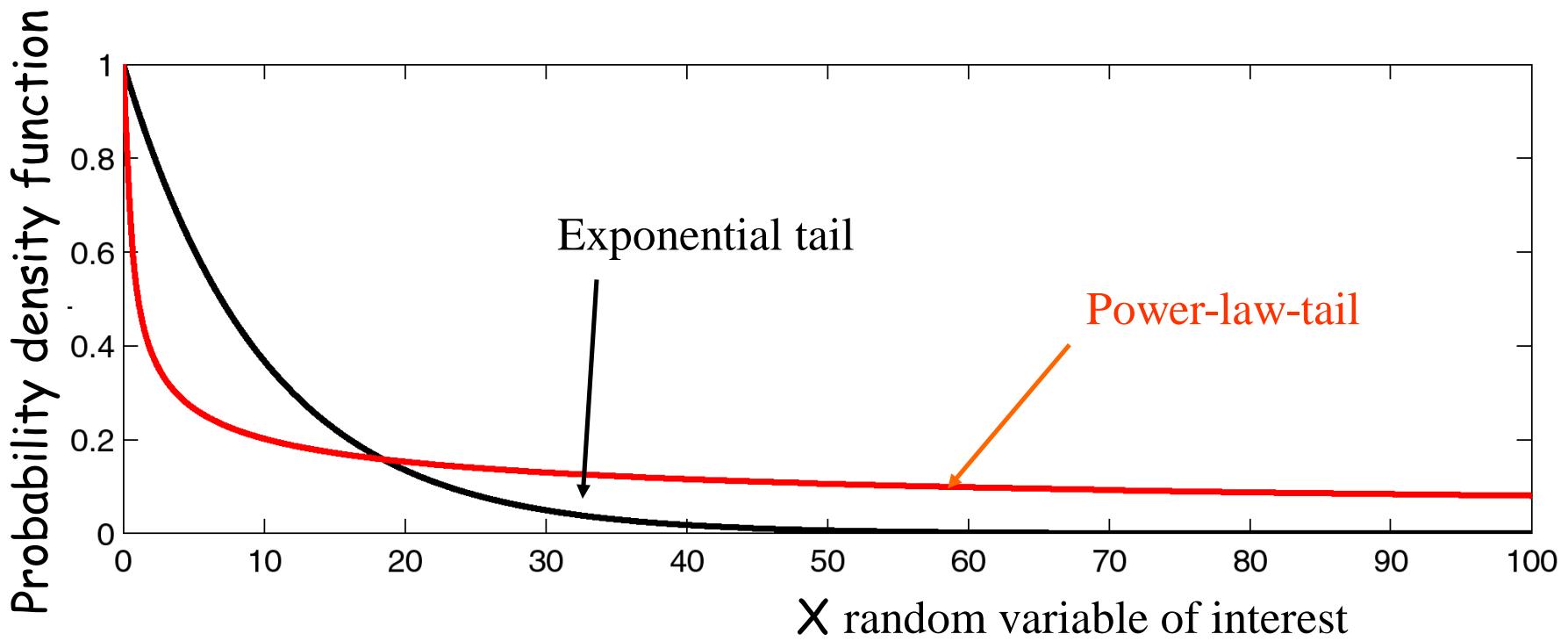
- $\xi > 0$ Frechet, 'heavy' tailed
- $\xi \rightarrow 0$ Gumbel, exponential tail;
- $\xi < 0$ Inverse Weibull, upper bounded.



Quantiles corresponding to return periods of interest can be retrieved:

$$F_r = 1 - \frac{1}{T_r} \quad \hat{h} = H^{-1}(F_r)$$

Exponential and Power-law decays



Two radically different types of extremes!



Peak Over Threshold Method (POT)

[Balkema, De Haan & Pickand, 1975; Davison and Smith, 1990]

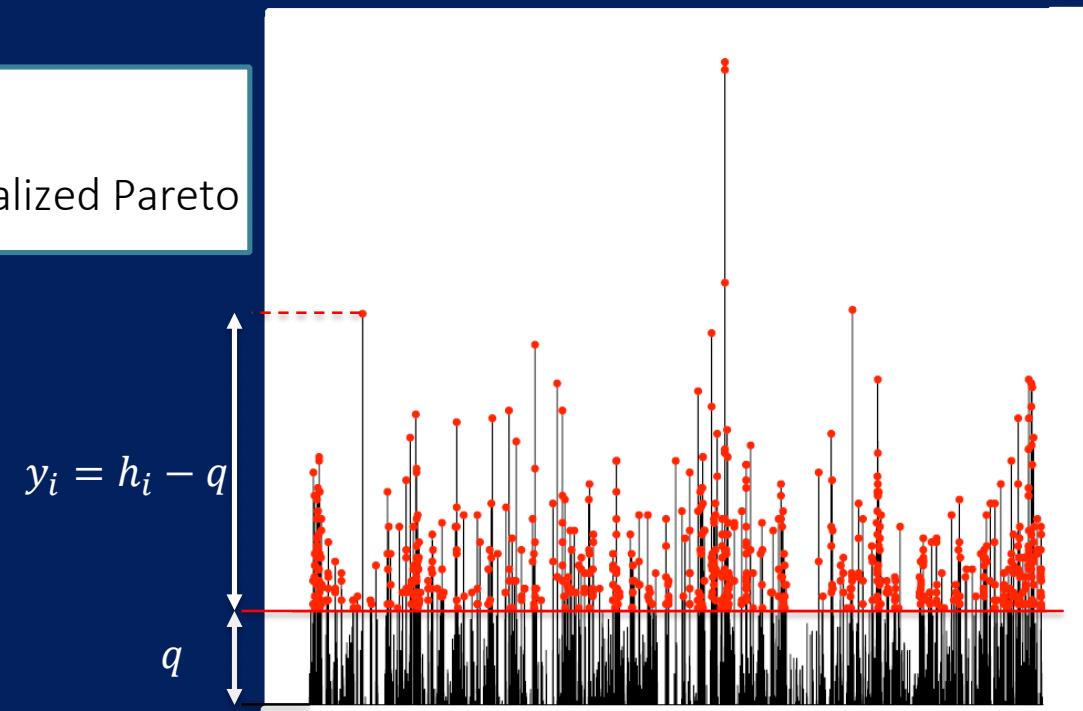
For a fixed threshold $q \rightarrow$ Exceedances $Y_i = H_i - q$ i.i.d. r.v.

$$P(Y_{max} < x) = \sum_{n=1}^{\infty} p(n) \cdot F(x)^n = \sum_{n=1}^{\infty} \frac{\lambda^n e^{-\lambda}}{n!} \cdot \left\{ 1 - \left[1 + \frac{\xi}{\psi} \cdot (x - q) \right]^{-1/\xi} \right\}^n$$

- Exceedances arrivals \longrightarrow Poisson
- Distribution of excesses \longrightarrow Generalized Pareto

Advantages:

1. Better description of the 'tail'
2. Consistent with GEV



Considerations on the application of the classical EVT

- Uncertain convergence of actual distribution to limiting one (some suggest observed Frechet is a result of incomplete convergence to Gumbel, e.g. Koutsoyiannis, 2013; Serinaldi and Kilsby, 2014).
- Fitting of GEV using Maximum Likelihood only uses yearly maxima and neglects most of the data. Important in the presence of short observations.
- Use of POT requires identification of threshold, uses more data but still a fraction of all available information.
- When number of events is small (dry climates, hurricanes, ...), yearly maxima also come from bulk of distribution, not just the tail (we are far from a limiting form)

A Metastatistical Extreme Value Distribution (MEVD)

The Block-maxima distribution

$$x_n = \max_n(x_i)$$

$$H_n(x) = F(x)^n$$

for independent, identically distributed X'_i 's

The compound block-maxima distribution accounting for stochastic n and parameters of “ordinary” events pdf

$$\zeta(x) = \sum_{n=1}^{\infty} \int_{\Omega_{\vec{\Theta}}} F(x; \vec{\theta})^n g(n, \vec{\theta}) d\vec{\theta}$$

$G(n, \theta)$ = joint prob distrib. of the parameters.

A Metastatistical Extreme Value distribution (MEV)

The MEV block-maxima distribution:

$$\zeta(x) = \sum_{n=1}^{\infty} \int_{\Omega_{\vec{\Theta}}} F(x; \vec{\theta})^n g(n, \vec{\theta}) d\vec{\theta}$$

$G(n, \theta)$ = joint prob distrib. of the parameters.

Can be approximated using sample averages:

$$\zeta(x) \cong \frac{1}{T} \sum_{j=1}^T F(x; \vec{\theta}_j)^{n_j}$$

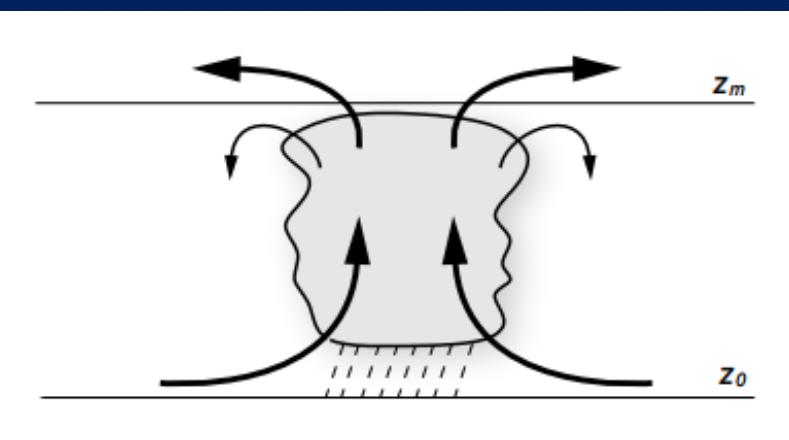
T = # sub-periods over which n and θ are estimated

Rainfall: what $F(x)$ - the pdf of daily «ordinary» rainfall?

[Wilson e Tuomi, 2005]

Assumptions

- Heavy precipitation events dominated by moisture advection
- Negligible contribution from local evaporation



- Simple two-layers atmospheric model
- Temporal average

Weibull \in domain of attraction of Gumbel, \longrightarrow
but convergence is very slow!

\bar{k} =precipitation efficiency

\bar{q} =specific humidity

m =advection mass

$$R_{acc} = \bar{k}\bar{q}m$$



$$F(x) = 1 - e^{(\frac{x}{C})^w}$$

Weibull Parent
distribution

It would need a disproportionately large number of events N for Gumbel to hold. Frechet is improperly used instead.

MEV-Weibull distribution

The MEV expression:

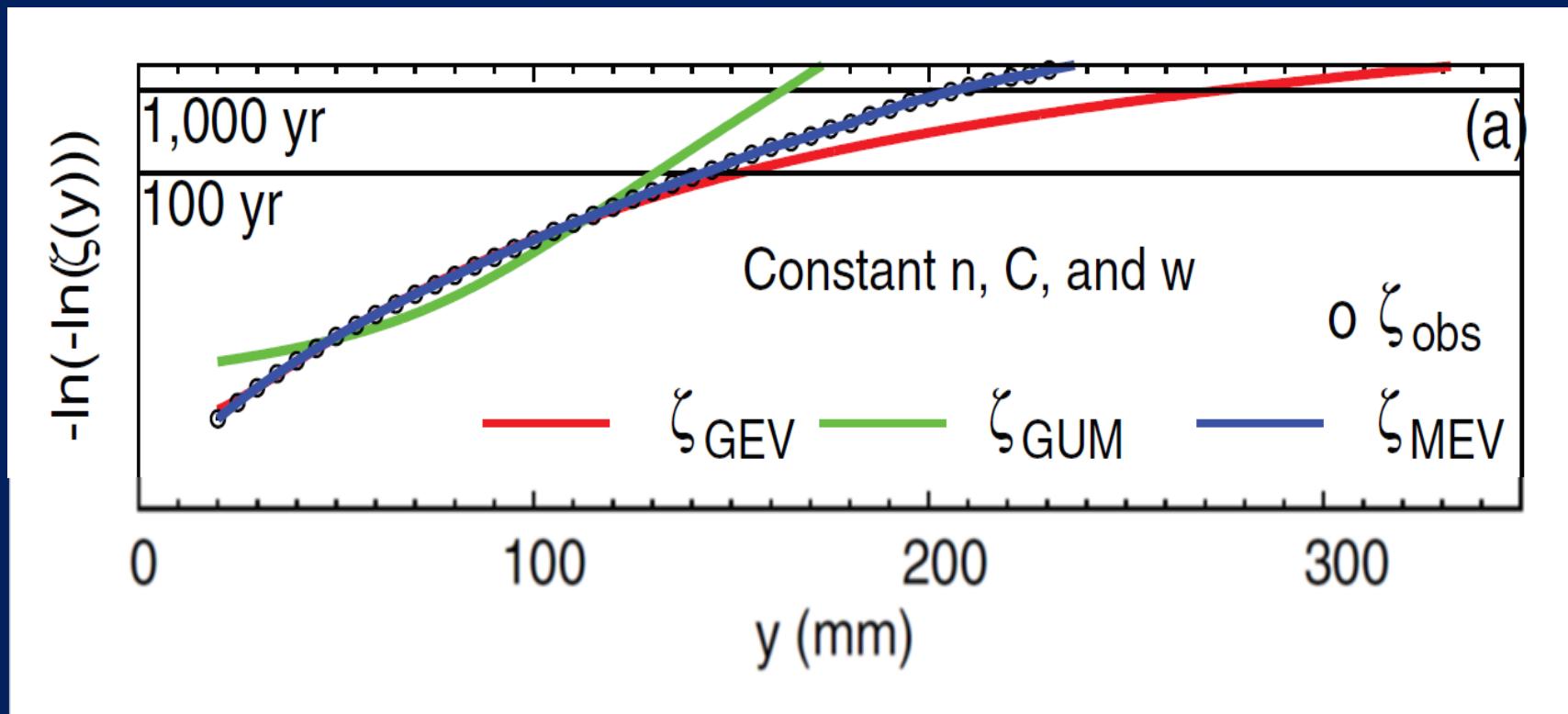
$$\zeta(x) \cong \frac{1}{T} \sum_{j=1}^T F(x; \vec{\theta}_j)^{n_j}$$

T = # sub-periods over which n
and θ are estimated

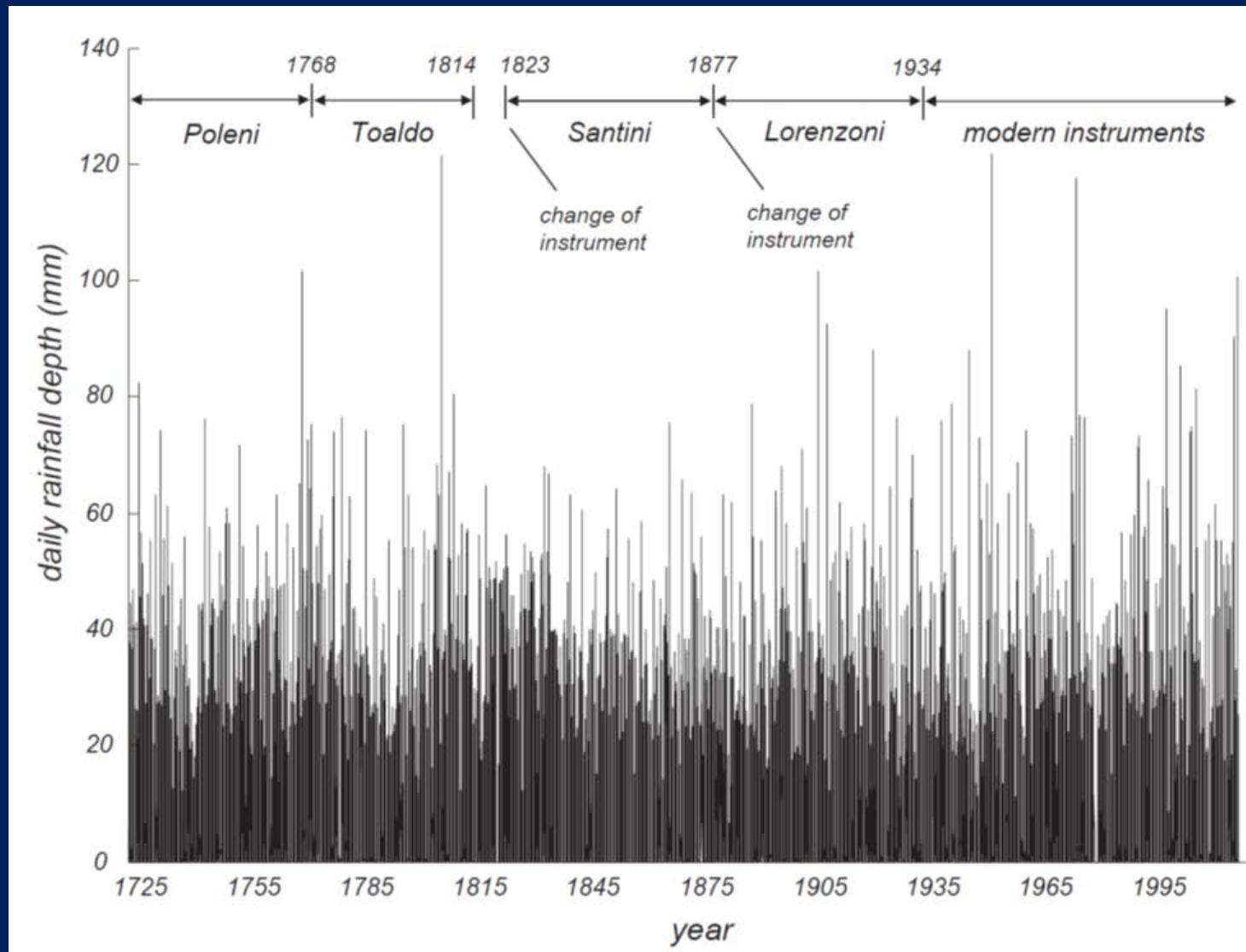
In the Weibull case becomes:

$$\zeta(x) \cong \frac{1}{T} \sum_{j=1}^T \left[1 - e^{\left(\frac{x}{c_j} \right)^{w_j}} \right]^{n_j}$$

Weibull-distributed synthetic data

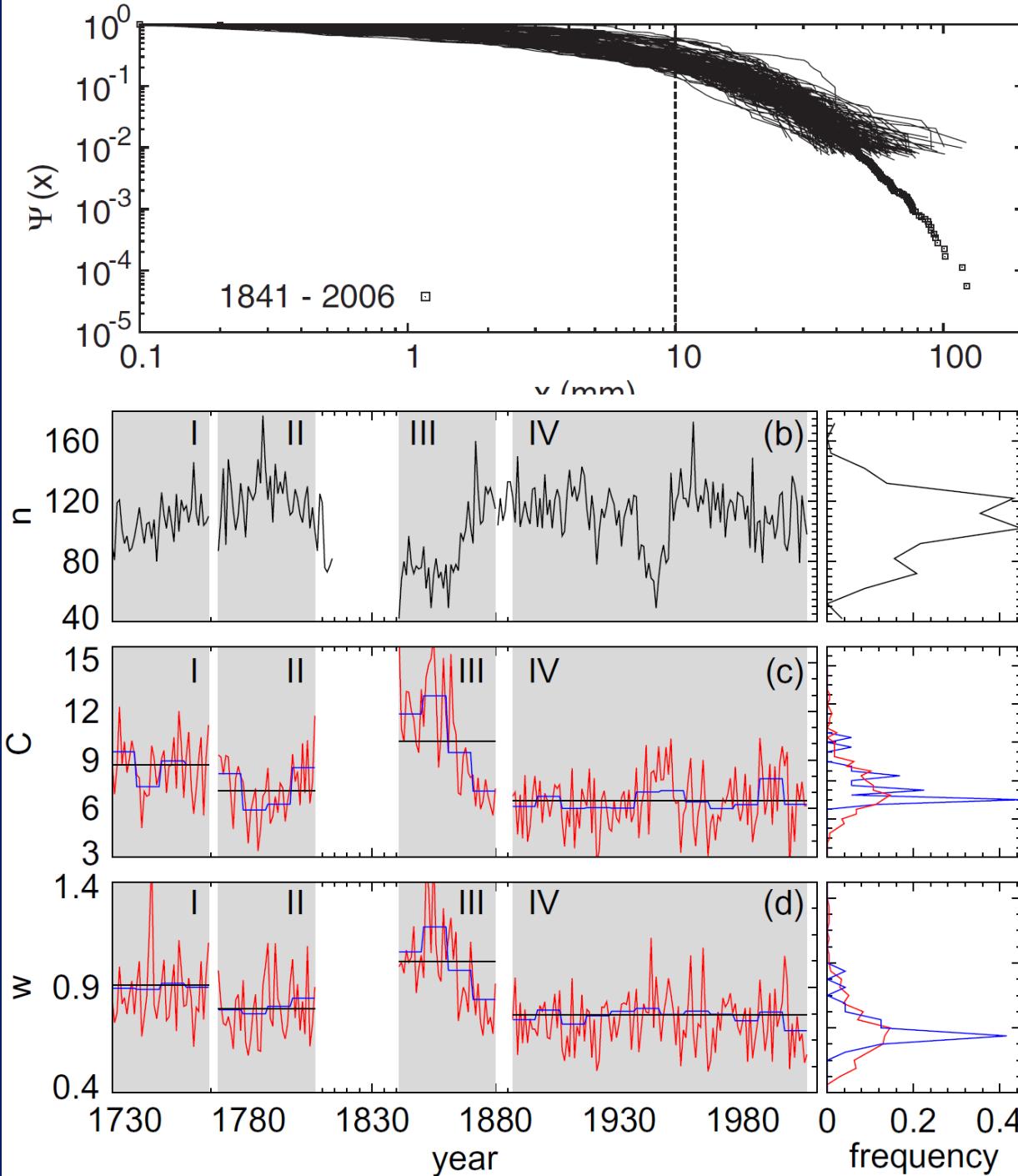


The Padova daily precipitation time series 1725-2006



(Marani and Zanetti, 2015)

Padova series:
Wide fluctuations in
pdf parameters
and in number
of events.

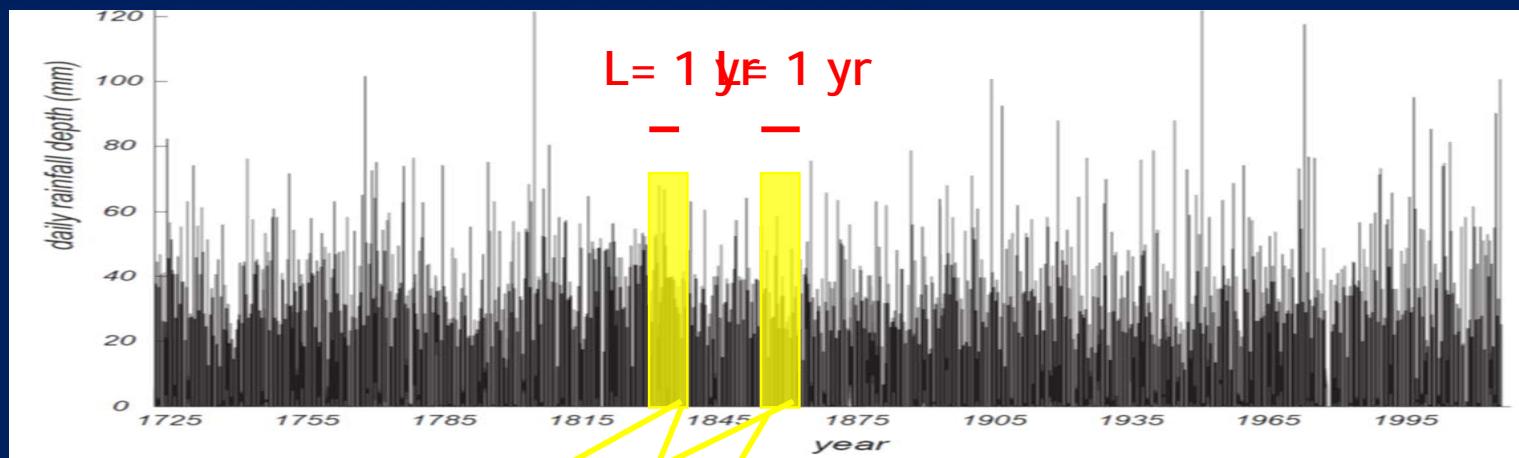


(Marani and
Ignaccolo, 2015)

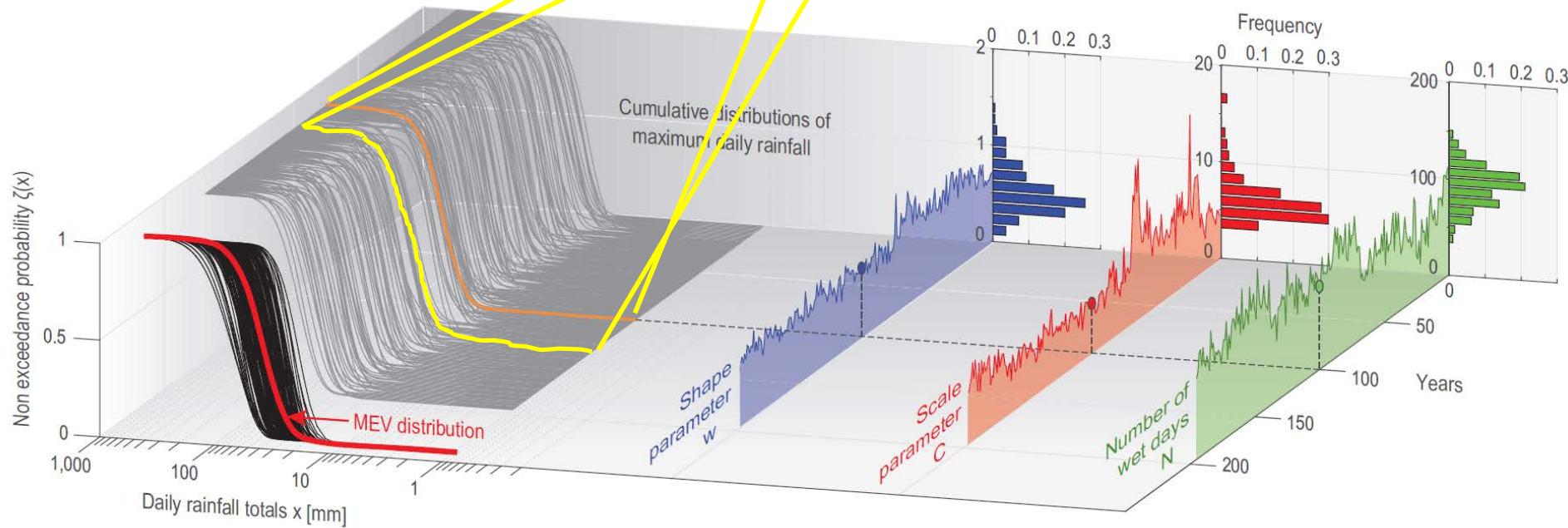
The Metastatistical Extreme Value distribution (MEV)

(Marani and Ignaccolo, AWR, 2015; Zorzetto et al., GRL, 2016)

Padova:
Daily rainfall
1725- today



Marani and Zanetti, WRR, 2015



Reality Check



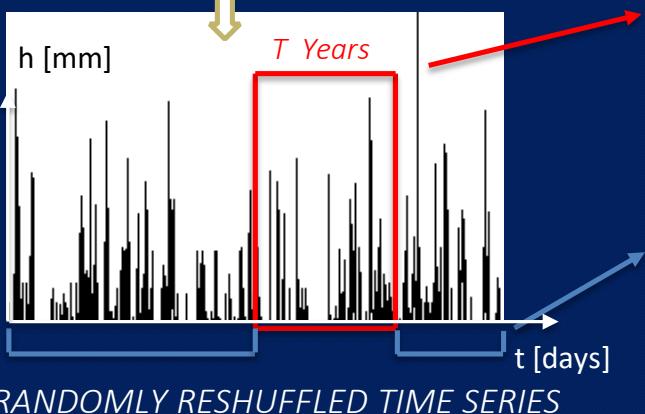
36 datasets, 106 -275 years of daily observations, ($\langle L \rangle = 135$ yrs)
Less than 5% of missing data

Method of analysis

Reshuffling of daily data preserving

- (1) yearly number of events, and
- (2) observed values (i.e. Pdf's)

- To eliminate correlation and non-stationarity
- Preserving the true (unknown) distribution of the parameters and numbers of wet days.



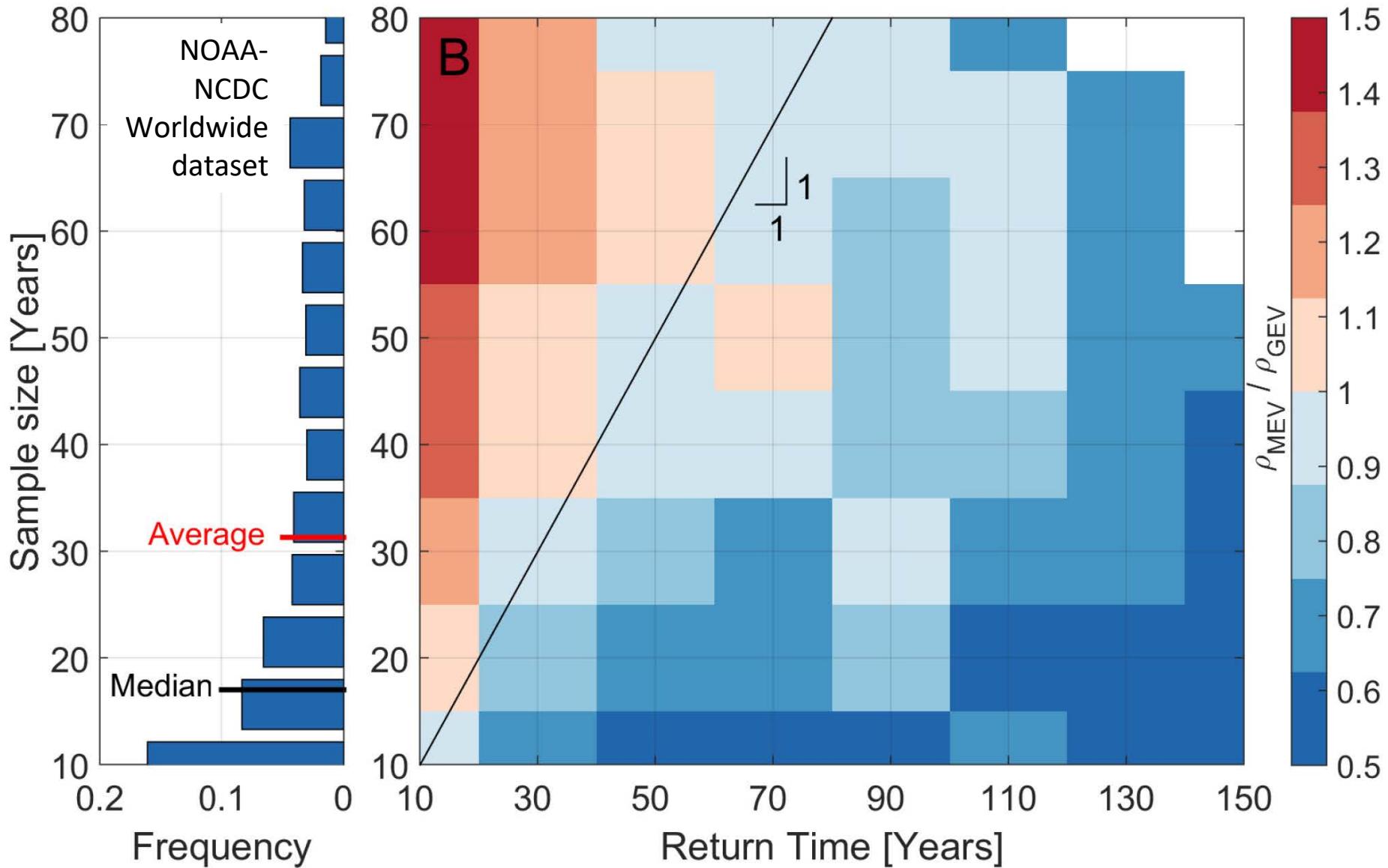
Selection of a T-year sample from which
compute $\hat{x}(T_r, T)$ with GEV, POT and MEV.

The remainder of the dataset is used to
observe the 'true value' $x_{obs}(T_r)$ to assess the
performance of EV methods

Non dimensional estimation error

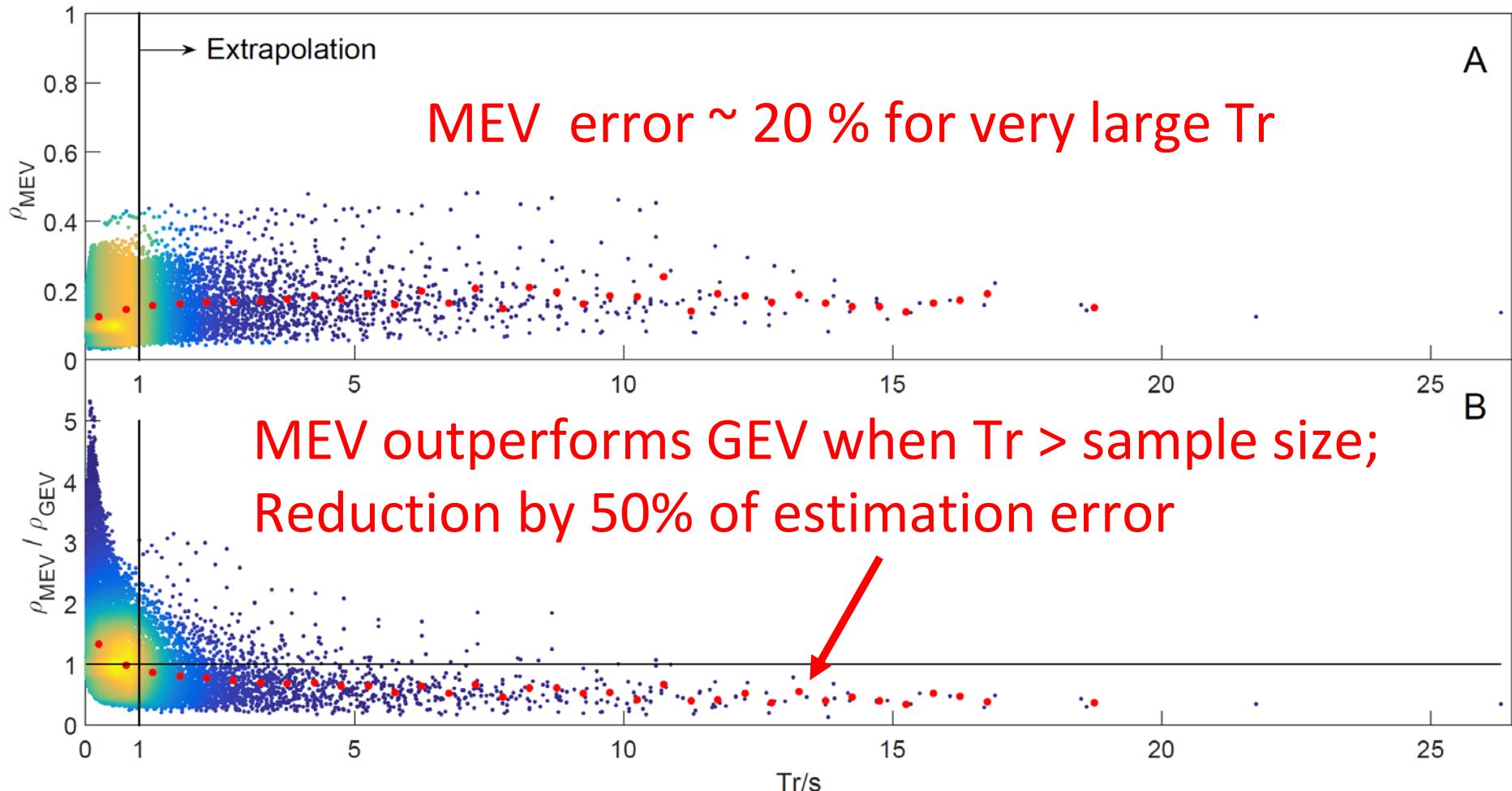
$$\epsilon = \sqrt{\frac{1}{N} \sum \left(\frac{\hat{x} - x_{obs}}{x_{obs}} \right)^2}$$

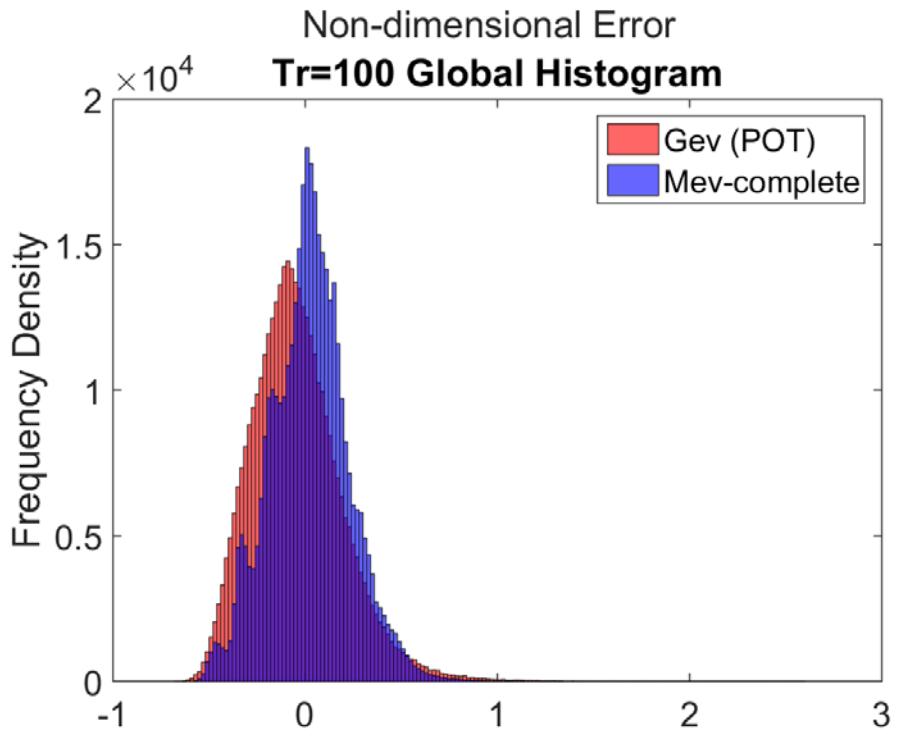
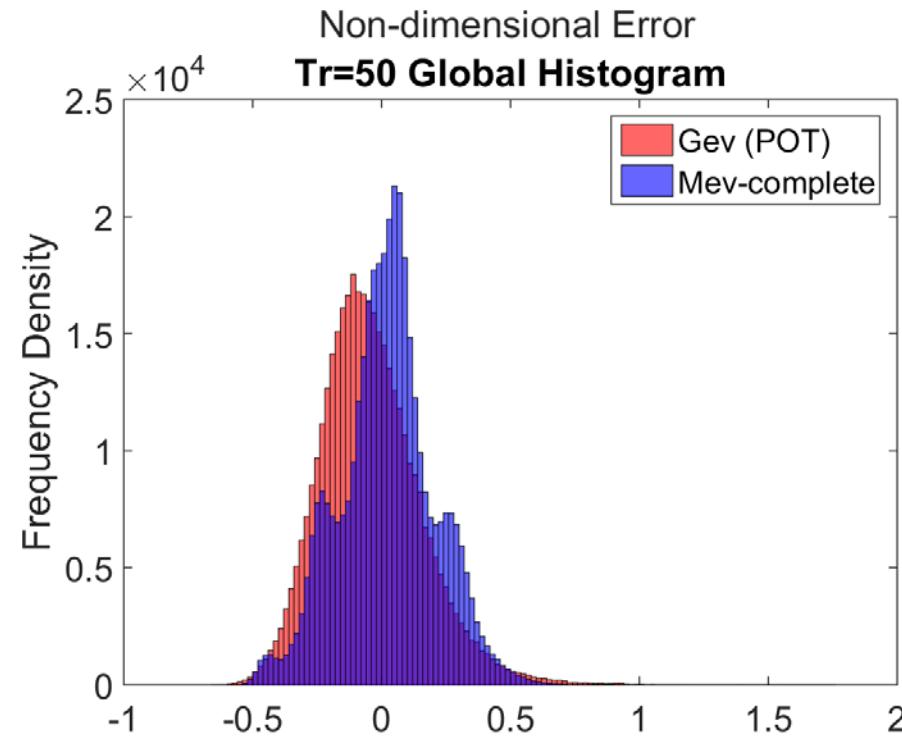
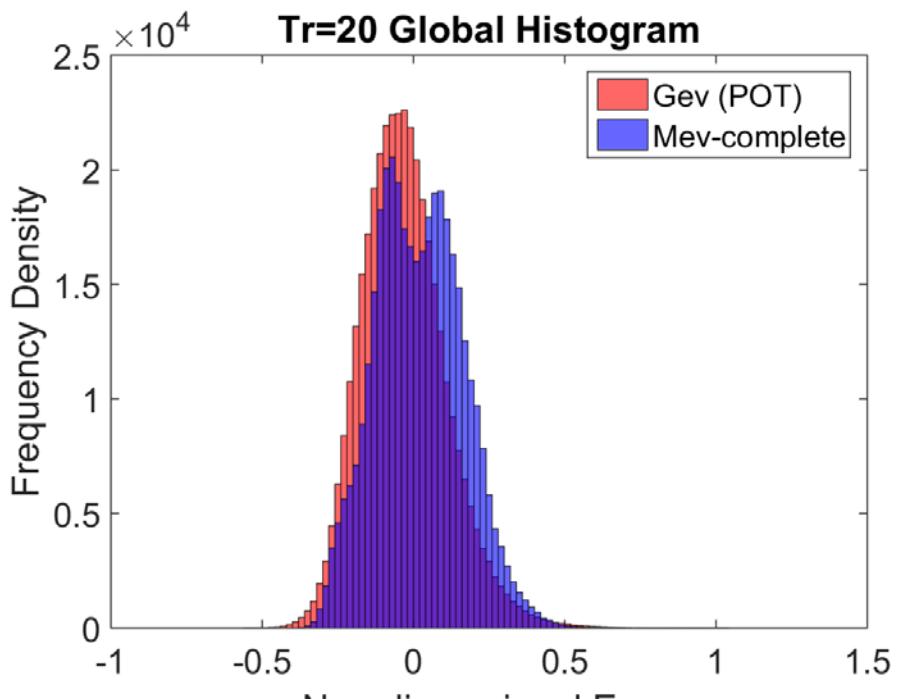
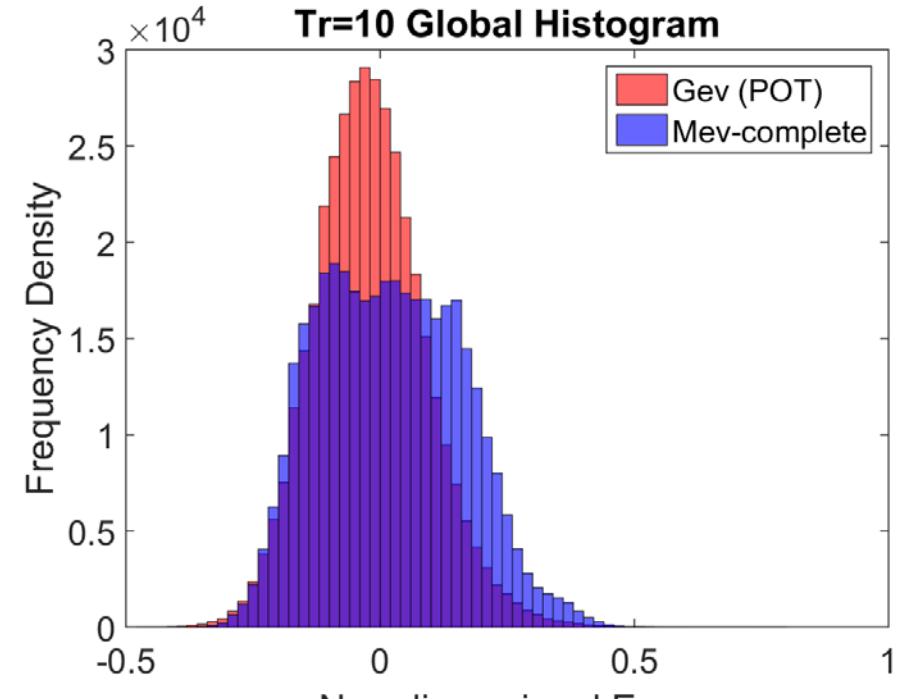
Ratio of MEV estimation error to GEV-LMOM error



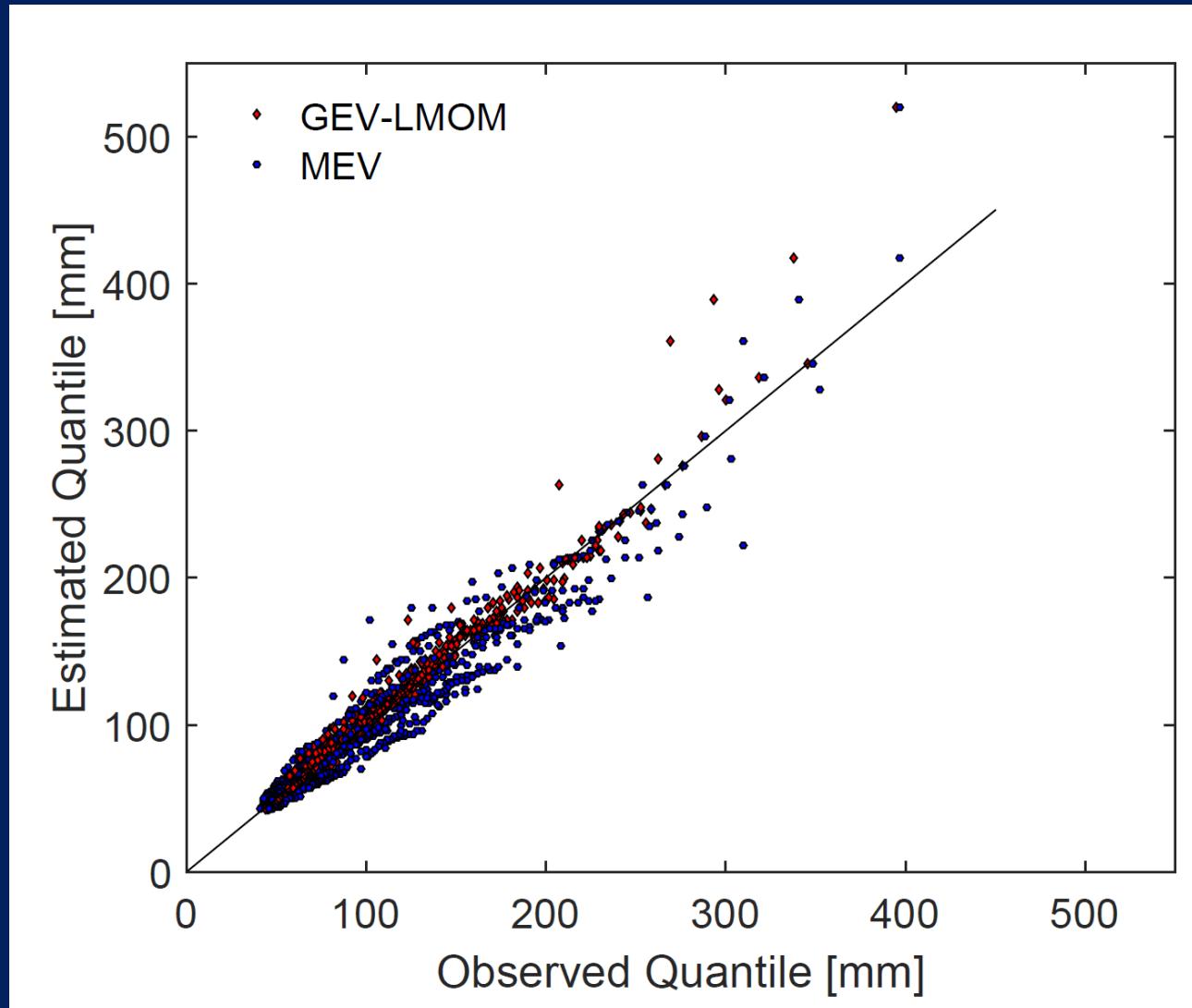
Relative estimation error as a function of $\text{Tr}/(\text{sample size})$
when training and testing data are separate

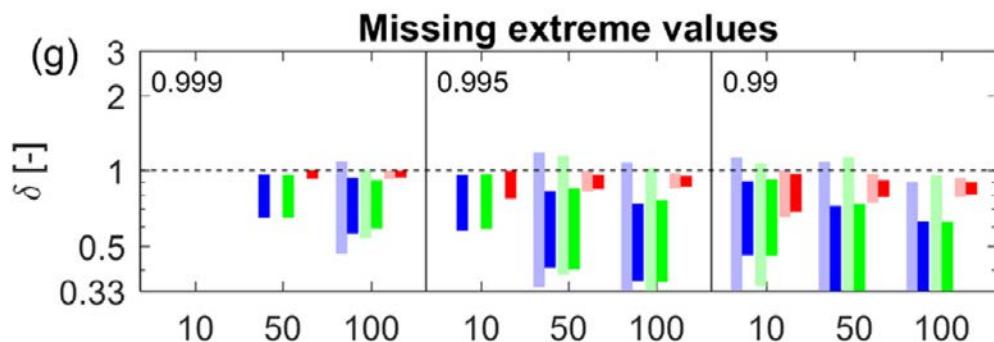
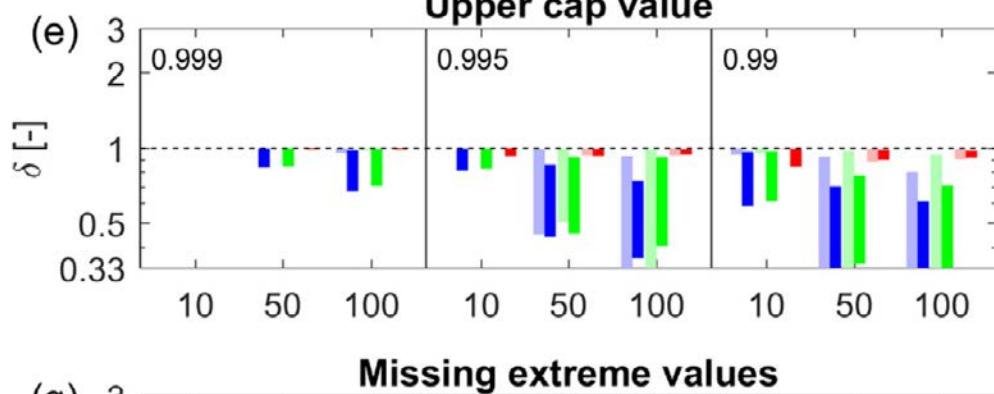
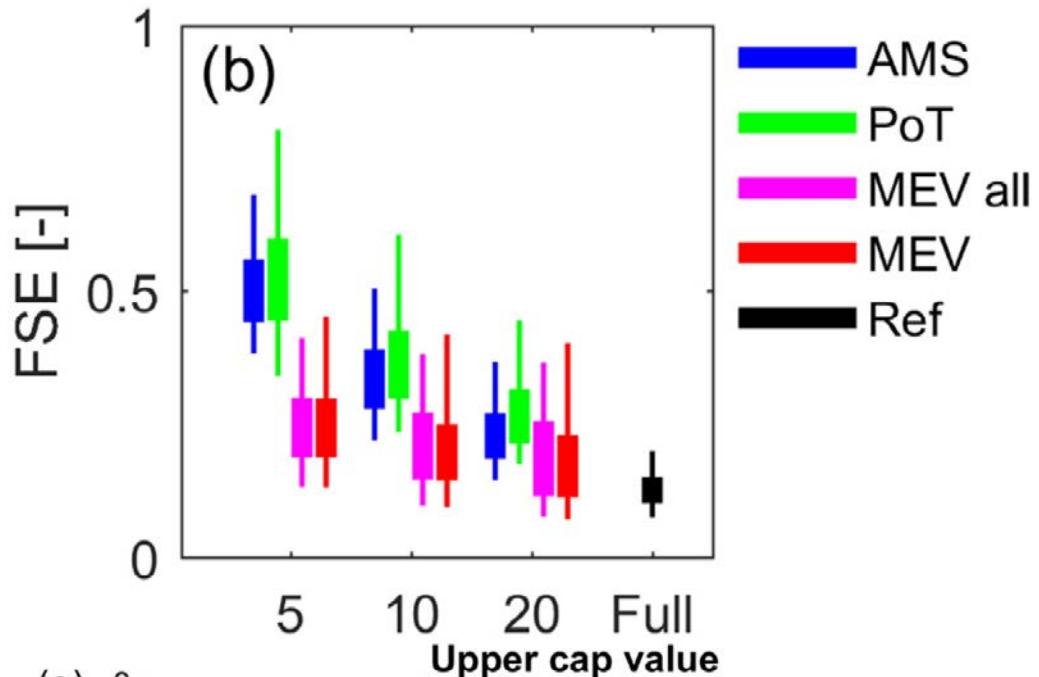
MEV vs. GEV-LMOM





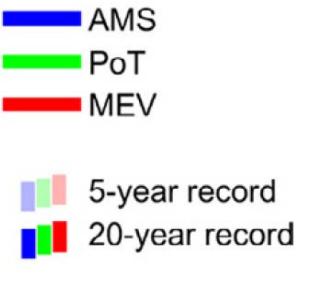
An interesting observation: GEV performs better if calibration data=testing data





Hourly rainfall (Marra et al.,
AWR 2018):

- MEV uncertainty
 $\sim 50\% < \text{GEV}$
- MEV less sensitive than GEV to errors in data (saturation at high values, missing largest values, etc.)

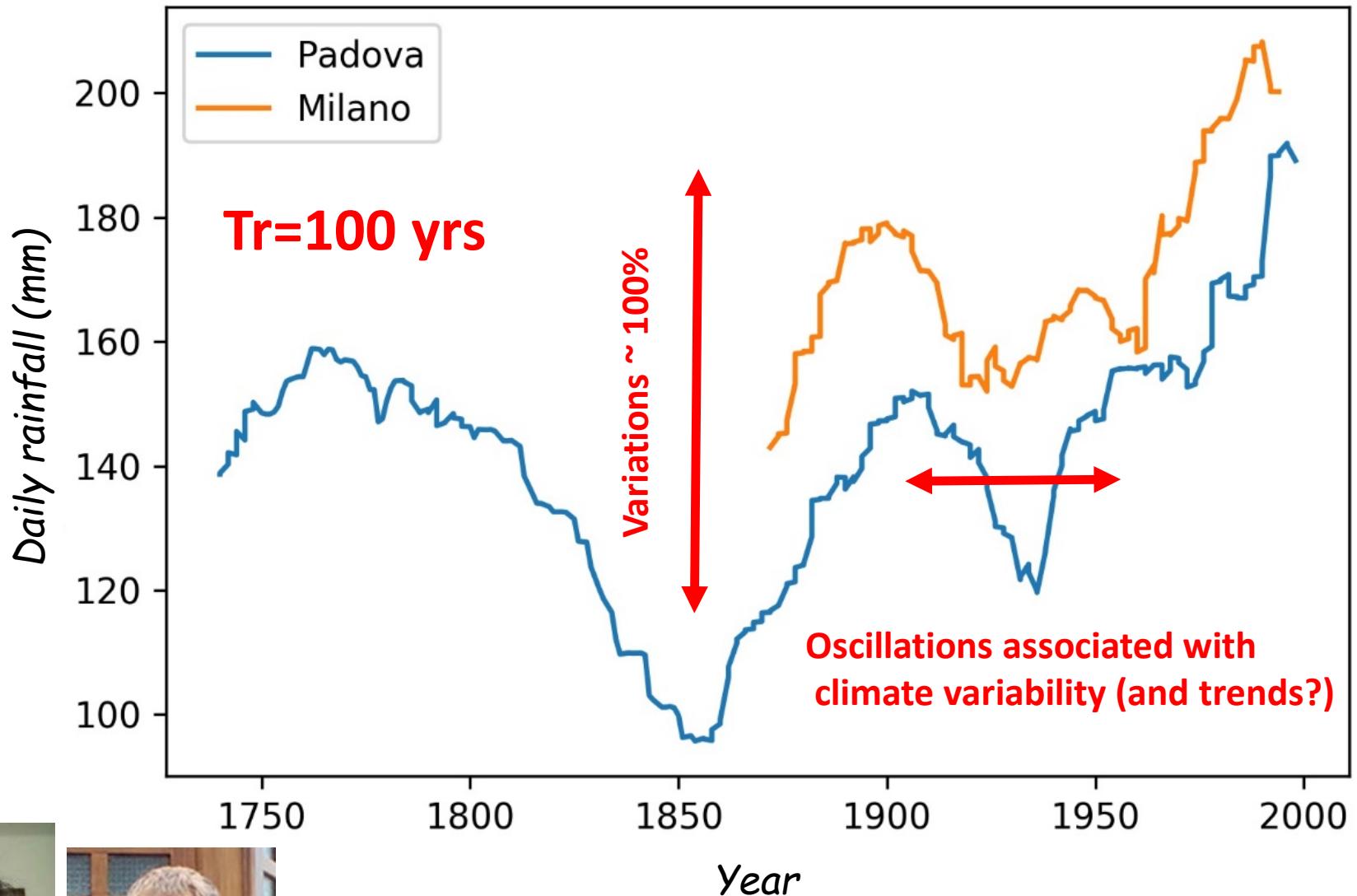


Conclusions (1/3)

MEV ouperforms classical EV distributions providing:

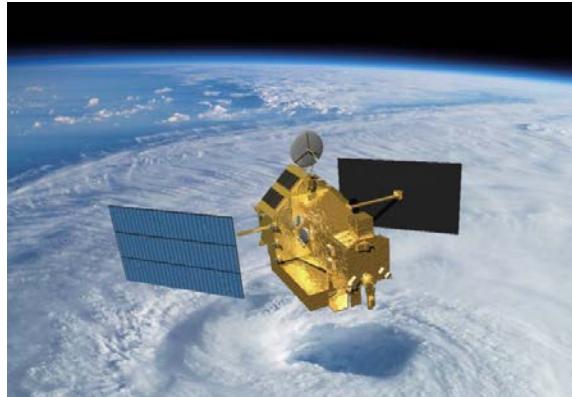
1. Reduced uncertainty (50%) for $\text{Tr} >$ sample size
2. Better use of the available daily data
3. Removal of asymptotic/Poisson hypotheses

Slight detour: return time under climate fluctuations

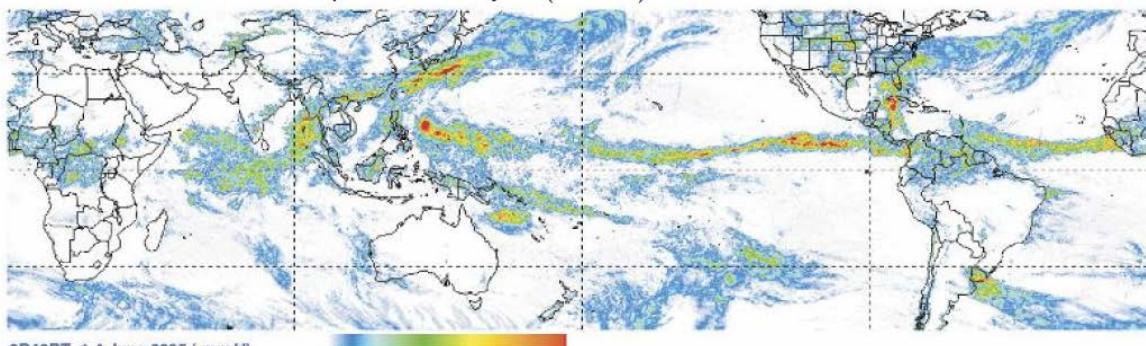


(Marani and Zanetti, WRR 2015;
Zorzetto et al., in preparation)

TRMM Multi satellite Precipitation Analysis (TMPA)



(source: NASA)



(Huffman et al, 2007)

3B42 V7 TMPA: research quality Combination of multiple sensors, chiefly microwave and infrared.

Fairly coarse $0.25^\circ \times 0.25^\circ$ spatial resolution

Up to 3 hours temporal resolution

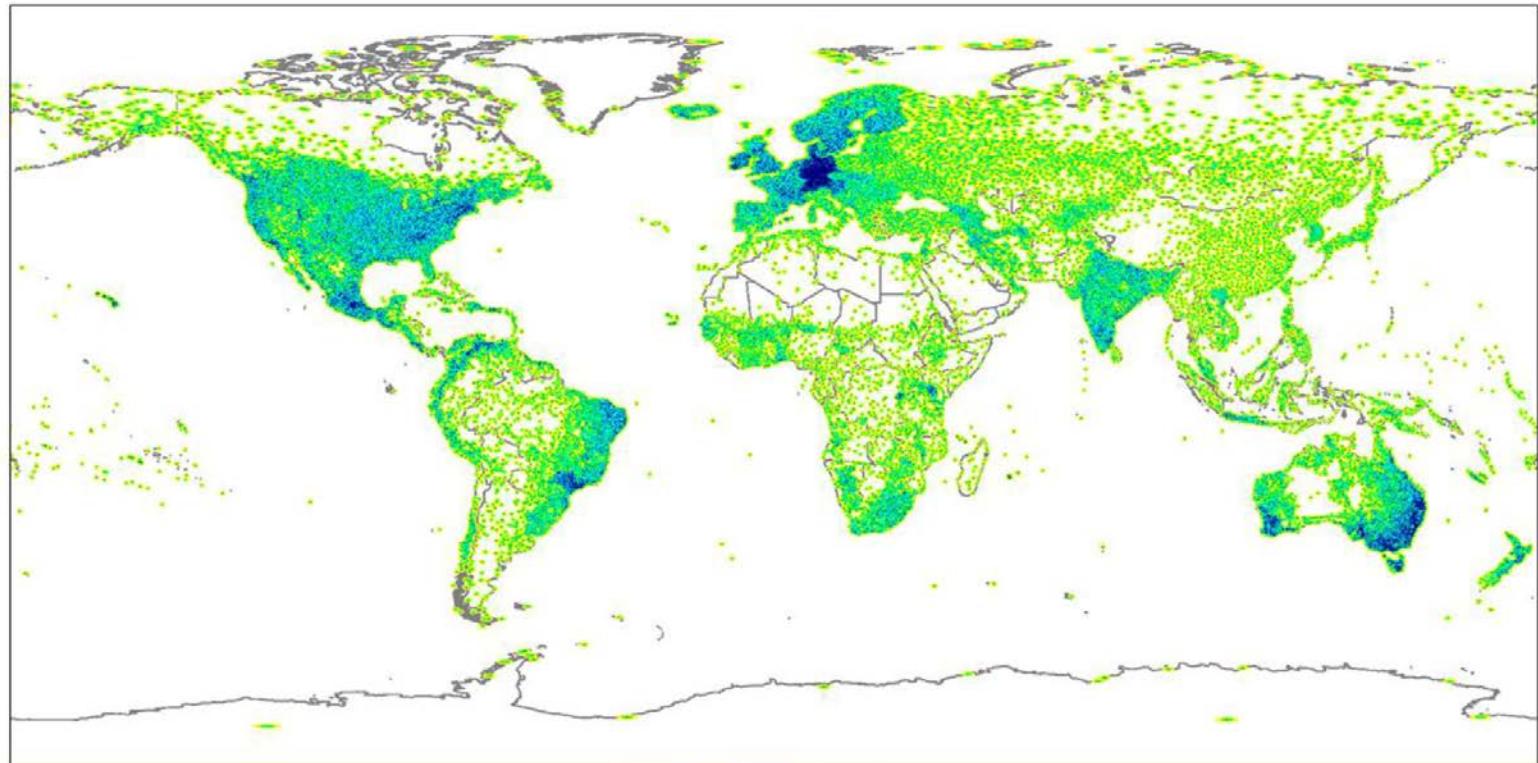
Useful for Inference on rainfall pdfs and extremes?

- QPE errors
- Short record length

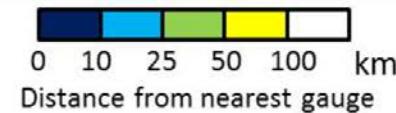


(Enrico Zorzetto,
under NASA Fellowship NESSF 17-EARTH17F-270)

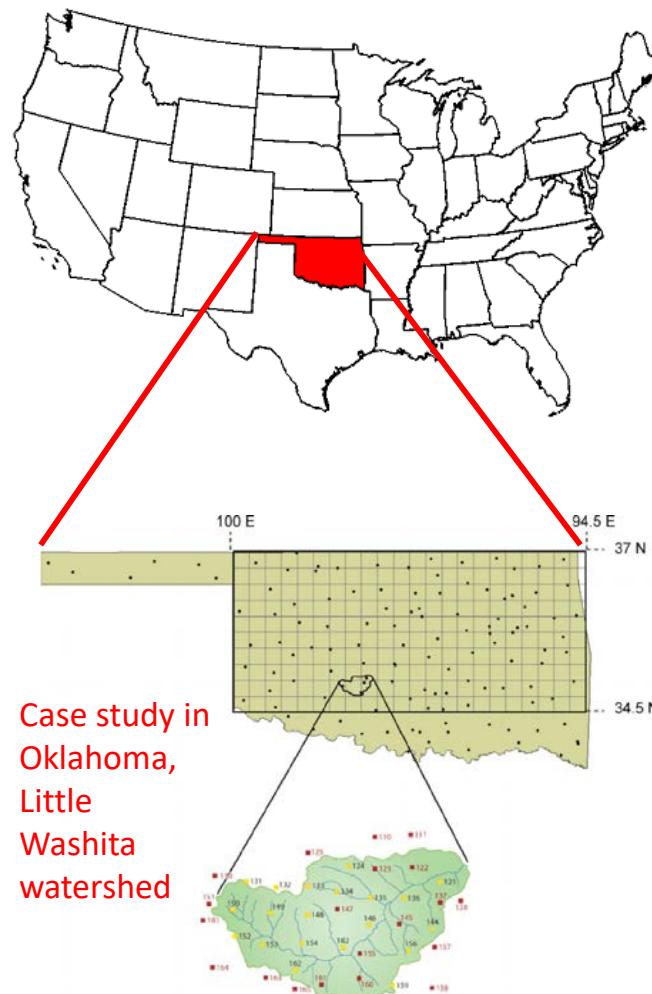
Global Rainfall Extremes: Gaps in global observational networks



But fewer stations are long enough for
Extreme rainfall frequency analysis

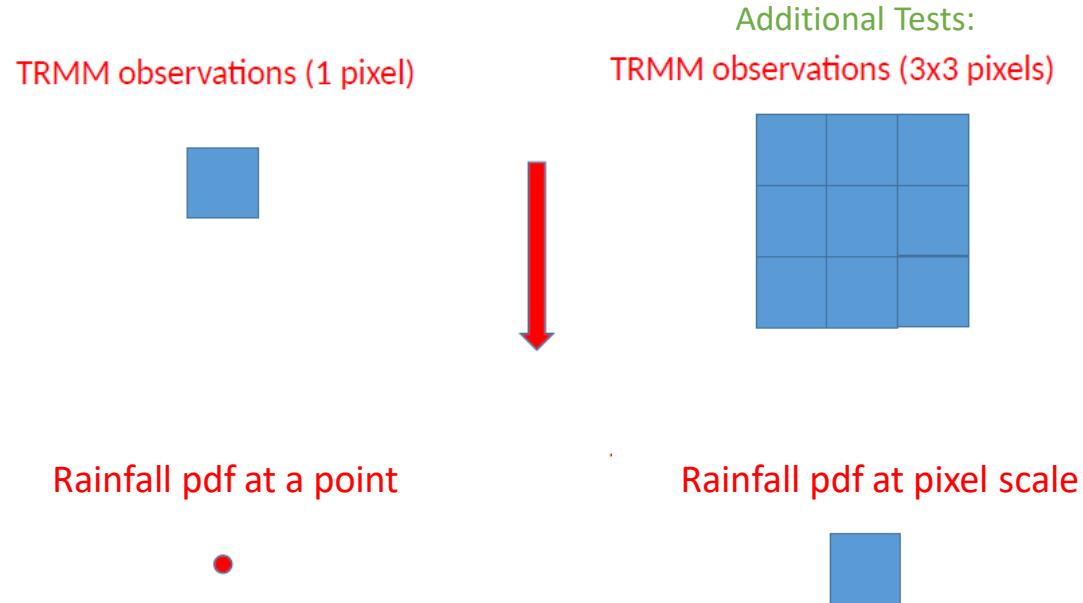


Kidd et al., BAMS, 2016



How to validate QPE rainfall pdfs using point measurements?

Parameters of rainfall pdf downscaled using closed-form relations from space-time stochastic rainfall models



A Metastatistical Extreme Value distribution (MEV)

The MEV block-maxima distribution:

$$\zeta(x) = \sum_{n=1}^{\infty} \int_{\Omega_{\vec{\Theta}}} F(x; \vec{\theta})^n g(n, \vec{\theta}) d\vec{\theta}$$

$G(n, \theta)$ = joint prob distrib. of the parameters.

Can be approximated using sample averages:

$$\zeta(x) \cong \frac{1}{T} \sum_{j=1}^T F(x; \vec{\theta}_j)^{n_j}$$

T = # sub-periods over which n and θ are estimated



A downscaling approach for TMPA gridded data

We transform across scales the following quantities

The mean is conserved across scales

$$\mu_0 = \mu_L.$$

The variance increase of a factor γ

Effect of spatial correlation

$$\sigma_0^2 = \frac{1}{\gamma} \sigma_L^2, \quad \gamma = \frac{\sigma_{c_L}^2}{\sigma_{c_0}^2} = \frac{4}{L_x^2 L_y^2} \int_0^{L_x} \int_0^{L_y} (L_x - x)(L_y - y) \rho(x, y) dx dy$$

[Vanmarcke, 2010]

Variation of the yearly number of daily events/year

$$\gamma_s(L) = \frac{N_L}{N_0}$$

[Zorzetto and Marani, 2018]

Can we infer these quantity from TMPA alone?

Precipitation spatial correlation at the point scale

$$\Delta(L_{x,k}, L_{y,l}) = 4 \int_0^{L_{x,k}} \int_0^{L_{y,l}} (L_{x,k} - s_1) (L_{y,l} - s_2) \rho(s_1, s_2) ds_1 ds_2$$

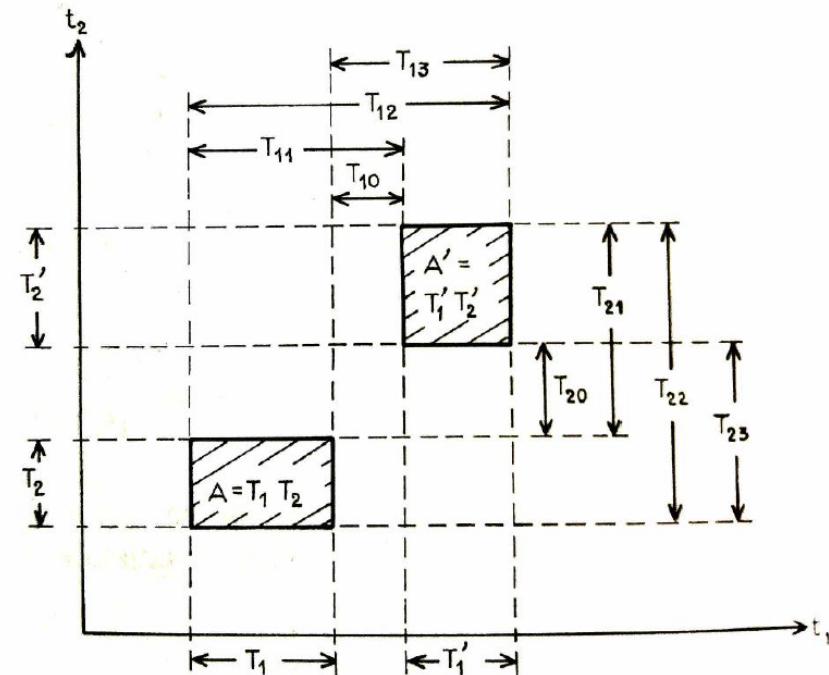
Correlation of local averages

$$\text{Cov}[h_L, h_{L'}] = \frac{\sigma_0^2}{4(L_x L_y)^2} \sum_{k=0}^3 \sum_{l=0}^3 (-1)^k (-1)^l \Delta(L_{x,k}, L_{y,l})$$

Distances $L_{x,k}$ and $L_{y,l}$ with $k, l = 0, 1, 2, 3$ encode the relative position of the two pixels. Assume correlation function:

$$\rho(s) = \begin{cases} e^{-\frac{\alpha s}{\epsilon}} & s < \epsilon \\ \left(\frac{\epsilon}{es}\right)^\alpha & s \geq \epsilon \end{cases} \quad s = \sqrt{x^2 + y^2}$$

From Marani, 2003



From Vanmarcke, 2010

Precipitation spatial correlation at the point scale

$$\Delta(L_{x,k}, L_{y,l}) = 4 \int_0^{L_{x,k}} \int_0^{L_{y,l}} (L_{x,k} - s_1) (L_{y,l} - s_2) \rho(s_1, s_2) ds_1 ds_2$$

Correlation of local averages

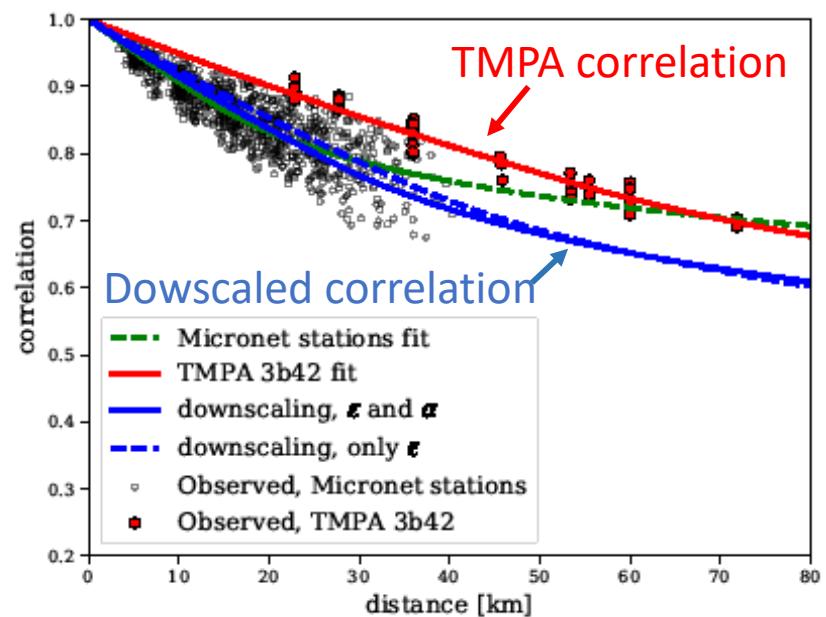
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Distances $L_{x,k}$ and $L_{y,k}$ with $k, l = 0, 1, 2, 3$ encode the relative position of the two pixels. Assume correlation function:

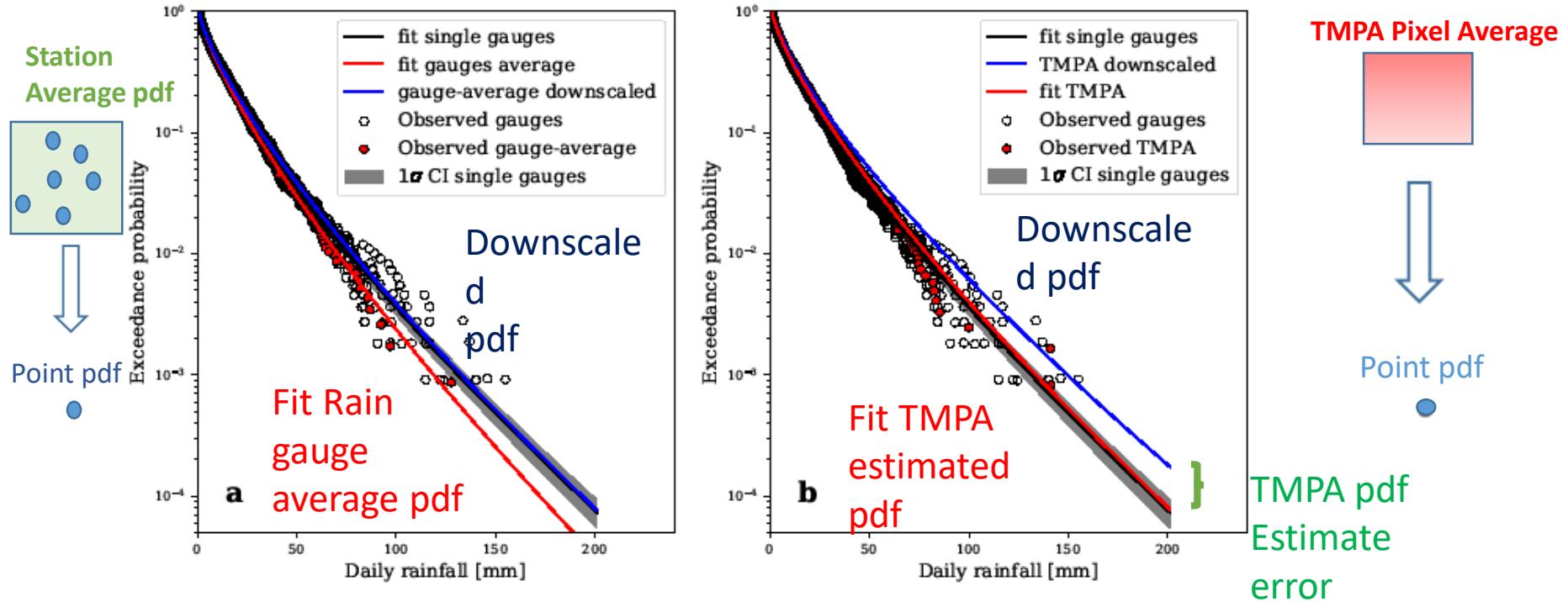
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From Marani, 2003

Assumed Correlation function shape



Downscaling of the pdf of ordinary daily rainfall



(Zorzetto and Marani, WRR, 2018)



A downscaling approach for TMPA gridded data

We transform across scales the following quantities

The mean is conserved across scales

$$\mu_0 = \mu_L.$$

The variance increase of a factor γ

$$\sigma_0^2 = \frac{1}{\gamma} \sigma_L^2, \quad \gamma = \frac{\sigma_{c_L}^2}{\sigma_{c_0}^2} = \frac{4}{L_x^2 L_y^2} \int_0^{L_x} \int_0^{L_y} (L_x - x)(L_y - y) \rho(x, y) dx dy$$

[Vanmarcke, 2010]

Variation of the yearly number of daily events/year

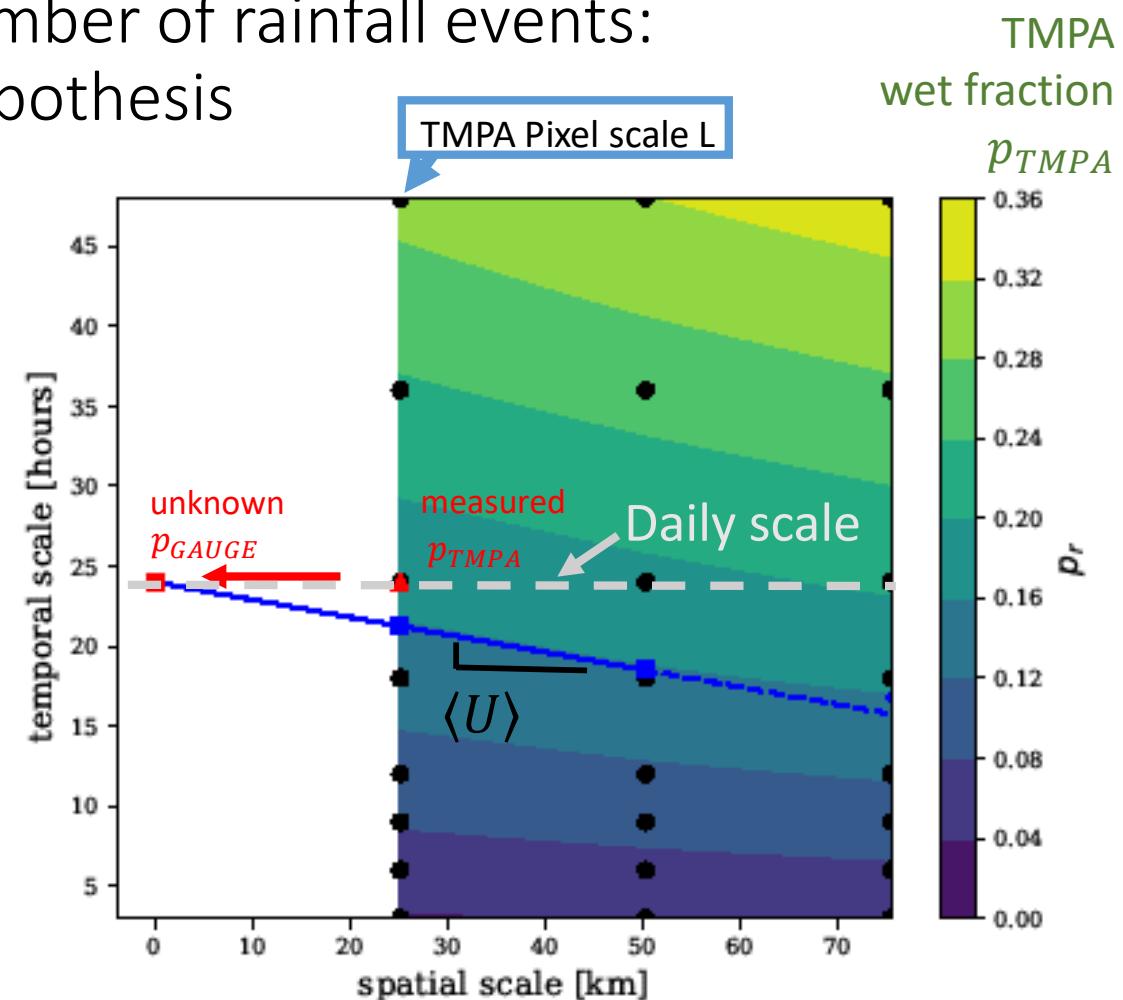
$$\gamma_s(L) = \frac{N_L}{N_0} \quad \text{Effect of spatial intermittency}$$

[Zorzetto and Marani, 2018] Can we infer these quantity from TMPA alone?

Downscaling of the Number of rainfall events: “frozen turbulence” hypothesis

Effect of spatial intermittency

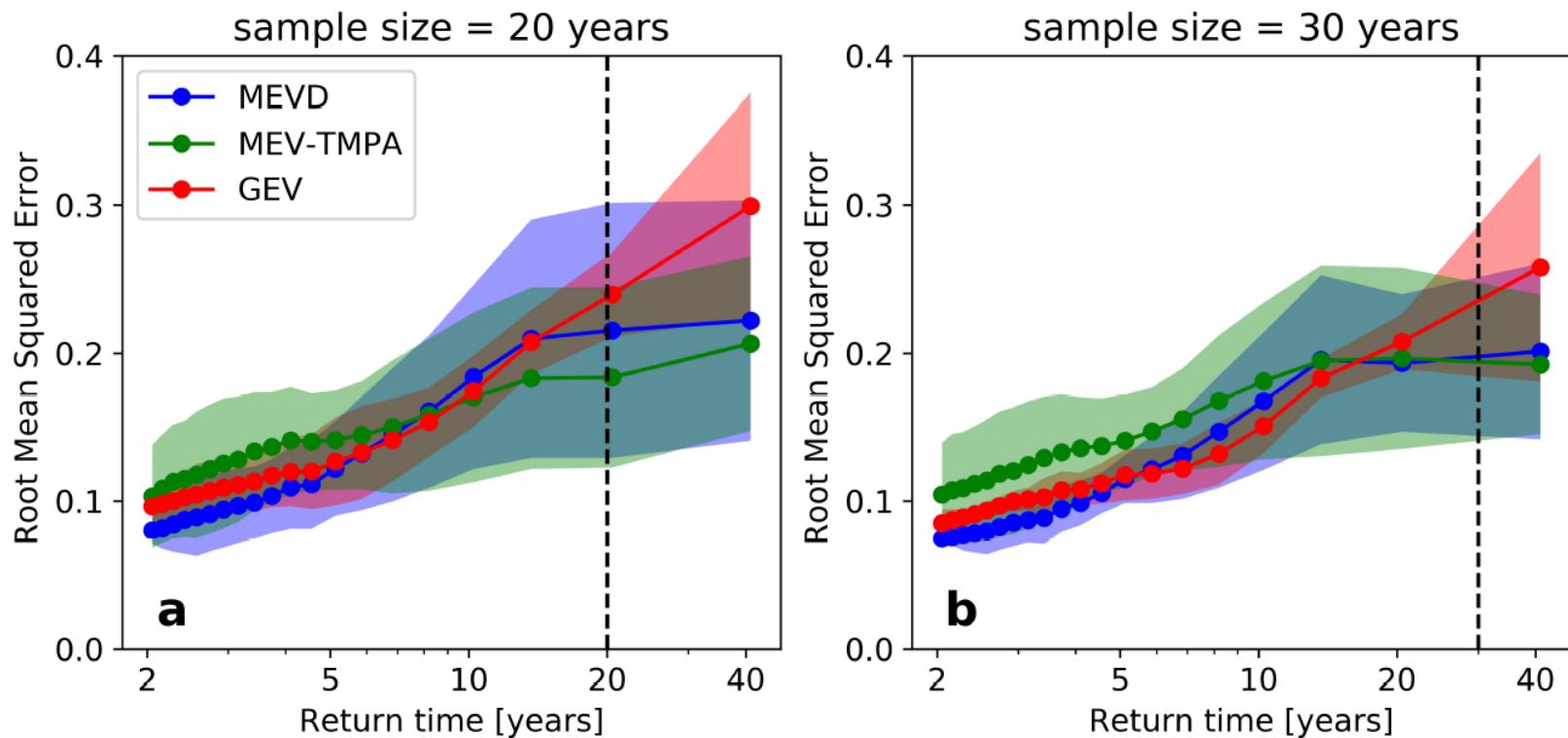
$$\gamma_s(L) = \frac{p_{TMPA}(L)}{p_{GAUGE}}$$



(Zorzetto and Marani, WRR, 2018)

Cross-Validation Analysis

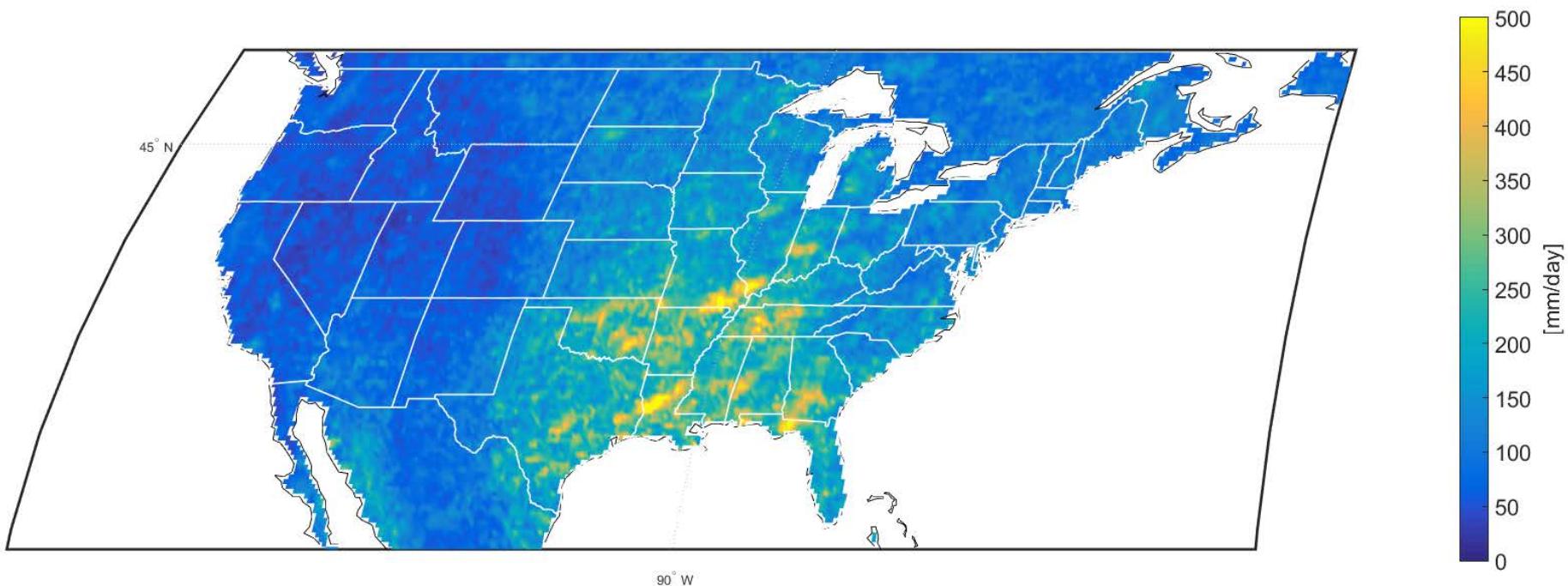
$$\text{Root Mean Square Error } R(s, Tr) = \left(\frac{1}{n_g} \sum_{i=1}^{n_g} \left[\frac{\hat{h}_i(Tr) - h_i^{obs}(s_s, Tr)}{h_i^{obs}(Tr)} \right]^2 \right)^{1/2}$$



Error of the downscaling technique (green) vs ground observations (red, blue).

From TMPA satellite estimates alone

Continental USA - GEV, 100-year daily rainfall event
from all TRMM-TMPA observations (3B42 V7 - 17 yrs)

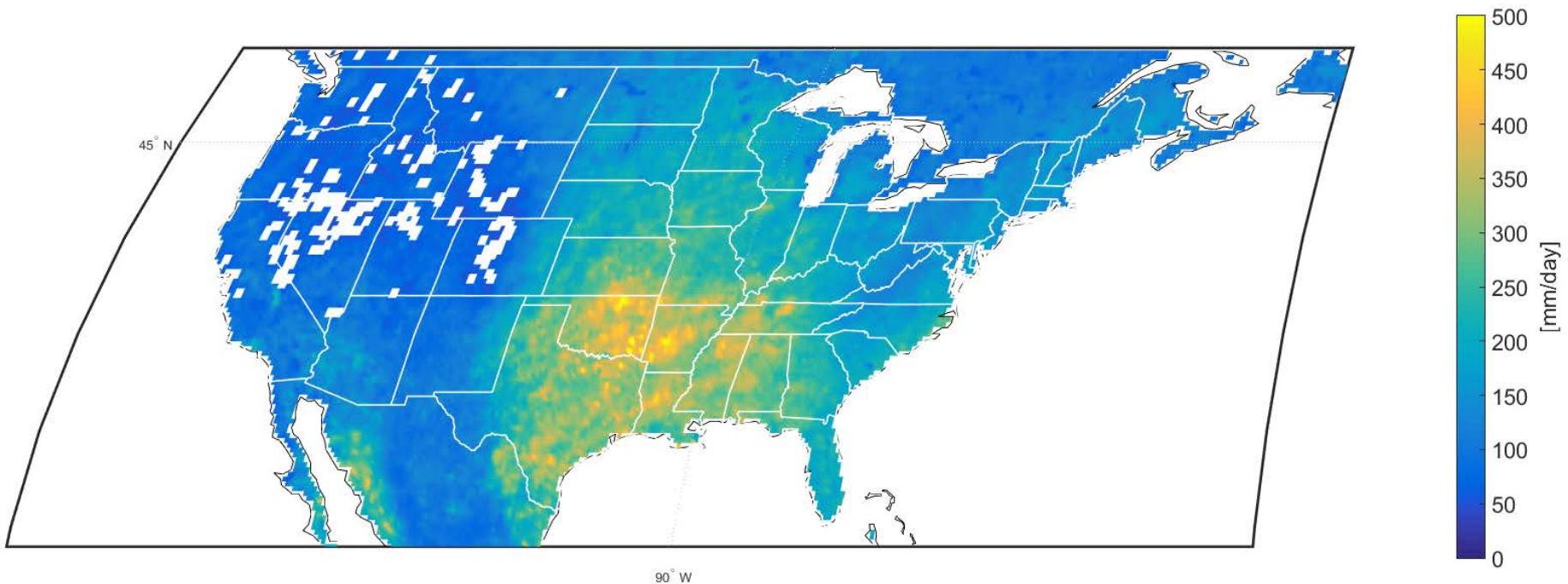


Daily TRMM 3B42 V7 data:
Quasi global (+50°N, -50°S);
0.25°x0.25° spatial resolution

Zorzetto and Marani,
in preparation, 2019

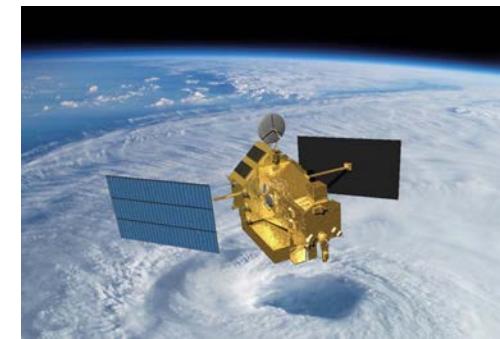


Continental USA – MEV, 100-yr daily rainfall event from all TRMM-TMPA observations (3B42 V7 - 17 yrs)



Zorzetto and Marani,
In preparation, 2019

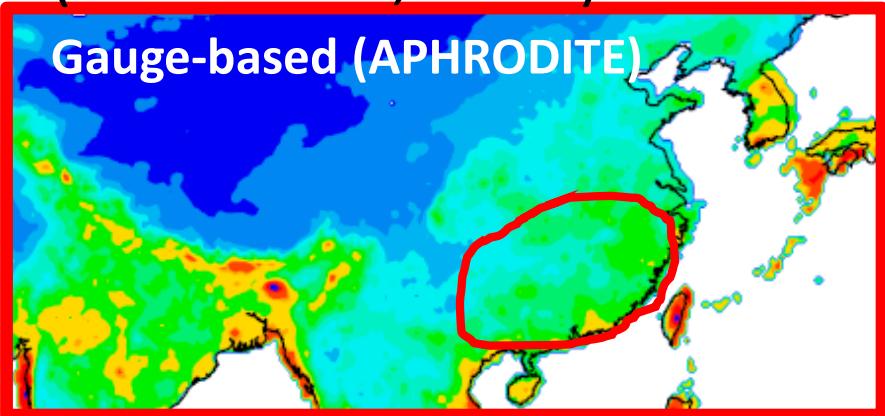
Daily TRMM 3B42 V7 data:
Quasi global (+50°N, -50°S);
0.25°x0.25° spatial resolution



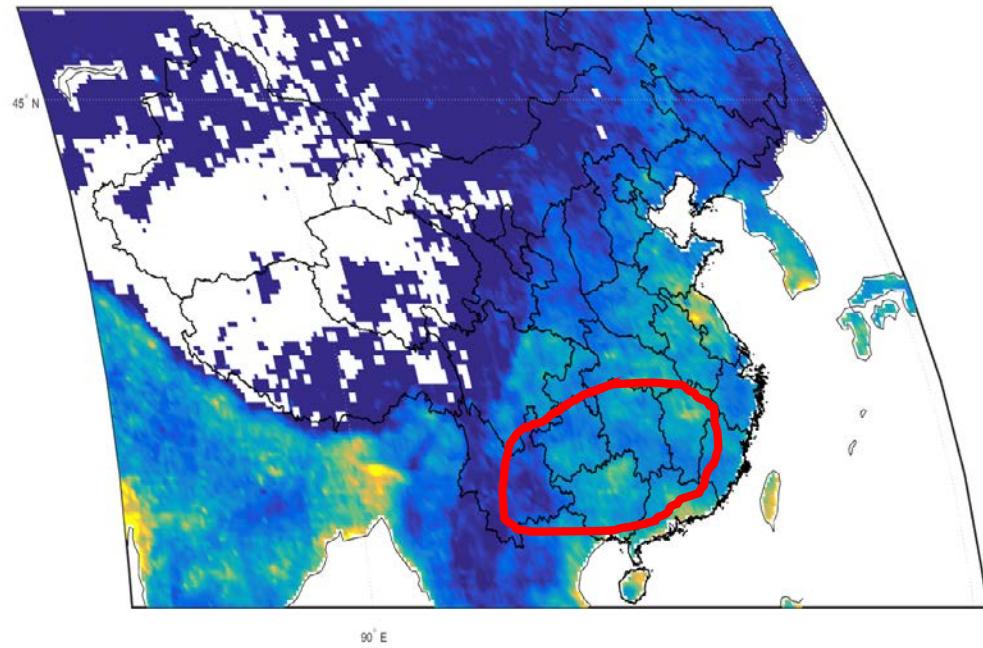
Comparisons of extreme event ($Tr=100$ yrs) distribution patterns

(Zhou et al., 2015)

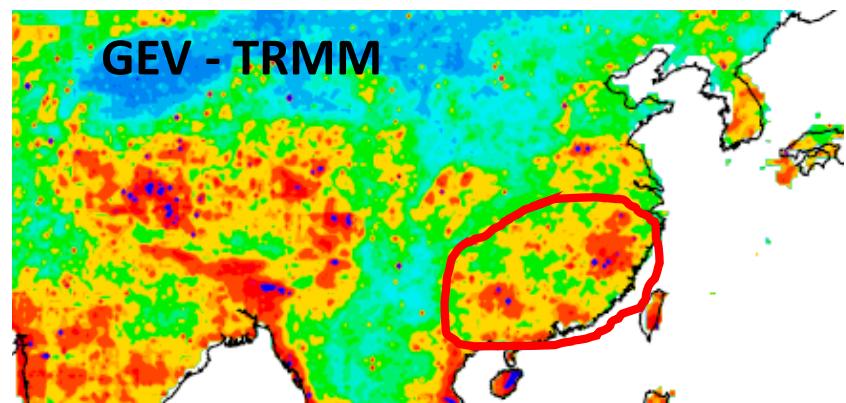
Gauge-based (APHRODITE)



MEV-TRMM (this work)

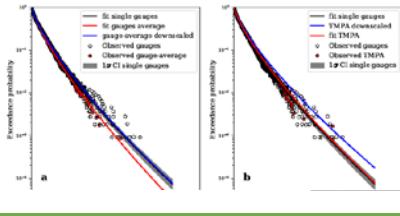


GEV - TRMM



Zorzetto and Marani, in preparation, 2019

Conclusions (2/3):



- Downscaling yields rainfall pdf at a point. **Validation of TMPA feasible even with sparse ground observations.**
- We can reconstruct **rainfall statistics at a point in space, uniquely based on areal-average satellite precipitation estimates**
- ongoing work: investigation of method performance over **inhomogeneous and complex terrain**



(Source:NASA)

U.S. HURRICANE LANDFALLS

2017

HARVEY (CAT. 4)
ROCKPORT, TEXAS
AUG. 25

NATE (CAT. 1)
LA/MS
OCT. 7-8

IRMA (CAT. 4)
SOUTH FLORIDA
SEP. 10

MARIA (CAT. 4)
PUERTO RICO
SEP. 20

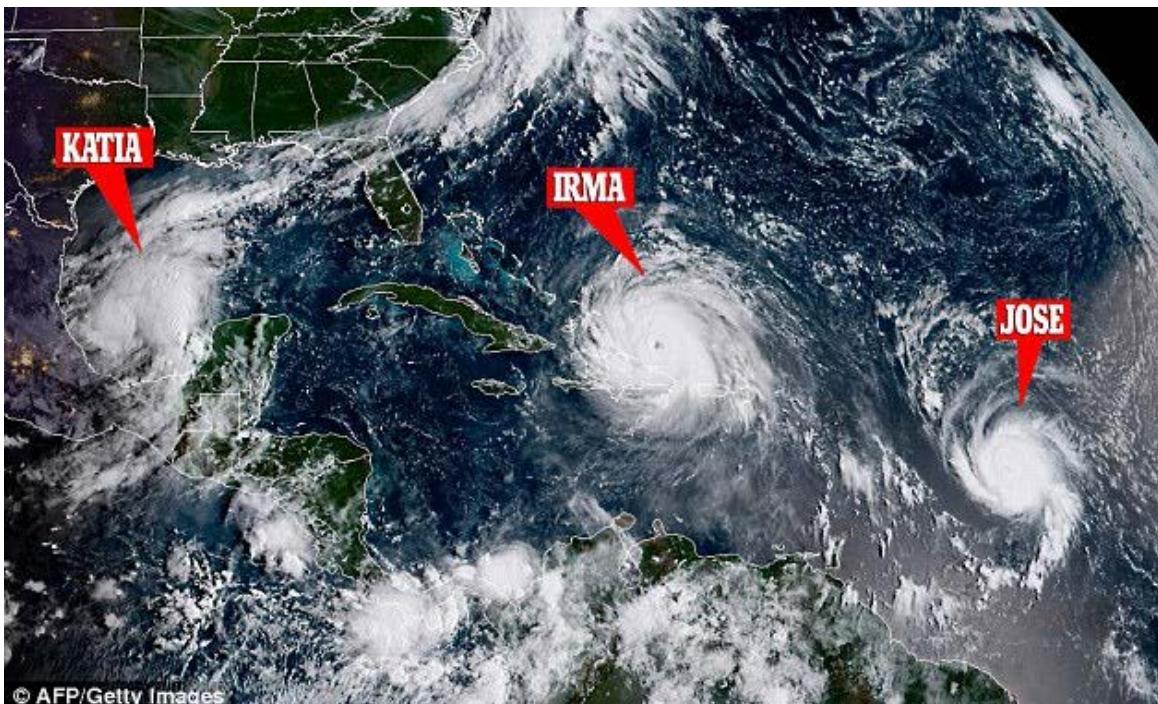
The Weather Channel

Extreme Atlantic hurricanes

KATIA

IRMA

JOSE

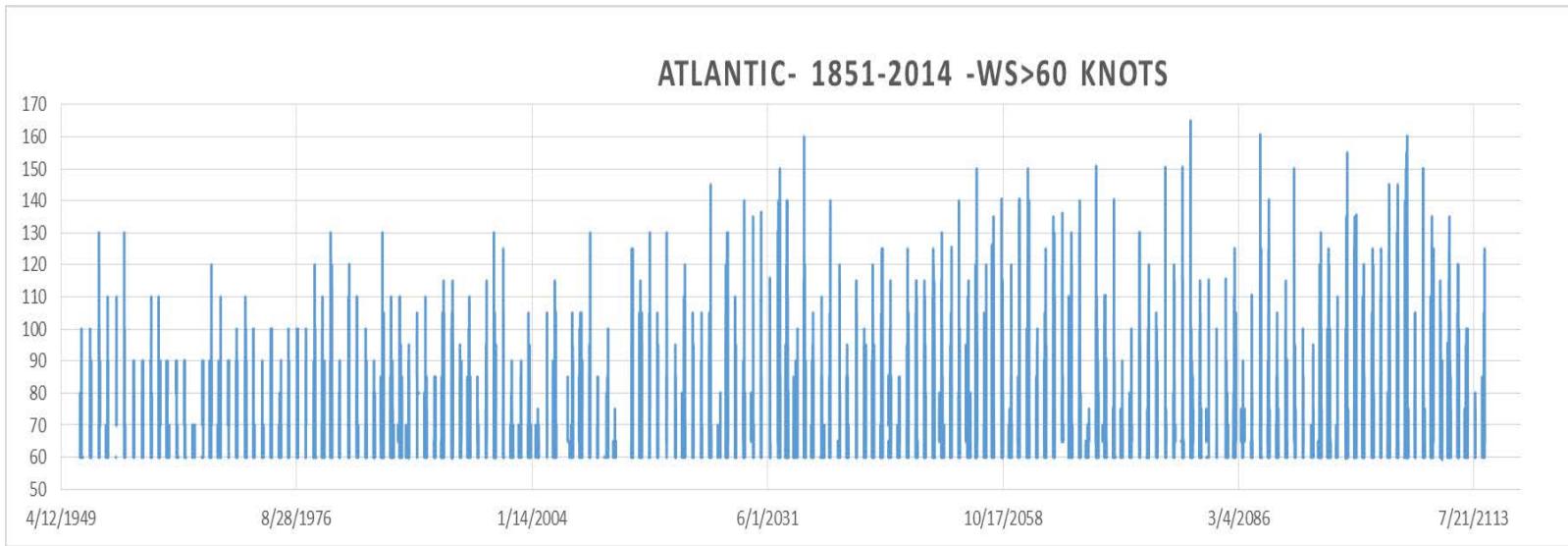


Reza Hosseini,
Polytechnic
Milan



Marco Scaioni,
Polytechnic
Milan





Category	Wind speeds
Five	≥70 m/s, ≥137 knots ≥157 mph, ≥252 km/h
Four	58–70 m/s, 113–136 knots 130–156 mph, 209–251 km/h
Three	50–58 m/s, 96–112 knots 111–129 mph, 178–208 km/h
Two	43–49 m/s, 83–95 knots 96–110 mph, 154–177 km/h
One	33–42 m/s, 64–82 knots 74–95 mph, 119–153 km/h



«Ordinary» Hurricanes: Generalize Pareto Distribution

The MEV expression is:

$$\zeta(x) = \frac{1}{T} \sum_{j=1}^T F(x; \vec{\theta}_j)^{n_j}$$

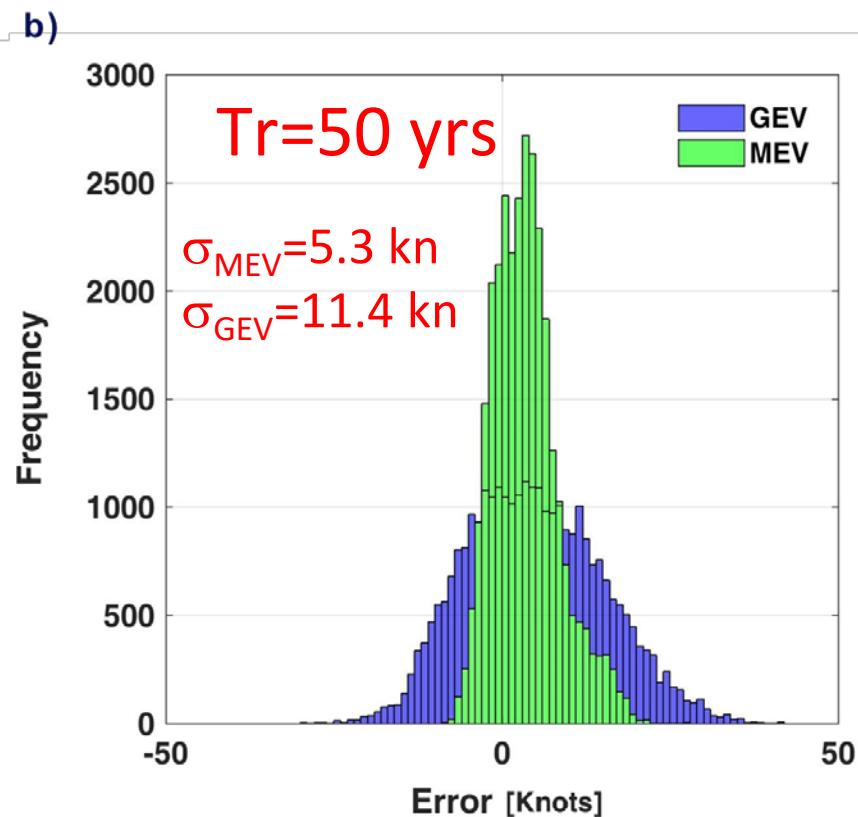
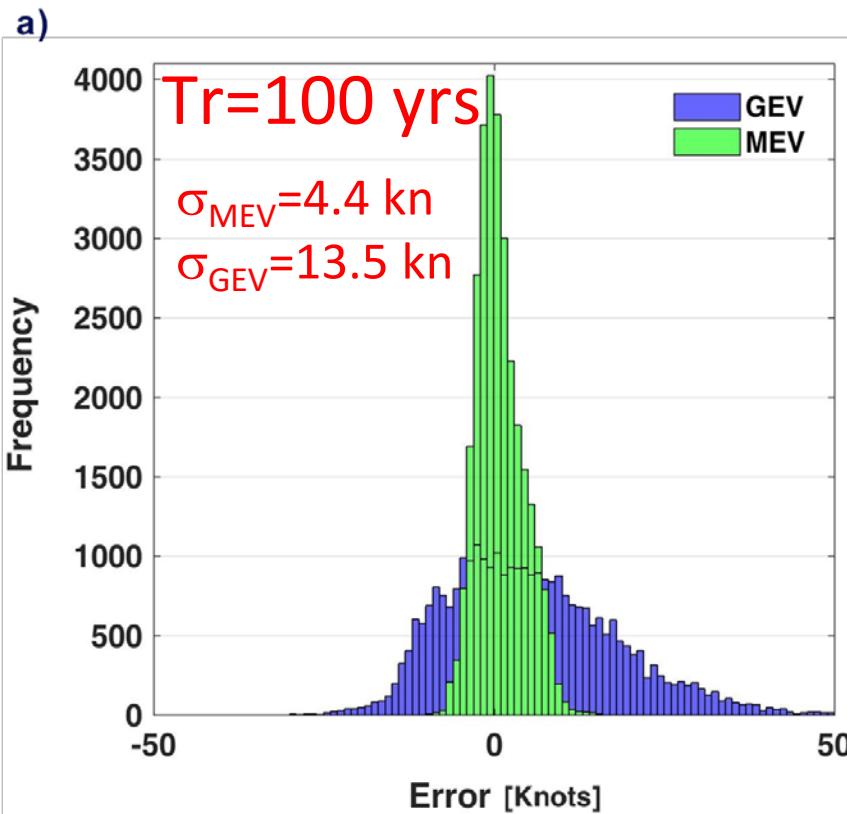
T = # sub-periods over which n
and θ are estimated

We find that the Generalized Pareto Distribution (GPD) best describes ‘ordinary’ Hurricanes:

$$\zeta(x) = \frac{1}{T} \sum_{j=1}^T \left[1 - \left(1 - k_j \frac{(x - u)}{\sigma_j} \right)^{\frac{1}{k}} \right]^{n_j}$$

Hurdat 2 analysis (cross validation with independent randomized calibration and test datasets).

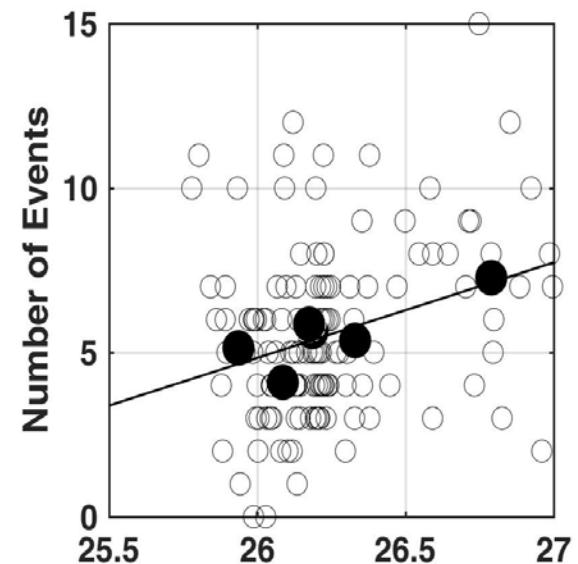
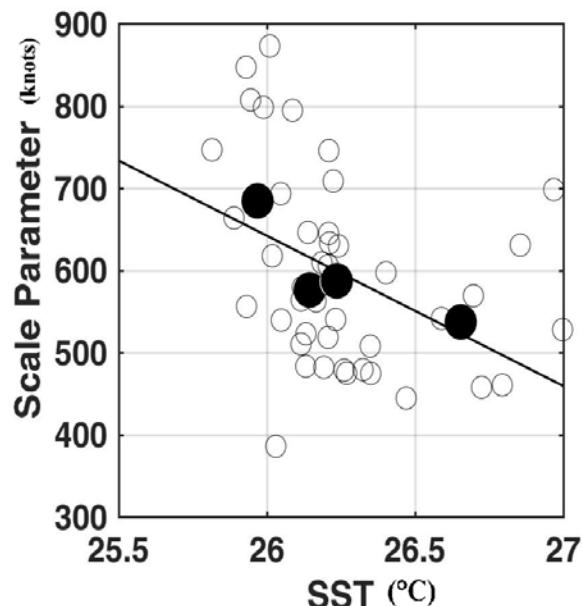
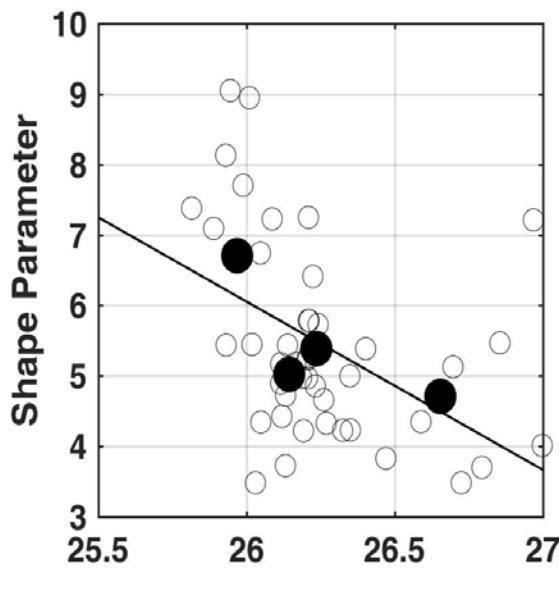
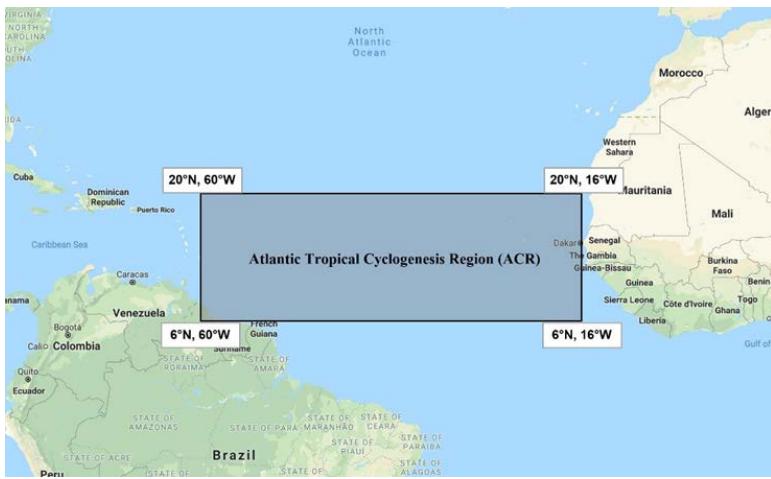
Extreme Hurricanes: MEVD (GPD) vs. traditional GEV



50-70% Improvement in estimation uncertainty

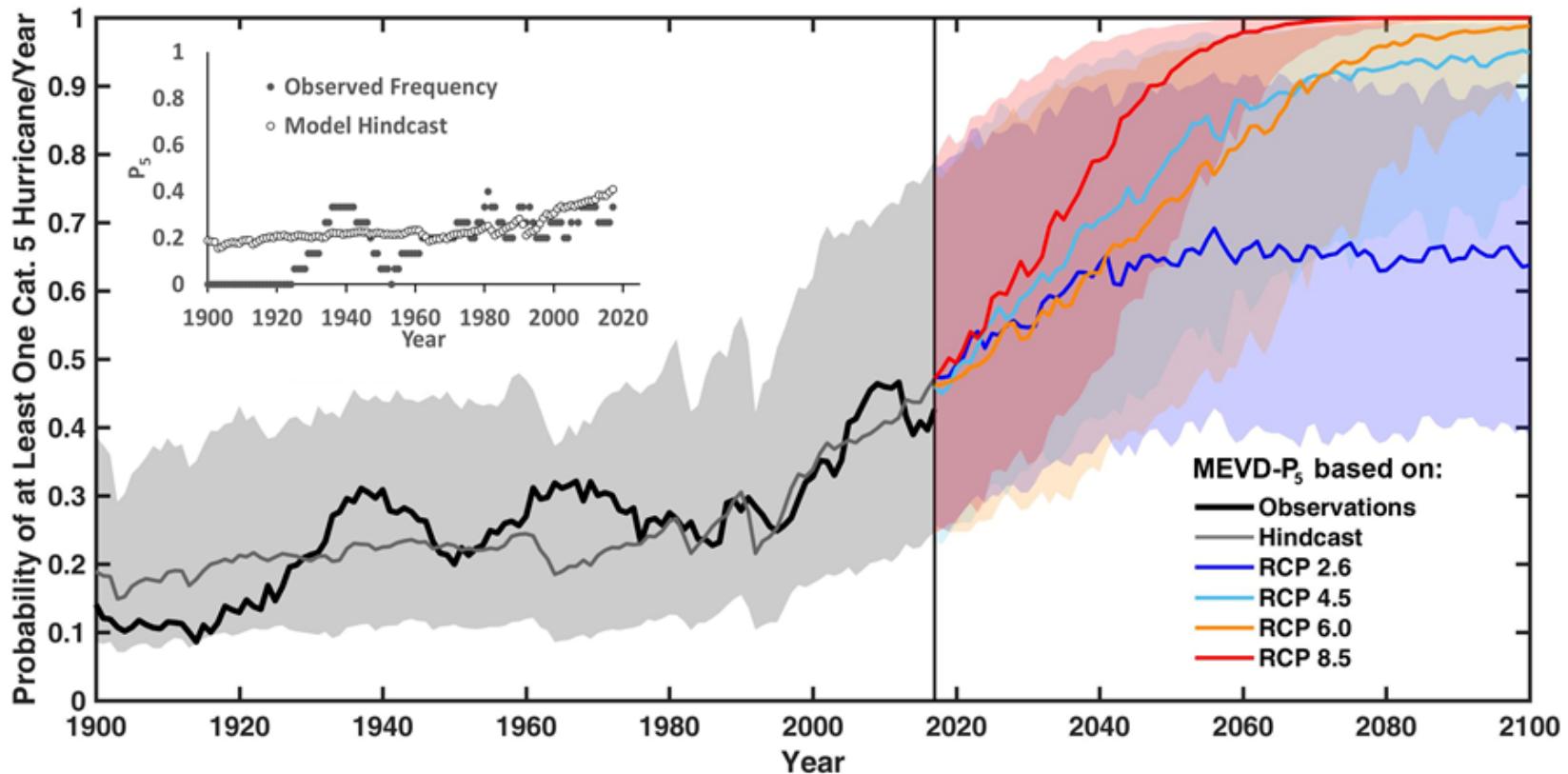
Hosseini et al, submitted, 2019

GPD parameters and # hurricanes/year as a function of Sea Surface Temperature (SST in the ACR)



Hosseini et al, submitted, 2019

Hindcast and projection of extreme-hurricane probability



Conclusions (3/3):

- Generalized Pareto Distribution of ordinary hurricane intensity;
- MEVD-GPD improves estimation uncertainty by 50%-70% in the case of atlantic hurricane intensity.
- MEVD can be successfully used to infer the effects on changes in ordinary events on the distribution of large extremes

Summary:



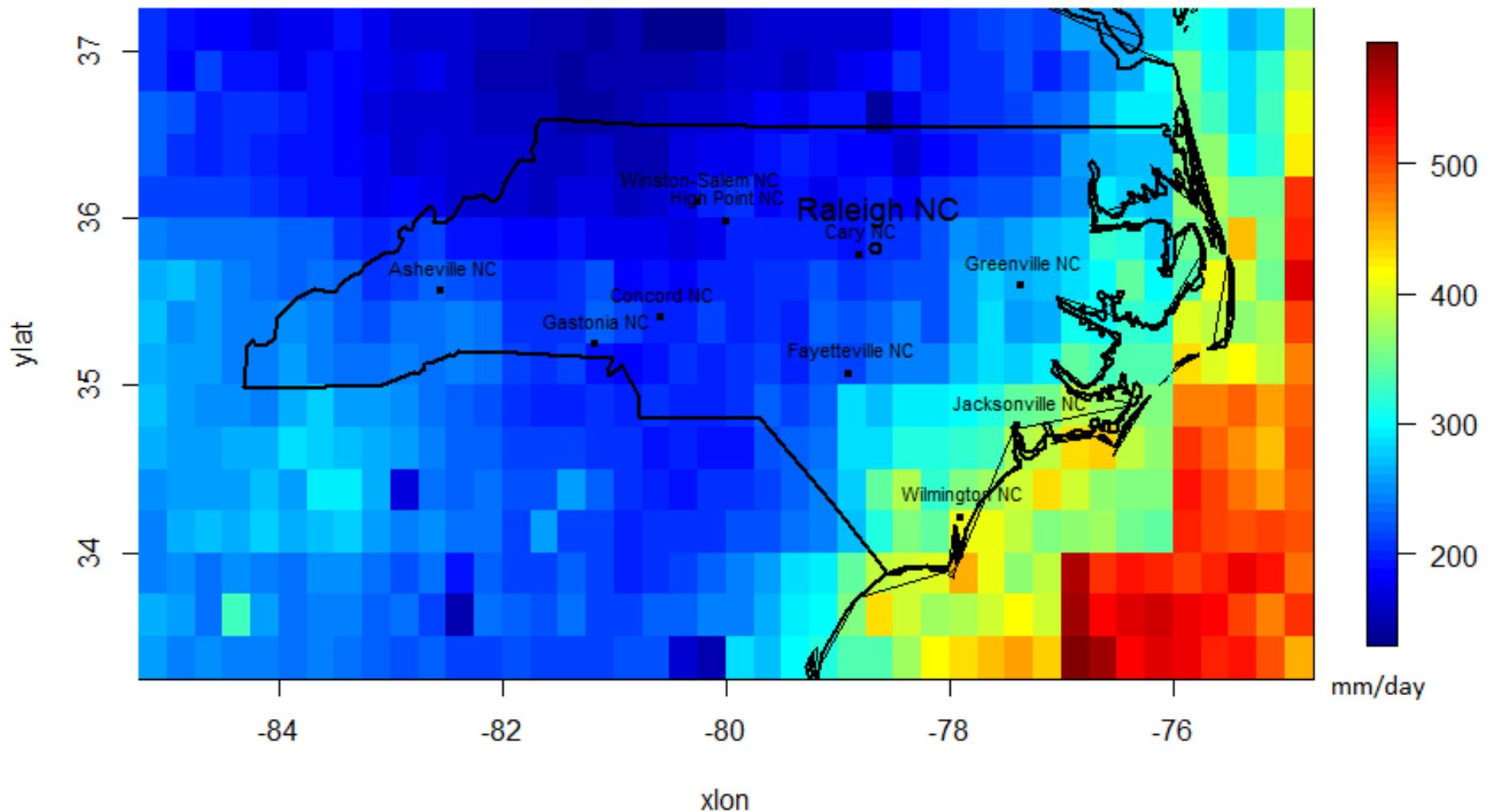
MEVD shows practical and conceptual advantages over traditional EVT (rainfall, floods, storm surges, wind)

- links extremes to “ordinary” events:
large samples, interannual
variability/trends, physical processes

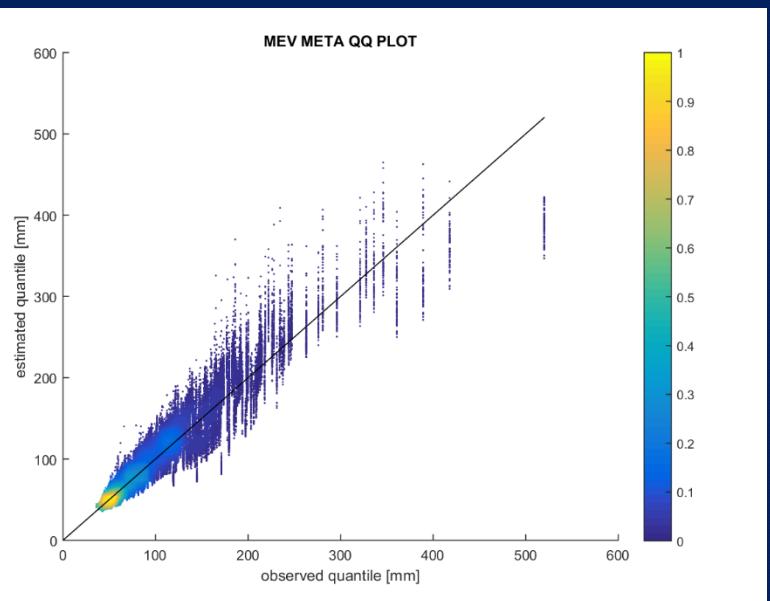
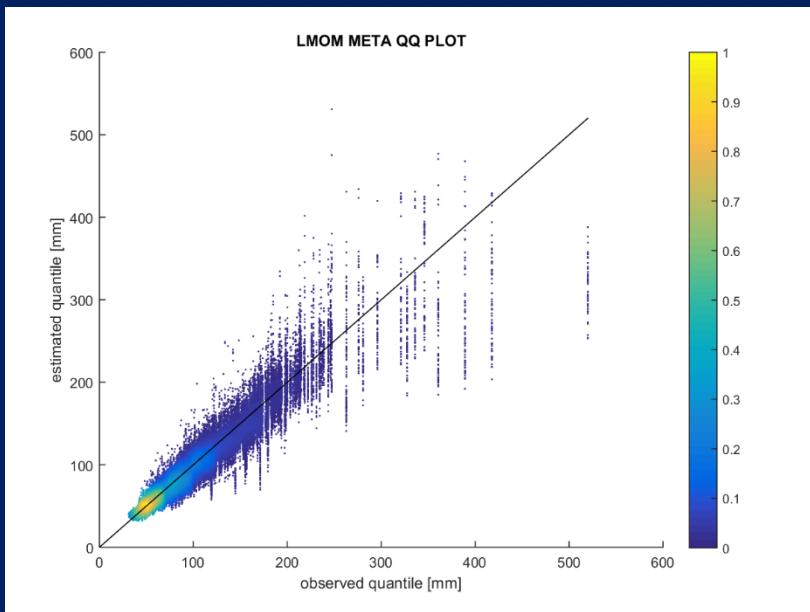
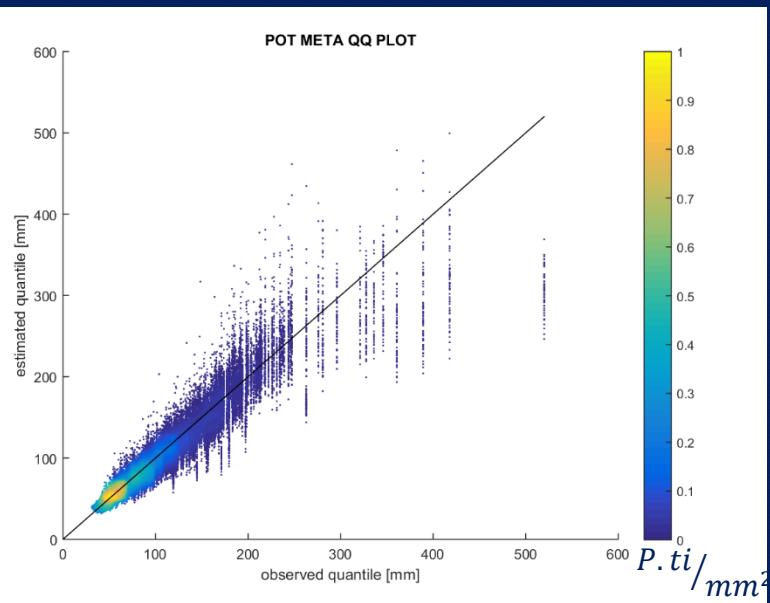


Time to move beyond limiting distributions and Poisson arrivals...

Tr=100 daily rainfall from TRMM observations (17 yrs)



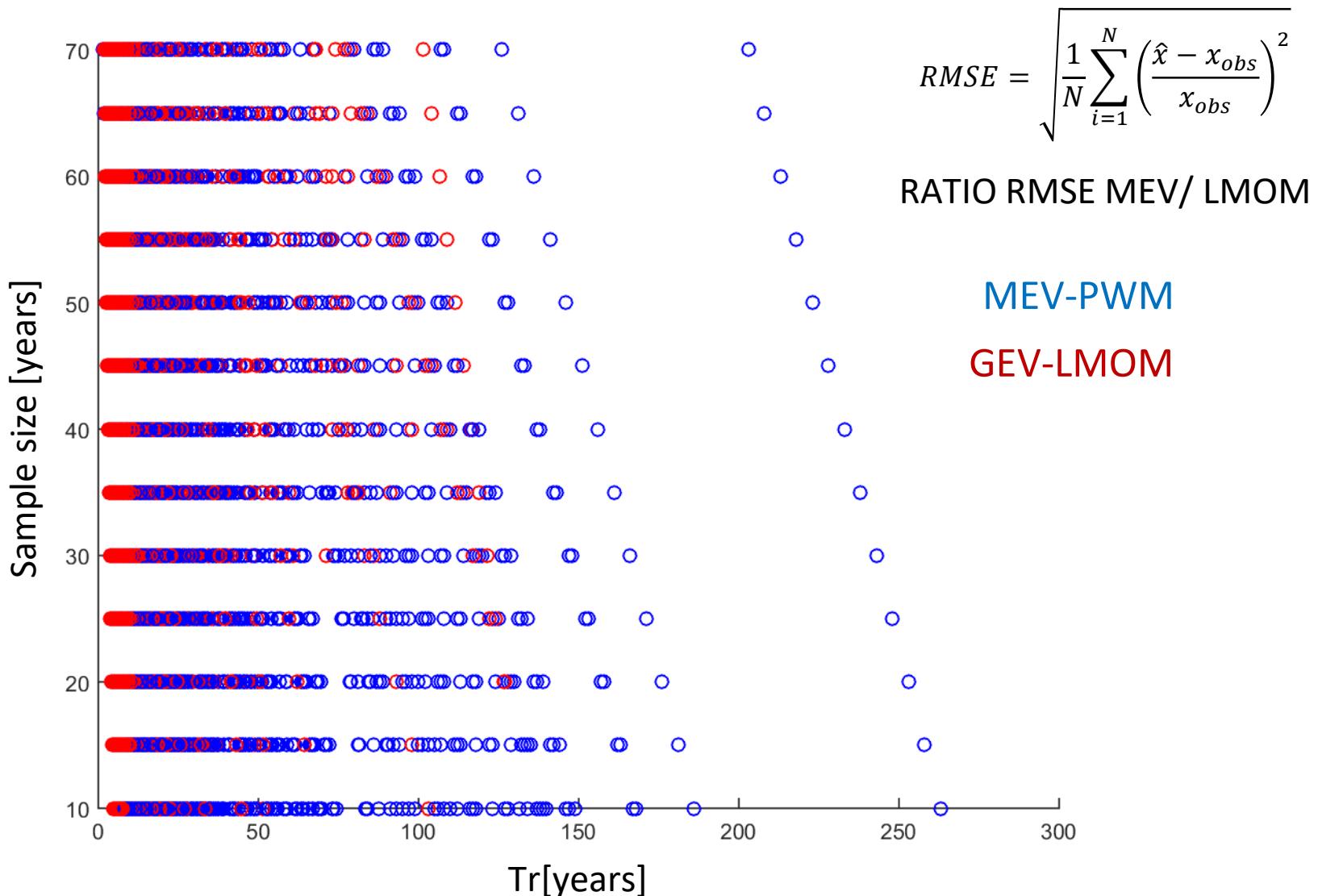
Global QQ plots



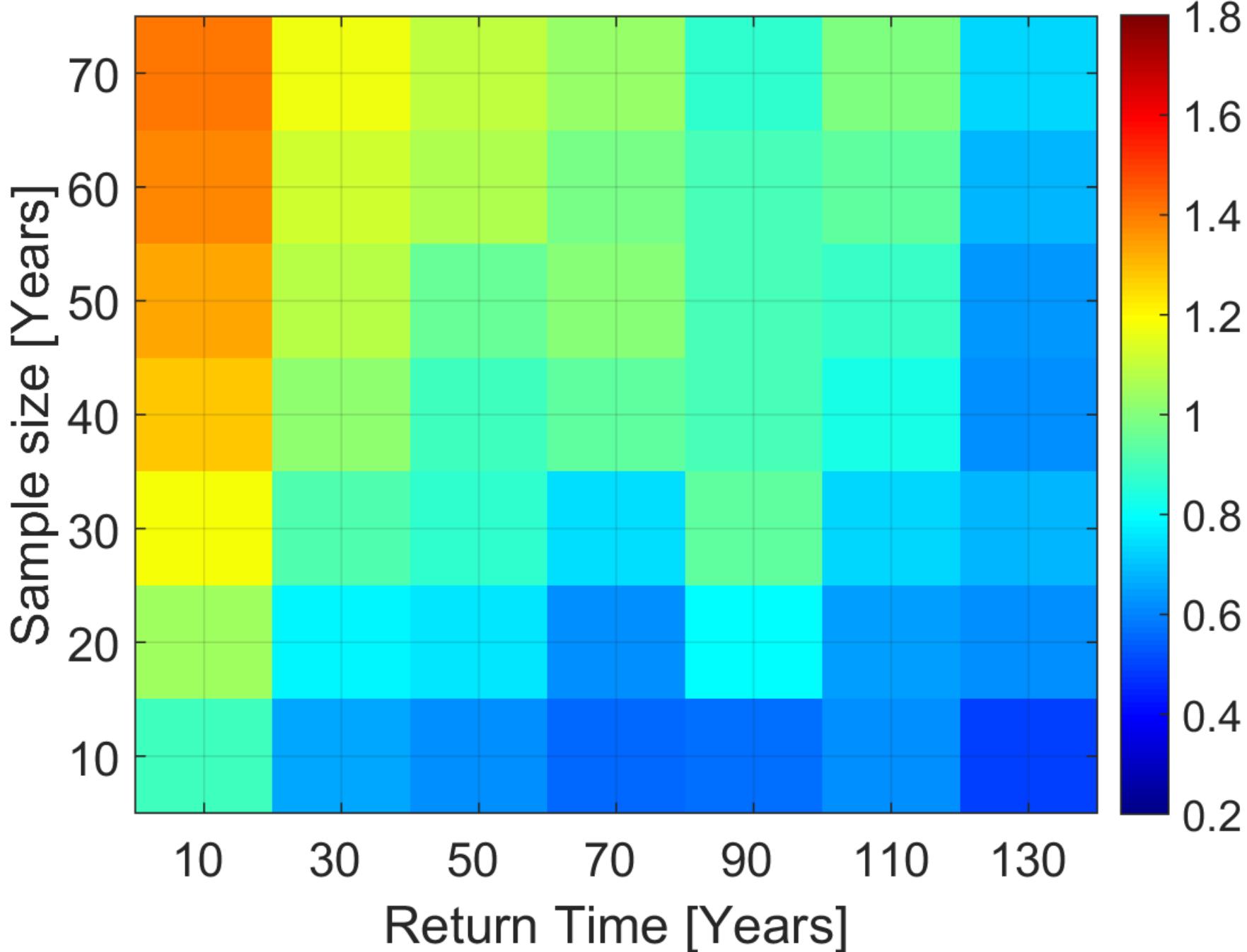
GEV/ POT are a good fit for the calibration sample but they fail in describing the stochastic process from which the sample has been generated

MEV allows a better description of the underlying process; less variance in high quantile estimation

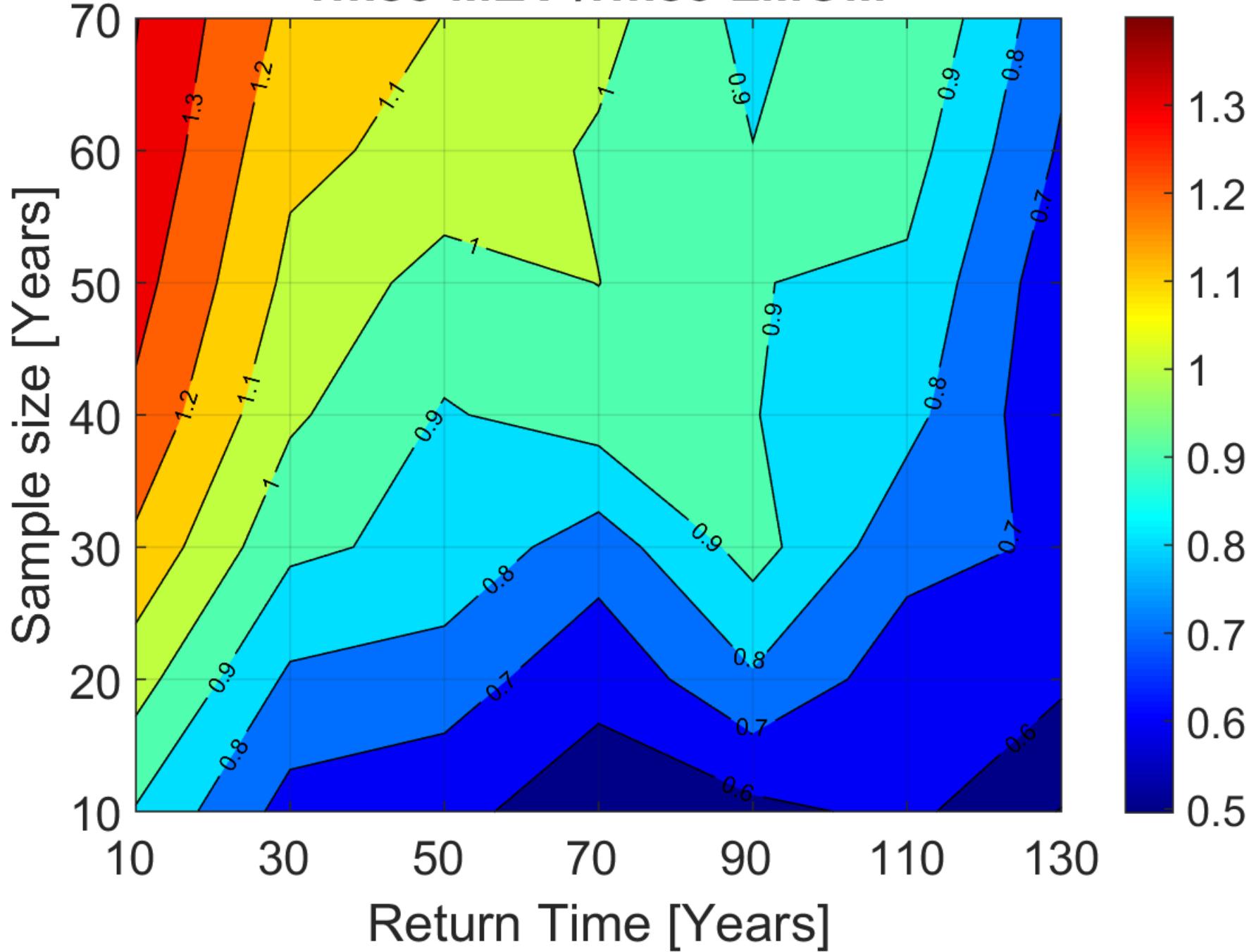
RMSE ratios for several Sample sizes and return times



rmse MEV /rmse LMOM



rmse MEV /rmse LMOM



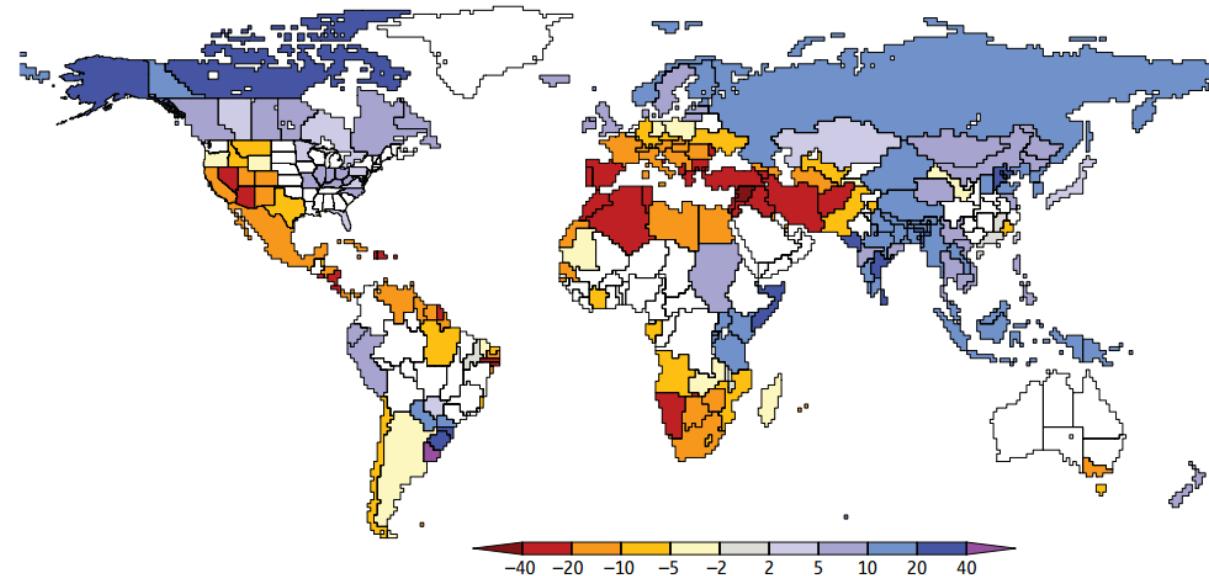
CLIMATE CHANGE

Stationarity Is Dead: Whither Water Management?

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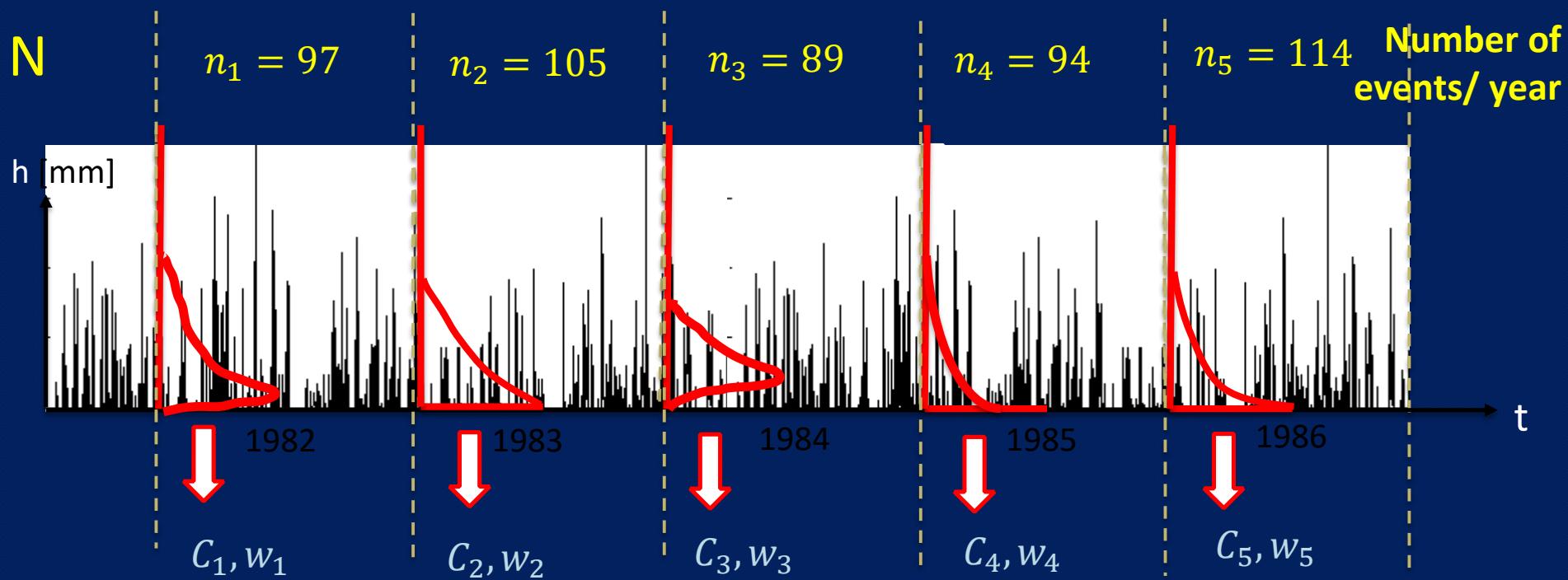
Systems for management of water throughout the developed world have been designed and operated under the assumption of stationarity. Stationarity—the idea that natural systems fluctuate within an unchanging envelope of variability—is a foundational concept that permeates training and practice in water-resource engineering. It implies that any variable (e.g., annual streamflow or annual flood peak) has a time-invariant (or 1-year-periodic) probability density function (pdf), whose properties can be estimated from the instrument record. Under stationarity, pdf estimation errors are acknowledged, but have been assumed to be reducible by additional observations, more efficient estimators, or regional or paleohydrologic data. The pdfs, in turn, are used to evaluate and manage risks to water supplies, water-

Climate change undermines a basic assumption that historically has facilitated management of water supplies, demands, and risks.



The MEV distribution

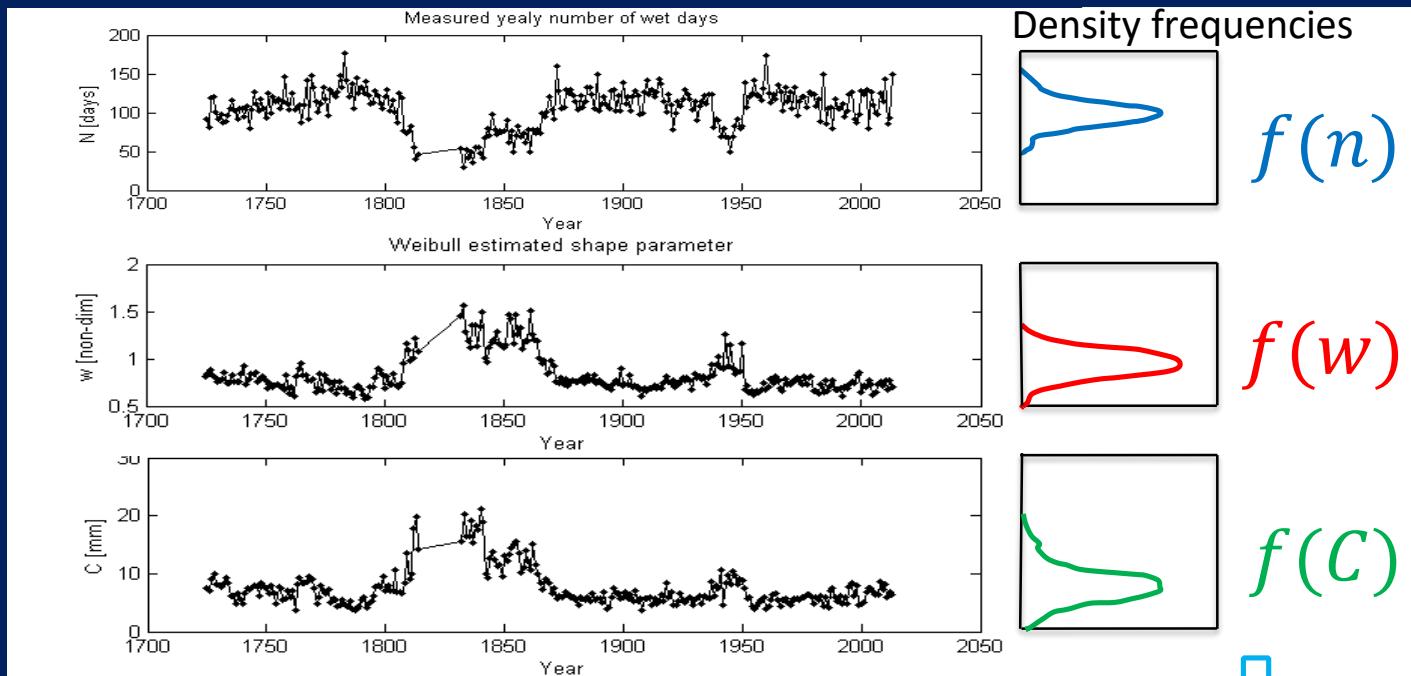
- Assuming Weibull as a pdf for daily rainfall
- Fit performed using Probability Weighted Moments (Greenwood et al, 1979)



$$F(x) = 1 - e^{\left(\frac{x}{C}\right)^w}$$

- Sampling n from the distribution $p(n | C, w)$
- Fit Weibull to the single years $\longrightarrow C_i, w_i$

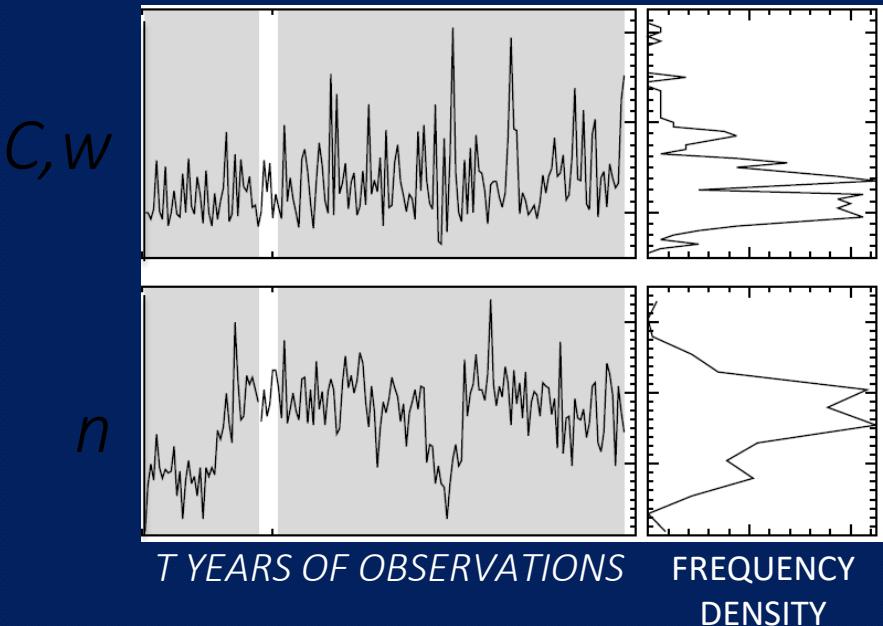
The Metastatistical Extreme Value Distribution



$$\zeta(x) = \sum_{n=1}^{\infty} \iint_{\mathcal{C} w} g(n, C, w) \left[1 - e^{-\left(\frac{x}{C}\right)^w} \right]^n dCd w$$

- Weibull parameters $\vec{\theta} = [C, w]$ and N are random variables themselves
- The CDF of annual maximum is the mean on all their possible realizations

A discrete expression for MEV



$$\int p(x)f(x) \cong \sum_i f(x_i) \text{ if } x_i \sim p(x)$$

$$f(C) \rightarrow \\ f(w) \rightarrow \\ \zeta(x) \cong \frac{1}{T} \sum_{j=1}^T \left\{ 1 - e^{\left(\frac{x}{C_j}\right)^w} \right\}^{n_j}$$

$$f(n)$$



Instead of an analytical model for $g(n, C, w)$
we use the observed frequencies

An alternative approach consists in simplifying the MEV expression (e.g.
Hypothesizing independence among N, C, w):

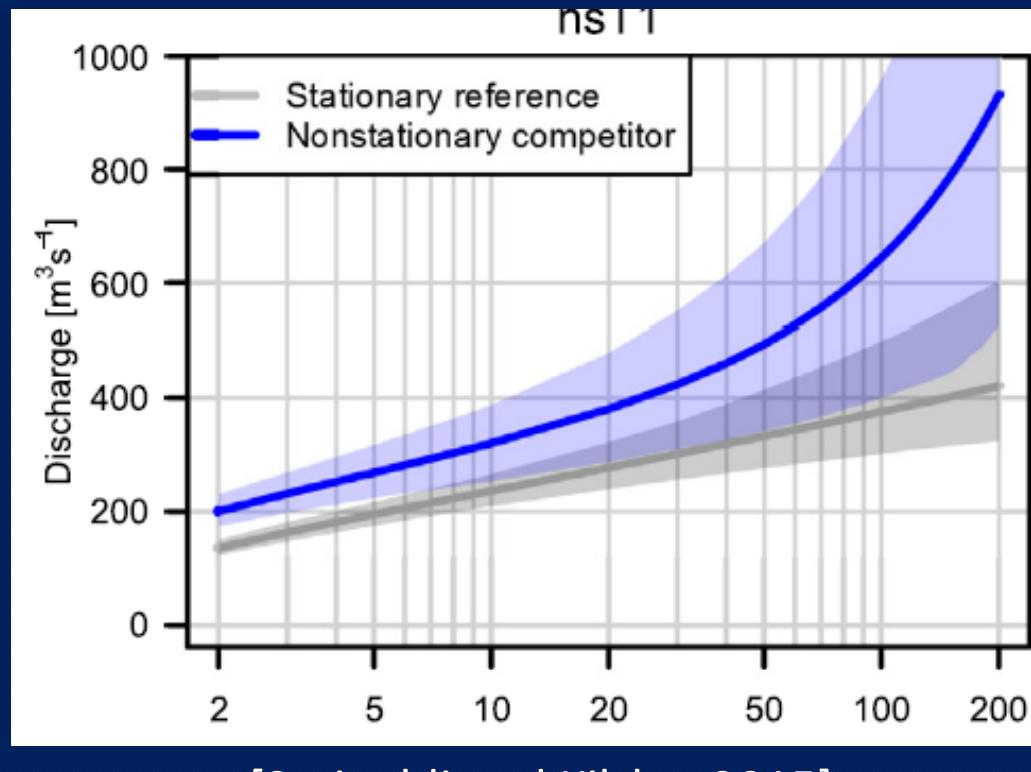
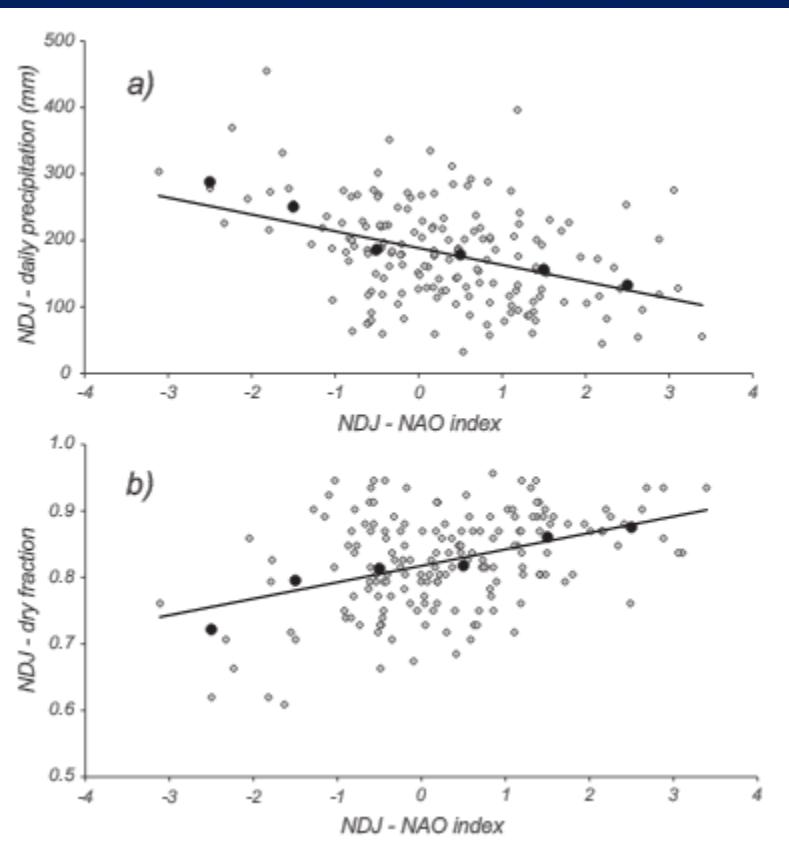


$$\zeta(x) = 1 - \sum_n n \cdot p(n) \cdot e^{\left(\frac{x}{C}\right)^w}$$

Penultimate
Approximation

What to do next?

Nonstationary EV analysis: Dependence of the parameters {N,C,w} on climatological parameters (downscaling)

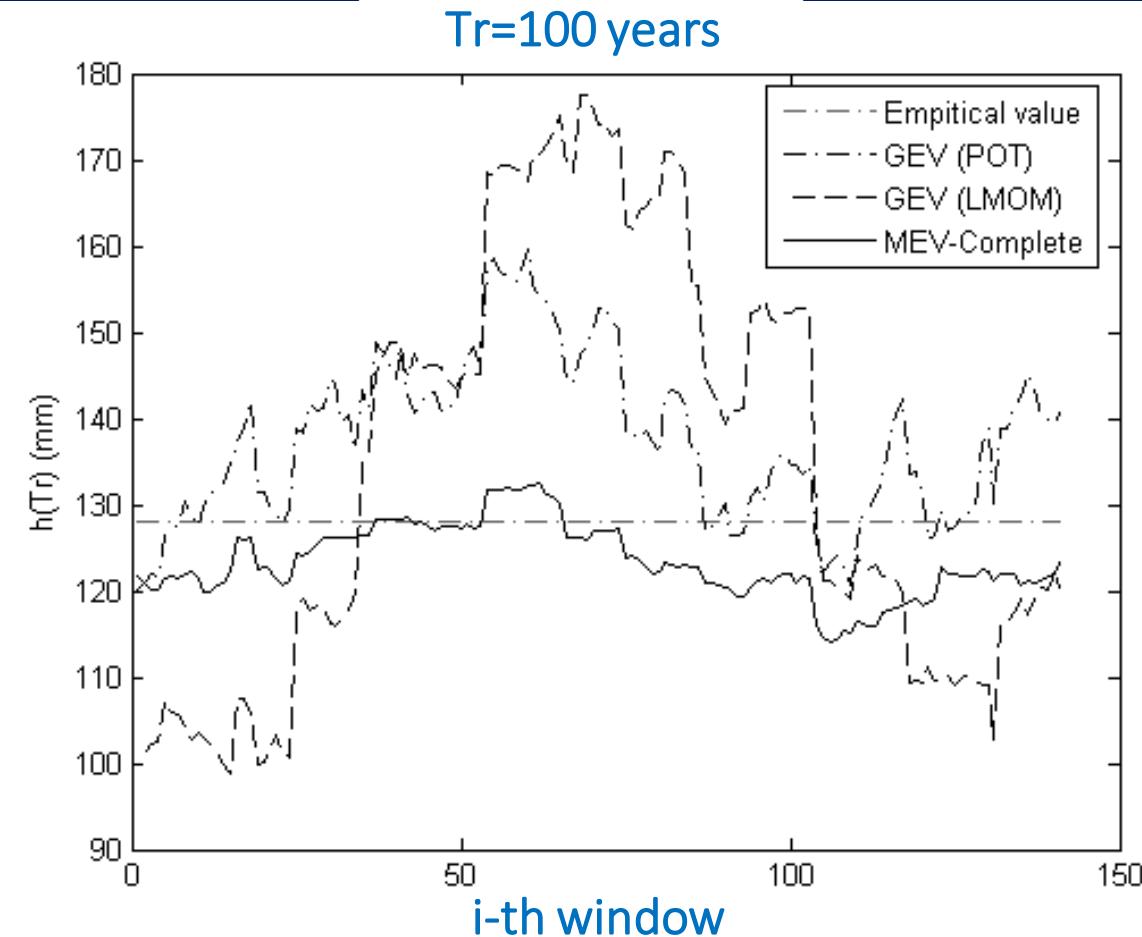


[Serinaldi and Kilsby, 2015]

[Marani and Zanetti, 2014]

Non stationary analysis

Bologna (Italy) randomly reshuffled time series
Sliding and overlapping windows analysis



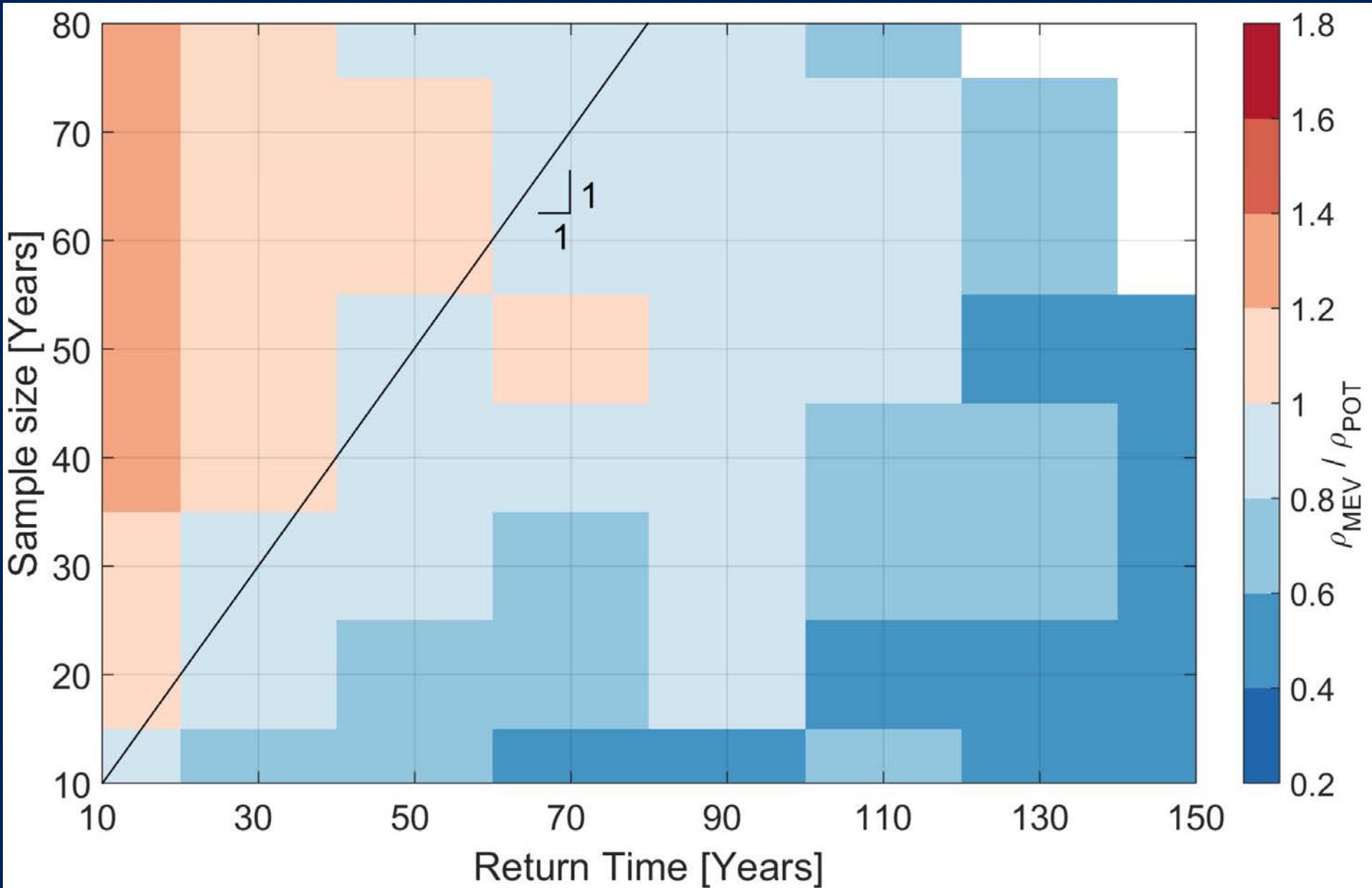
GEV and POT estimated quantiles show oscillations with same amplitude



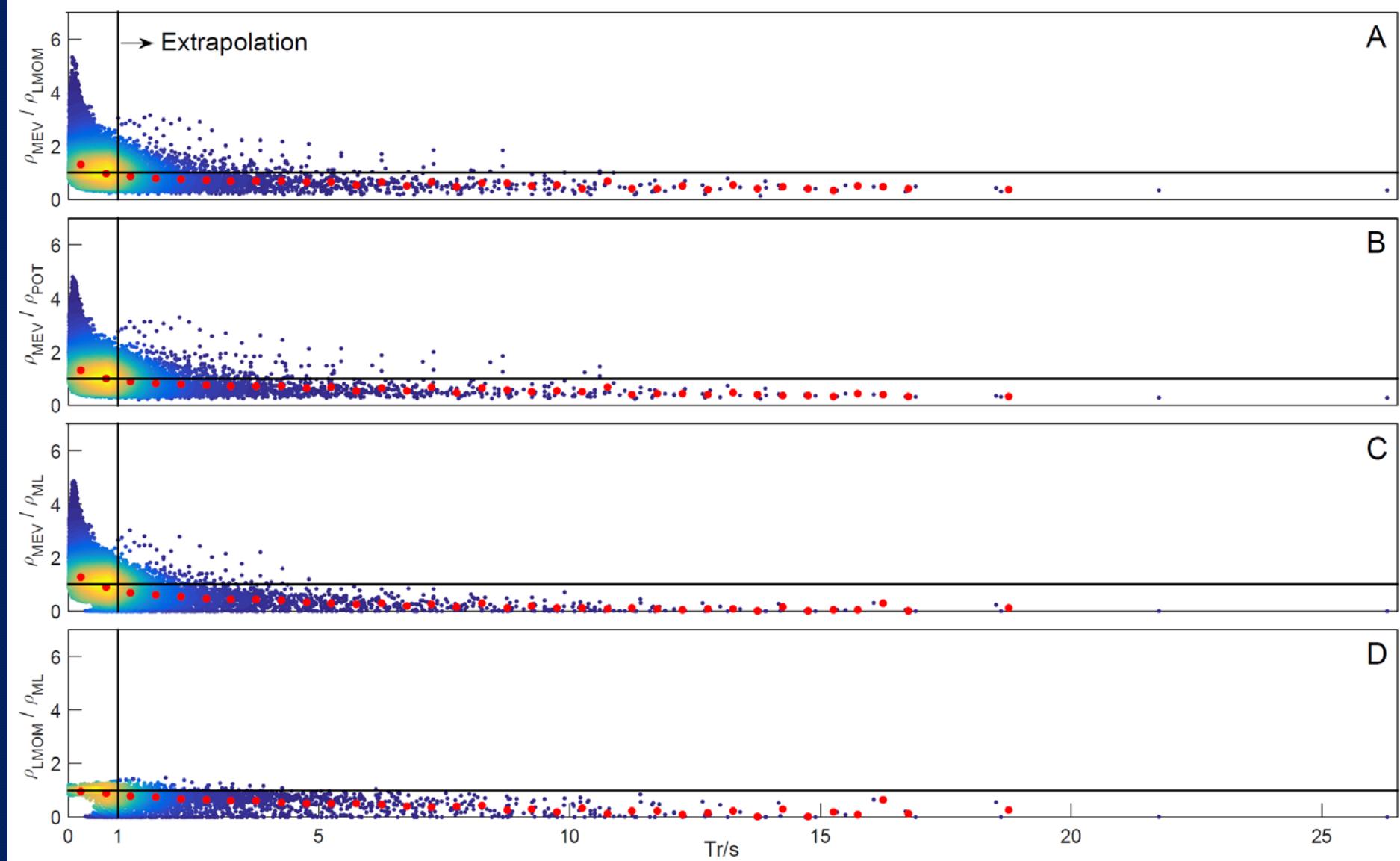
Due to the variance in the parameter estimates

In the case of MEV the variance of estimated quantiles is much smaller; Stationary behaviour

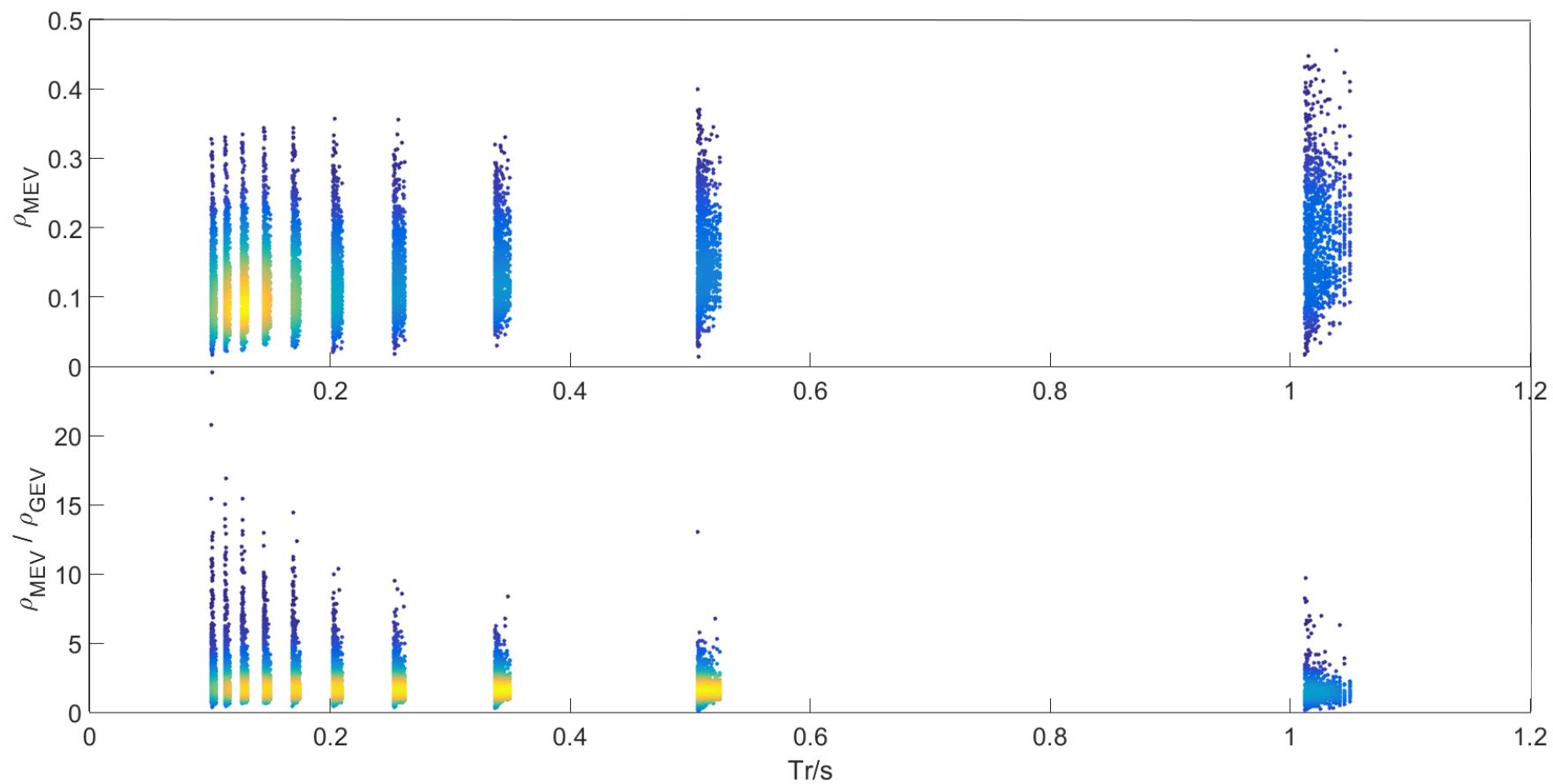
Ratio of MEV estimation error to GEV-POT error



Estimation error as a function of Tr/(sample size)



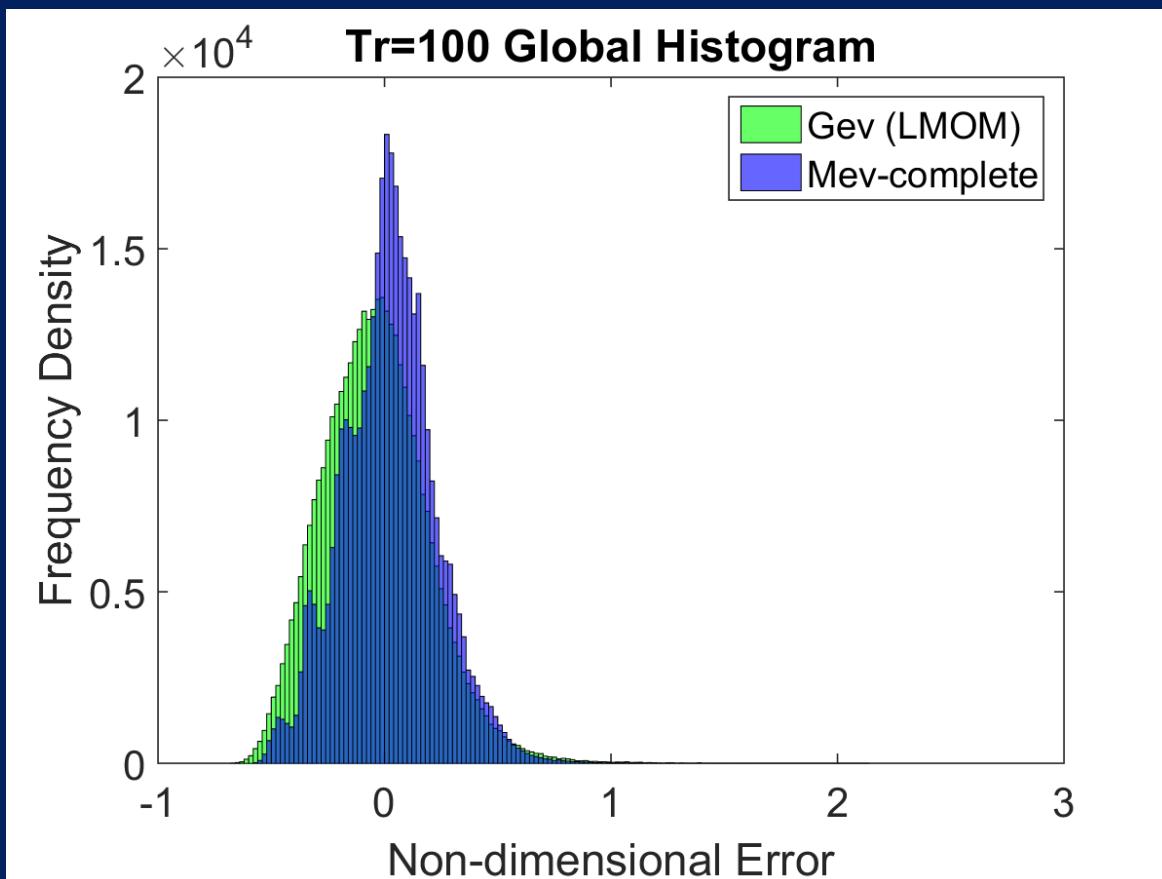
Performance when testing sample = calibration sample



Error distribution

Distribution of the error computed over 1000 random reshuffling, for all the analyzed datasets.

Quantiles ($Tr=100$ yrs) estimated by **GEV**, **POT**, **MEV**
calibrated over 30-years samples

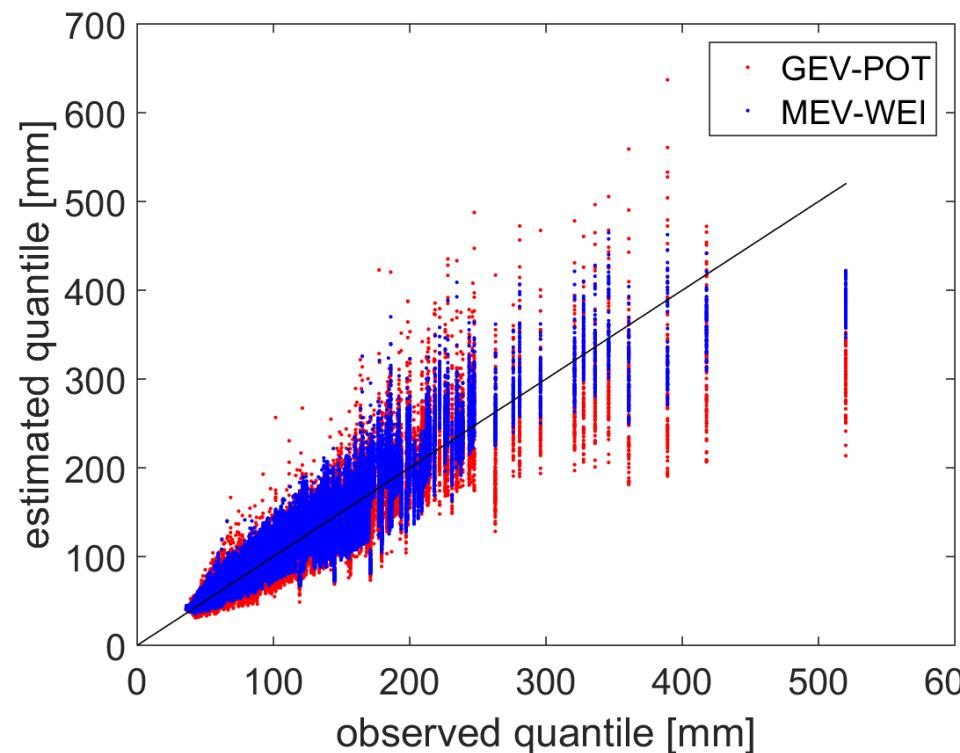
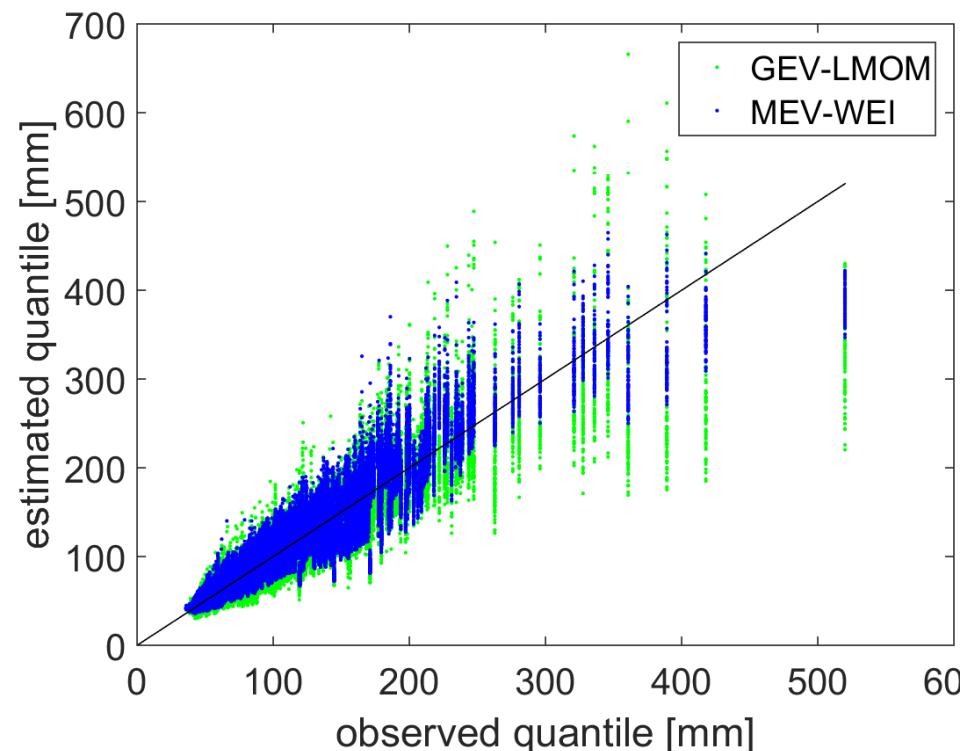


$$\epsilon = \frac{\hat{x} - x_{obs}}{x_{obs}}$$

$$\hat{x} = F^{-1} \left\{ 1 - \frac{1}{Tr_i} \right\}$$

x_{obs} from the observational
(independent) sample

Global QQ-Plots

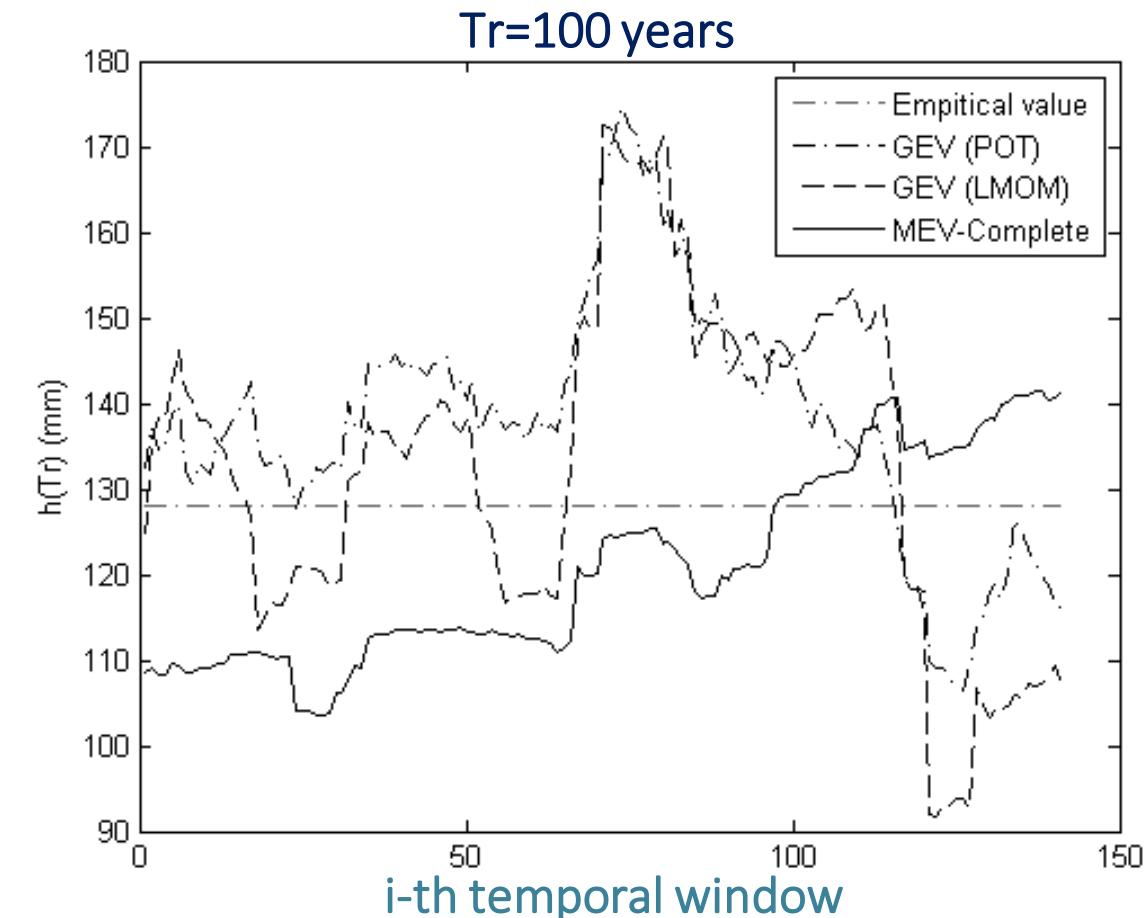


Sample size=45 years

100 random reshuffling

Some thoughts on non stationarity

Bologna (Italy) original 180 years time-series
Sliding and overlapping windows analysis



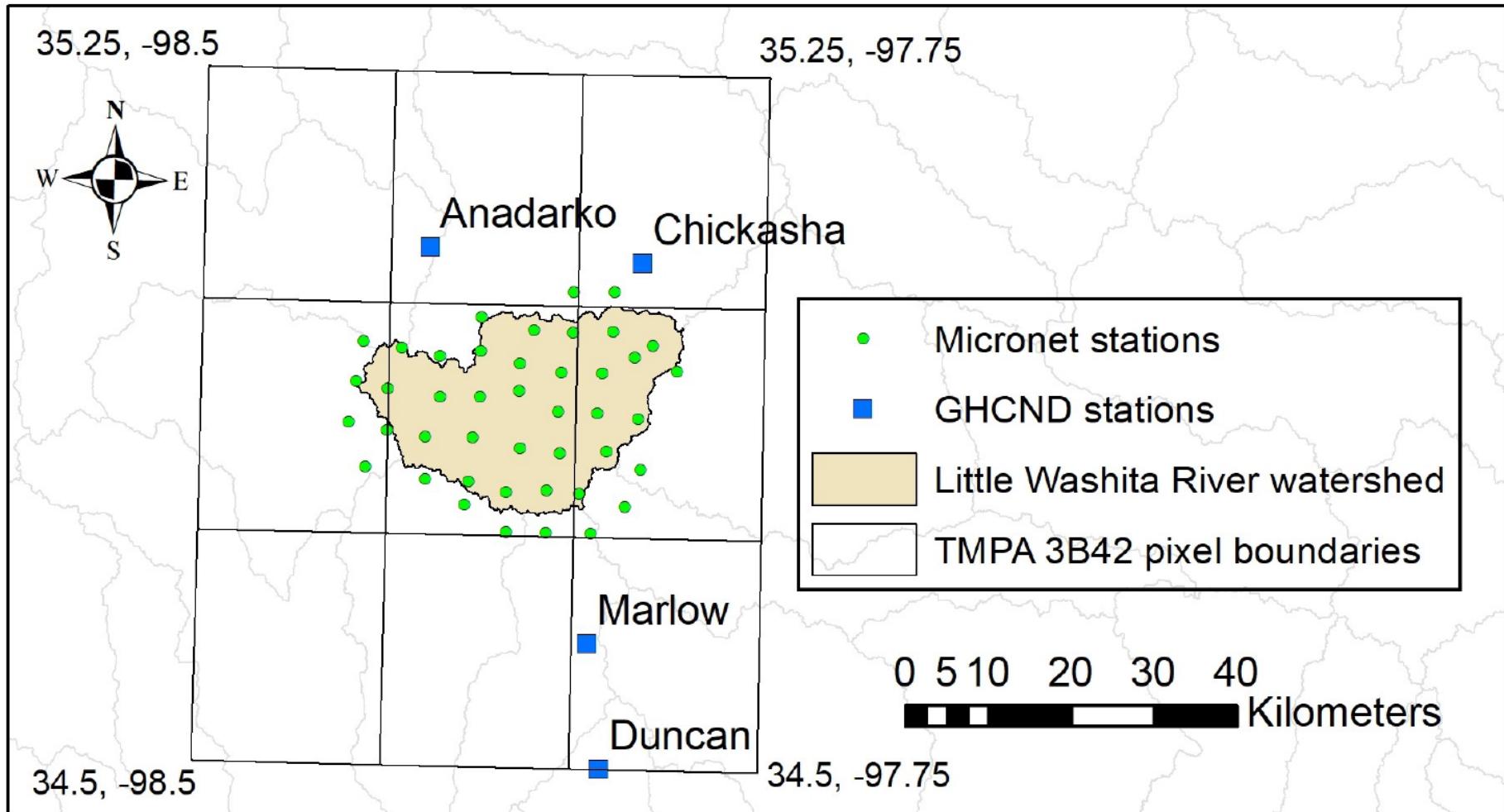
GEV and POT estimated quantile shows higher variance

MEV shows a positive trend in est. quantiles



Due to trend in the parameters of Weibull C, w and n

A case study over Oklahoma



[Elliot et al., 1993]

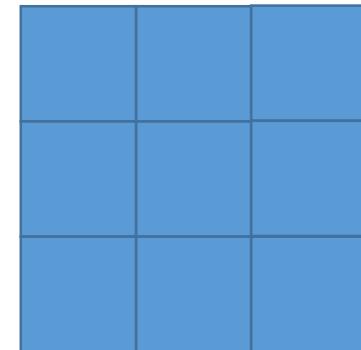
Zorzetto and Marani, WRR, 2018

Parameters of rainfall pdf downscaled using closed-form relations from space-time stochastic rainfall models (Marani, 2003)

TRMM observations (1 pixel)



TRMM observations (3x3 pixels)



Point-scale MEV distrib. (C, w, n)



TRMM 1-pixel MEV distrib. (C, w, n)





Billion Dollar Weather/Climate Disasters

1980 - 2011
NOAA/NESDIS/NCDC

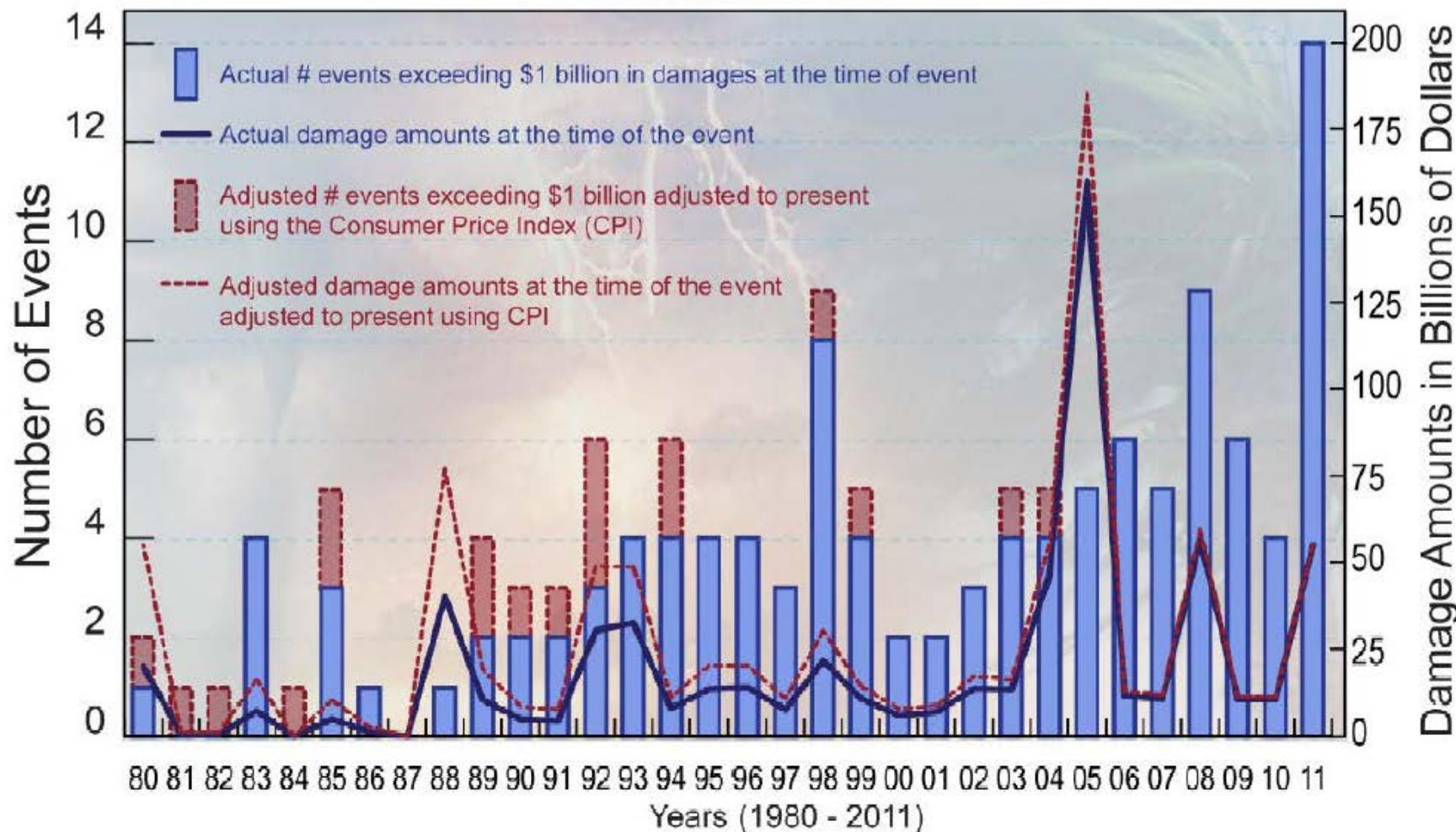


Fig. 1 US Billion-dollar Weather and Climate Disaster time series from 1980-2011 indicates the number of annual events exceeding \$1 billion in direct damages, at the time of the event and also adjusted to 2011 dollars using the Consumer Price Index (CPI)